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“Landmines: the Local Effects of Demining”

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# LANDMINES: THE LOCAL EFFECTS OF DEMINING

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**ABSTRACT.** Anti-personnel landmines are one of the main causes of civilian victimization in conflict-affected areas and a significant obstacle for post-war reconstruction. Demining campaigns are therefore a promising policy instrument to promote long-term development. We argue that the economic and social effects of demining are not unambiguously positive. Demining may have unintended negative consequences if it takes place while conflicts are ongoing, or if they do not lead to full clearance. Using highly disaggregated data on demining operations in Colombia from 2004 to 2019, and exploiting the staggered fashion of demining activity, we find that post-conflict humanitarian demining increases economic activity and students' performance in test scores, especially in areas with better market access. In contrast, economic activity does not react to post-conflict demining events carried out during military operations, and it *decreases* if demining takes place while the conflict is ongoing. Rather, demining events that result from military operations are more likely to exacerbate extractive activities and promote deforestation.

**Keywords:** Landmines, demining, conflict, peace, local development.

**JEL codes:** D74, P48, Q56, I25.

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## 1 INTRODUCTION

Landmines –explosives buried under the surface that trigger upon contact–are one of the most pressing challenges to post-war recovery and to long-term development. Together with unexploded ordnance (UXO), landmines threaten people’s life and mobility, therefore affecting agricultural investment, access to markets and basic services, and schooling (UN General Assembly, 1996; CNMH, 2017). Since they are cheap to fabricate and their use is widespread in internal armed conflicts, landmines also constitute an obstacle to humanitarian aid and post-conflict reconstruction (Parker, 2018). Estimates suggest that the current global stock of buried anti-personnel landmines is 110 million, distributed across 60 countries.<sup>1</sup> The stockpile of landmines yet to be planted is more than double that figure. Every year, 26,000 people on average get killed or injured by landmine blasts (Hall, 2017), and about 42% of the victims are children. Since landmines are hard to detect and costly to remove, their damage extends well after the end of the war.<sup>2</sup>

The colossal current and expected costs of landmine explosions imply that demining campaigns are one of the most pressing and socially profitable post-conflict endeavors. However, and perhaps surprisingly, research on the economic effects of demining is rather scarce and largely based on statistical associations. A recent notable exception is Chiovelli et al. (2019), who study the causal effect of the efforts undertaken since the end of the civil war in 1992 to strip Mozambique from its landmine threat. Exploiting municipal and yearly variation on the implementation of landmine clearance campaigns, the authors document a large positive impact on nighttime lights.

Building upon this finding, and using detailed geo-referenced information for the case of Colombia, this paper shows that, in addition to increasing economic activity, peacetime mine clearance also increases population density and improves students’ performance by reducing the costs of school attendance and increasing key school inputs. In sharp contrast, we also document that, a different type of demining that is not present in the Mozambican case, namely mine removal in military operations, *does not* boost any economic or social outcome. Instead, this alternative demining “treatment,” which is neither aimed at relieving entire mined areas nor at helping affected communities, exacerbates violent territorial contestation and reduces nighttime lights and population density if carried out while the internal conflict is still ongoing. In short, in this paper, we both provide external validity to Chiovelli et al. (2019) and complement their evidence with a wide range of additional –and more

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<sup>1</sup>See <http://www.landminefree.org/2017/index.php/support/facts-about-landmines> (last accessed 8/22/2021).

<sup>2</sup>While building a landmine can cost between \$3 and \$75, removing it requires an investment of up to \$1,000 (Doswald-Beck et al., 1995). If landmine fabrication and planting would halt, it would take over a thousand years to strip the entire planet of landmines at the current demining rate.

disaggregated–socio-economic indicators and land use measures. But, most importantly, we offer a much more nuanced view to the idea that demining is unambiguously desirable.

Colombia is an interesting laboratory to study the economic and social effects of demining, as it is the country with the second-highest number of landmine victims after Afghanistan ([Landmine Monitor, 2019](#)), and that with the highest number of victims of improvised (handmade) mines, which are cheaper to fabricate but much more dangerous than traditional antipersonnel mines.<sup>3</sup> Colombia has experienced different demining strategies, some of which had not been previously studied despite being quite common in internal armed conflicts throughout the world. First, we examine comprehensive humanitarian mine clearance campaigns, which refer to the thorough efforts of local and international non-governmental organizations (NGOs) to locate mine fields and work with the support of local communities to remove all the existing landmines until the area can confidently be called mine-free. Second, we look at demining events that result from military anti-insurgency operations, advancement, or maneuvers. Such cases seldom result in the official clearance of entire areas.<sup>4</sup>

Our identification strategy exploits the exact coordinates of all demining events and relies on the timing of demining campaigns that took place both throughout the conflict and after the start of peace negotiations with the *Revolutionary Armed Forces of Colombia* (FARC from its Spanish acronym). Specifically, we compare the evolution of various outcomes of interest in areas subject to demining and areas that are known to host anti-personnel mines but that had yet-to-be or were never demined during our sample period. Because our data on demining is geo-referenced, we focus on highly disaggregated local effects within a 5 Km radius in the baseline results.

Our estimation of the local causal effects of demining takes into account the recently documented problems of using two-way fixed effects to estimate causal effects in difference-in-differences settings with staggered adoption and heterogeneous treatment effects. First, we assess how relevant this is for our context by computing the decomposition suggested by [Goodman-Bacon \(2021\)](#), and we find evidence against using standard linear techniques. Second, our baseline specification uses the estimator proposed by [Callaway and Sant’Anna \(2020\)](#) that is based on a parallel trends assumption, and computes group-time average treatment effects (ATTs) that are later aggregated to compute an overall ATT. Third, we explore the robustness of our results to using alternative estimators, such as those proposed

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<sup>3</sup>Colombia is quite different than Mozambique. According to the World Bank’s Open Data (see <https://data.worldbank.org> –last accessed 09/14/21), Colombia’s GDP per capita is over 11 times that of Mozambique; Mozambique’s economy is much more dependent on agriculture (the share of agriculture value added over GDP is 26% versus 6% in Colombia); and Mozambique receives 3.5 times more official development assistance per capita than Colombia.

<sup>4</sup>Section 2 provides a detailed account of each type of treatment.

by Borusyak et al. (2021), De Chaisemartin and d’Haultfoeuille (2020), and Wooldridge (2021). Fourth, to assess the validity of our main identifying assumption (namely, that the post-treatment potential outcomes of treated cohorts have the same trend as those of the never treated and the soon-to-be-treated cohorts), we report the dynamics of the estimator, as well as corrections for potential bias coming from pre-treatment differential trends (Roth, 2021), and the robustness of our results to moderate linear and non-linear violations of the parallel trends assumption (Rambachan and Roth, 2021). Fifth, we show the robustness of our results to: i) adding municipality-specific time trends, as well as pre-treatment municipality and event-level covariates in a doubly-robust fashion as suggested by Sant’Anna and Zhao (2020); ii) using different radii around the demining event; iii) using different comparison units and periods; iv) accounting for potential spatial spillovers coming from the fact that current demining is likely to take place in the vicinity of past demining; and v) winsorizing outliers in our dependent variables.

The economic and social effects of demining are manifest in a wide set of outcomes. Because the presence of landmines hinders rural investment, market access, and mobility, we first look at whether demining affects economic activity and other social interactions likely to result in nighttime light density changes. Second, we look at the effect of demining on students’ performance. The relevance of studying both economic activity and learning outcomes is exacerbated by the observation that, according to our database, a large share of landmines are planted on or near rural roads or in the vicinity of schools. Third, because demining lowers the entry barriers of both productive and extractive rural investments, it may also increase deforestation. We thus also explore the effect of demining on forest loss. Fourth, in the Colombian context, the new rural investors that demining attracts may be associated with illegal armed groups, and particularly paramilitary militias. Therefore, we also study how demining affects conflict outcomes, alluvial illegal gold mining, and illicit extractive activities such as coca crops.<sup>5</sup> Moreover, the latter may trigger further deforestation pressures. This set of outcomes may also respond to the dynamics of demining inasmuch as illegal groups often protect their strongholds and illicit crops with mine fields (Fundación Seguridad y Democracia, 2006; CNMH, 2017).

We find that humanitarian demining campaigns that took place during the post-conflict period led to improvements in socio-economic conditions. In particular, we find a 12.4% increase in nighttime luminosity and a 2.7% increase in population density. A back-of-the-envelope calculation based on Henderson et al. (2011) yields that each humanitarian demining event increases the municipal GDP by 0.8%, and each dollar invested in humanitarian demining yields \$7 in benefits after only one year.

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<sup>5</sup>Coca leaves are the main input of cocaine, of which Colombia is the main exporter worldwide, including about 90% of the U.S. market.

We also find a 6.7 (8.1) percentage points increase in the probability that students obtain at least a satisfactory score in standardized national tests, specifically in math and reading. The performance of students in national standardized exams improves as soon as demining takes place and persists thereafter. We argue that this immediate effect is related to mechanisms pertaining to both school composition and students' performance. On the one hand, humanitarian demining attracts new migrants to the cleared region (as noted by its positive effect on population density), which also results in an increase in school enrollment and the opening of new schools in the area. On the other, demining increases the share of students promoted to the next grade and generates a net reduction in the students-to-teachers-ratio.

We show that the positive effect that demining has on both nighttime lights and students' performance is larger in areas that are better connected to labor, inputs, and output markets through the available road network. Indeed, connectivity is key to reap the economic potential that follows mine clearance campaigns. Moreover, it facilitates school attendance and reduces absenteeism. We also complement the reduced-form analysis of the importance of road connectivity with a market access general equilibrium framework à la [Donaldson and Hornbeck \(2016\)](#), which allows us to compute economy-wide effects. We find that in the absence of humanitarian demining efforts, Colombia would have forgone 0.7 percent of annual GDP between 2013 and 2019.

In sharp contrast with the case of humanitarian demining, demining in military operations during the same period, which did not have the objective of achieving the complete clearance of targeted areas, had no statistically significant effects neither on nighttime luminosity nor on students' performance. Instead, it decreased population density. Something similar occurred during the conflict period when no humanitarian demining took place, and the only demining resulted from military operations. In these instances, demining also reduced population density. Importantly, it also caused a differential *reduction* in nighttime luminosity. We posit that one potential mechanism through which demining during conflict decreased population density and nighttime light is that it exacerbated violent territorial disputes. Indeed, because armed groups use landmines to prevent the territorial advancement of enemies, demining can trigger violent confrontations between groups as well as the victimization of civilians thought to collaborate with the enemy ([CNMH, 2017](#) and [Procuraduría, 2011](#)). We thus explore the municipal-level correlation between demining and variables related to the incidence of violence and forced internal displacement, and find that: i) these variables are positively correlated with military demining only and; ii) this correlation is much stronger during the conflict period.

Regarding the effects of demining on forest cover, we find that while post-conflict humanitarian demining had no significant effects on deforestation, demining events resulting from

military operations both throughout the conflict and after its ending caused large increases in deforestation. Relatedly, we document a differential increase in the use of wildfire, which by and large is associated with extractive agricultural activities, as well as an increase in illegal gold mining. To shed light on the potential mechanisms underlying this finding, we look at spatially disaggregated data on soil suitability. We find that the deforestation surge after military demining is more pronounced in areas that are suitable for extractive agricultural activities such as oil palm, cattle ranching, banana growing, rubber planting, and forestry. We interpret these findings as consistent with the idea that demining in military operations serves the interest of elites with stakes in extractive economies. Moreover, such investments are unlikely to result in the type of economic growth that is captured by nighttime luminosity. On the contrary, and as mentioned, demining during conflict seems to have reduced growth.

As for the effect of demining on coca crops, we find that demining events that took place during the post-conflict period decreased coca cultivation locally. This contrasts with demining during the conflict, which did not affect coca-growing and hence the potential production of cocaine. We also exploit the implementation of an illegal-crops substitution program that was implemented after the signature of the peace agreement with FARC to estimate heterogeneous effects parametrized by the program's presence. We find that the demining-led decrease in coca cultivation is driven by municipalities where the crop substitution program was implemented. This implies that the coexistence and mutual reinforcement of different policies (illegal-crops substitution and demining) can be an effective way to reduce illegal drugs production. This is particularly relevant given the failure of the War on Drugs in producing countries.<sup>6</sup>

In short, we find that the local effects of demining largely depend on the type of mine removal strategy as well as on its timing. When demining is carried out by domestic and international NGOs with the objective of protecting local communities and clearing entire areas from landmines, this increases productive economic activities, population density, and the quality of education as measured by students' test performance. Instead, if demining takes place in the context of military activity which does not result in mine-free zones, and especially if this occurs while the conflict is still active, then it hurts economic growth and reduces population density, while at the same time it increases the intensity of the conflict and exacerbates extractive economic activities. These stark differences are not entirely driven by the fact that humanitarian and military demining target different areas.<sup>7</sup> But the main goal

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<sup>6</sup>See Prem et al. (2021a) for a thorough review and a recent example for the Colombian case.

<sup>7</sup>In a robustness exercise, we use LASSO to select covariates and Crump et al. (2009) to find a common support in which the two are comparable according to targeted areas. Within such sub-sample, we find that humanitarian demining still improves the economic and social outcomes relative to its military counterpart.

of this paper is not to do a horse race of these two treatments. These are fundamentally different in nature and cannot be thought of as alternative policy instruments aimed at the same objective.

Our paper contributes to the recent evidence on the economic effects of demining, championed by [Chiovelli et al. \(2019\)](#). We do so in two key dimensions. First, we offer a comprehensive analysis of the effects of demining on various socio-economic and political outcomes, including –but not limited to–nighttime lights. For instance, studying the effect of demining on educational outcomes allows us to better understand the potential long-term socio-economic consequences of landmines’ removal, that cannot be accounted for by nighttime lights. Looking at population density helps us gauge the potential effects of demining on internal migration patterns. We also explore a variety of outcomes that provide evidence on the effects of mines’ clearance on desirable and adverse land use dynamics. Second, we study demining activity under different institutional environments and with different scopes, as well as distinguishing between demining during an ongoing conflict and during the post-conflict period. This approach allows us to put in perspective the idea that demining is unquestionably good.

Our paper also contributes to the literature that studies the long-term economic effects of the massive U.S.-led aerial bombing campaigns that took place during the Vietnam War (1955-1975), the Cambodian Civil War (1967-1975), and the Laotian Civil War (1959-1975). While some authors find large negative long-term effects in terms of economic activity and agricultural productivity (see, e.g., [Lin, 2020](#) for the case of Cambodia and [Riaño and Valencia Caicedo, 2020](#) for the case of Laos), others conclude that the Vietnam aerial bombing campaign had no long-term effects in terms of poverty ([Miguel and Roland, 2011](#)) and political attitudes ([Dell and Querubin, 2018](#)). Whether the short-term devastation caused by bombings persisted over time or dampened down seems to depend on post-war policy responses related to public investment and public goods provision, as well as on plausibly exogenous factors such as soil quality (since bombs were more likely to explode in barren and less fertile soil). We posit that the long-term negative effects of buried aerial UXO are likely only a fraction of those that come from mines that are cheaper to fabricate and intentionally hidden underground instead of accidentally unexploded. In fact, landmine contamination is still a problem in around 60 countries, while unexploded aerial bombs are currently prevalent only in Cambodia and Laos. Finally, we also contribute to the literature studying the effects of landmine contamination on health ([Arcand et al., 2015](#)), education ([Merrouche, 2011](#)), and poverty ([Merrouche, 2008](#)).



## 2 CONTEXT

**2.1 Colombia’s civil war and the peace process** The start of Colombia’s internal armed conflict dates back to the 1960s, when FARC and the *National Liberation Army* (ELN from its Spanish acronym) were founded. The other main set of illegal actors of the Colombian conflict are right-wing paramilitary groups, originally armed by the state in the early 1970s and trained as self-defense organizations. Both the guerrillas and the paramilitary groups have extensively used landmines as a way to secure their strongholds as well as their control over areas that host illegal crops.

In October 2012, the Colombian government and FARC started peace negotiations in Cuba. FARC’s offensive activity quickly dropped by 98% (CERAC, 2016) and humanitarian demining efforts picked up. As a result, victims from anti-personnel landmines plummeted in FARC-affected municipalities (Perilla et al., 2021). In Figure A1, we document a large drop in the incidence of various conflict-related outcomes since the beginning of the peace negotiations. Based on these dynamics, we distinguish between a “pure conflict” period (from the first year of our sample, 2002, until 2012) and a “post-conflict” period, from 2013 onward.

**2.2 Landmines in Colombia** Colombia is the country with the highest number of victims of improvised anti-personnel mines, homemade explosives that detonate by contact *or even in the proximity* of a person or object. They are harder to detect and remove without risking an explosion (Landmine Monitor, 2019). The main milestone in the fabrication and planting of improvised mines in Colombia came in 2008, when FARC’s secretariat launched a strategy that they called *Plan Renacer Revolucionario de las Masas* (Revolutionary Rebirth of the Masses). In an internal secret memorandum, commander ‘Alfonso Cano’ instigated all fronts to strengthen their production and planting of landmines in order to protect their strongholds (see Appendix Figure A2 for a picture of the memo in the original Spanish). By 2017, the area contaminated with landmines was officially estimated to be around 11,400 acres (Landmine Monitor, 2017), is equivalent to almost 80% of the size of Manhattan.

As a signatory of the 1997 Ottawa Convention, Colombia adopted in 2002 the Information Management System for Mine Action (IMSMA) of the Geneva International Centre for Humanitarian Demining (GICHD), a registry of landmine explosions, suspicion of presence, and demining events. It also committed to clear all landmines from its territory by 2021 (which by now has been extended to 2025). However, due to the intensity and territorial reach of the ongoing conflict, large-scale humanitarian demining and full clearance operations did not pick up until after the start of peace negotiations with FARC.

**2.3 Humanitarian mine clearance versus demining in military operations** Humanitarian demining is the assistance provided by specialized NGOs to communities exposed

to the contamination of explosive devices (anti-personnel landmines and UXO, PAICMA, 2012). This entails three main activities. First, the NGO approaches the community and finds at least one local liaison who helps building links within the community. The objective is to better understand the local context and convey the NGO's principles of neutrality and impartiality to the broader community, in order to gain their trust and cooperation. Second, the NGO conducts community-level surveys to determine if there are Hazardous Areas (HA), suspected to be mined. Third, a technical team visits all HA to confirm the existence of landmines or UXO from a safe distance, using metal detectors and sound beams. If the HA is dismissed then it is certified as mine free. If it is confirmed then the technical team removes the existing mines or engages in controlled blasts until the area is declared as cleared. Demining NGOs also provide assistance to mine victims, as well as pedagogical seminars aimed at raising awareness and promoting safe behavior.<sup>8</sup> Because of the comprehensiveness of the process and the engagement of the community, the average time that the full clearance of an area takes with humanitarian demining is 16 months.<sup>9</sup>

Any national or foreign NGO may undertake humanitarian demining, and the first NGO that engaged in demining activity was the British NGO Halo Trust. It did so in 2013, the first year of the period that we label "post-conflict" in our statistical analyses.<sup>10</sup> In fact, since the start of peace negotiations, the parties agreed to allow the establishment of humanitarian demining campaigns, and in the final peace agreement, the involvement in demining activities was highlighted as a key activity for the reincorporation of former FARC combatants.<sup>11</sup>

The selection of targeted areas is done by the Inter-institutional Instance of Humanitarian Demining, a joint body composed by the Ministry of Defense, the General Inspection of the Military Forces, and the Office of the High Commissioner for Peace (Decree 3750 of 2011). The only variable that was used for prioritization was the number of past landmine accidents and victims.<sup>12</sup> We corroborate this in Figure 1, where the only pre-treatment characteristic that explains both the timing (Panel A) and intensity (in terms of the size of the targeted area, Panel B) of humanitarian demining is past landmine victimization. Other key variables, including the lagged value and first difference of our main outcomes, are not correlated with the timing or the intensity of the treatment. Importantly, however, in our regressions, we control flexibly for this selection variable and our results are unaffected.

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<sup>8</sup>Qualitative evidence suggests that humanitarian demining improves the perception of safety and the willingness to undertake productive activities in the land. Mine clearance also seems to have increased the value of land and attracted tourism (Mutual-Co, 2021).

<sup>9</sup>Own calculations using the data described in section 3.

<sup>10</sup>As of today Colombia hosts 7 demining organizations, 5 international and 2 local.

<sup>11</sup>See excerpts 3.2.2.6-part b, 4.1.3.1, and 6.2.3-part of the text of the agreement.

<sup>12</sup>For a detailed discussion, see [https://www.apminebanconvention.org/fileadmin/APMBC/clearing-mined-areas/work\\_plans/Colombia-strategic-plan-mine-action-2020-2025.pdf](https://www.apminebanconvention.org/fileadmin/APMBC/clearing-mined-areas/work_plans/Colombia-strategic-plan-mine-action-2020-2025.pdf) (last accessed 9/21/2021).

In sharp contrast to humanitarian demining, demining activities amid military operations have long existed as part of the dynamics of the internal conflict. The military constantly engages in mine removal and controlled explosions as part of their anti-insurgency operations and maneuvers, and as a way to clear paths for the advancement of troops and warfare equipment. Key to facilitating ground mobilization, military operations often include personnel from specialized anti-explosives groups, in charge of the detection and destruction of antipersonnel mines and UXO. But these activities are not linked to the government’s *Integral Action against Antipersonnel Mines*, they do not involve local communities and their objective is not to declare entire areas as cleared from mines. By March 2021, almost 25,000 demining events in military operations had been registered.

Humanitarian and military demining are therefore fundamentally different “treatments.” While in the former the clearance of an entire mined area is the main objective, in the second demining is an intermediate objective to achieve a strategic goal other than demining *per se*, and thus it is seldom comprehensive within demined areas and it is disassociated from wider community empowerment efforts. This implies that the results from looking at one or the other are not directly comparable, and it is important to highlight that such comparison is not our goal. However, in our empirical analysis, we do include, for robustness, an additional exercise in which we estimate the common support of the areas potentially affected by both humanitarian mine clearance and demining in military operations. We re-estimate our baseline models within that sub-sample and obtain qualitatively similar results.

### 3 DATA

**3.1 Demining** The IMSMA information system (see section 2) provides detailed geo-referenced data on all landmine demining events, as well as on the presence or suspected presence of antipersonnel mines from 2003 onward. As of March 31 of 2021, the database included 24,746 demining events in military operations, all geo-located with accurate military GPS technology.<sup>13</sup> Most of these events took place before 2013, when demining in military operations started dropping due to the de-escalation of the conflict that followed the start of the peace negotiations with FARC (see section 2). In contrast, humanitarian demining was very rare before 2013 due to the dangers of engaging in landmine clearance campaigns amidst an active internal armed conflict. The geo-coded records provided by all the humanitarian demining operations carried out by the seven active NGOs in charge of this activity coincides very accurately with the information included in IMSMA. As of June 31 of 2021, the NGOs had established 2,272 hazardous areas. Of these, 1,141 had been confirmed to

<sup>13</sup>Camp et al. (2016) show that in Ecuadorean Andes (a topography very similar to Colombia’s) the location error of the GPS information is 9.6 meters (with a standard deviation of 4.7 meters). Given the baseline radius that we use to estimate the local effect of demining (5 Km), a location error of such magnitude should not be a major concern.

host landmines and 645 had been cleared.

From these data, we code three treatments: i) humanitarian demining during the post-conflict period (2013-2019); ii) demining in military operations during the post-conflict period; and iii) demining in military operations during the period of active conflict (2004-2012). The coded information includes the exact location of all demining events and the year of occurrence. Moreover, and key for identification, we code the location of areas known to be contaminated but not yet demined (as of 2021).

**3.2 Outcomes** Since we focus on the local effects of demining (on areas located within a short radius from the demining event), our main set of outcomes is composed of variables that can either be geo-referenced or are available for relatively small spatial grids. This subsection describes such outcomes and discusses their measurement and sources.

*3.2.1 Nighttime lights* Nighttime light has been shown to be a reliable proxy for economic activity both nationally and in geographically small areas (Henderson et al., 2011; Bleakley and Lin, 2012; Michalopoulos and Papaioannou, 2013; Storeygard, 2016; Goldblatt et al., 2018; Martinez, 2021). It also may capture other types of human activities that result in higher electricity usage. We use the global harmonized nighttime light (NTL) dataset constructed by Li et al. (2020), which addresses all the known problems of nighttime lights, such as intercalibration, geometric correction, and blurring. Nightlight images are available for grids of 1Km<sup>2</sup>, so for our outcome of interest, we take all the pixels that intersect the buffer of a demining event and compute a weighted average of their luminosity value.<sup>14</sup>

Figure A4 corroborates that this measure is highly correlated with various socio-economic outcomes at the municipality level. These include value-added, mortality rate under five years old, an index for fiscal performance, literacy rate, and a poverty index. This correlation is high both for the entire country and for the relatively more rural municipalities, which host the vast majority of landmines and thus where most demining takes place.

*3.2.2 Population density* Population density is an alternative proxy of economic activity that has been widely used by economic historians. For instance, Acemoglu et al. (2002) argue that only prosperous areas can support dense populations, and that population density contributes to economic growth by encouraging the exchange of ideas. We, therefore, compute a buffer-specific population density measure. To that end, we use the 1Km<sup>2</sup> population rasters provided by WorldPop and the Center for International Earth Science Information Network

<sup>14</sup>The weights are given by the product of the luminosity value of each intersecting lit pixel and the fraction of the buffer area that overlays with that pixel. To deal with outliers and with observations with zero average light, we use the hyperbolic sine transformation of the average luminosity within the buffer.

(CIESIN).<sup>15</sup> We then compute the average of the estimated population density within each buffer of our sample.

*3.2.3 Schools and the performance of students in test scores* To assess the impact of demining on students' performance in standardized test scores, we start by constructing a novel geo-referenced database of all schools in Colombia.<sup>16</sup> We then merge to these data the average academic achievement of all the school students in the reading and math standardized national tests (called "Saber"), that are implemented yearly in selected grades (3rd, 5th, and 9th).<sup>17</sup>

We construct our measure of buffer level students' probability of obtaining at least a satisfactory score in the standardized test (which for simplicity we call students' performance) by computing the weighted average of the fraction of students enrolled in schools within the buffer that passed the test. The weight is the number of students that took the test in each school/year. Interestingly, we find that 96% (77%) of the buffers around a humanitarian (military) demining event have at least one school. Moreover, on average, nine schools fall within each (5 Km radius) buffer in our estimation sample.

*3.2.4 Extractive activity* We also measure buffer-specific yearly forest loss using the satellite-based estimates of the Global Forest Change (GFC) project, which includes information of changes in tree coverage with a resolution of approximately 30m<sup>2</sup>, estimated from Landsat images (Hansen et al., 2013). Deforestation is identified by GFC when a specific pixel changes its tree cover status from one year to the other. For each buffer of our sample, we compute the area occupied by pixels that became deforested in a specific year and use the hyperbolic sine transformation of this variable.

With the increasing use of fire as a deforestation tool, most of which is illegal, we explore the robustness of our deforestation results by using the data from NASA's Fire Information for Resource Management System (FIRMS). With this input, we compute the total number

<sup>15</sup>The data can be downloaded from <https://www.worldpop.org/geodata/listing?id=76>. The population raster is estimated by the source following the methodology proposed by Stevens et al. (2015), which uses disaggregated census data and a *Random Forest* machine learning model that includes remotely-sensed data and geographic administrative data to predict grid-level population values. The variables used for the prediction include various types of land use, nighttime lights, and climate and geographic characteristics, and the presence of local facilities.

<sup>16</sup>We did so by web-scraping the school information from the Education System of Educational Sites (SISE): <https://geoportal.dane.gov.co/SISE/sise/>. This is a cross-section database, but the change of school locations is very rare, especially in the more rural areas where demining takes place (Gómez Montoya et al., 2018).

<sup>17</sup>This information comes from administrative datasets of the Colombian Institute for the Evaluation of Education (ICFES from its Spanish acronym) and is available for the period 2012-2018 in the form of school-level averages. We do not include test scores from the end-of-high-school national test ("Saber 11"), as these are not comparable across years during our sample period, due to several methodological changes documented in ICFES (2019).

of fires that take place each year within each buffer.

We also have geo-referenced data on the existence and the geographic extension of illegal alluvial gold mining (EVOA from its Spanish acronym). This information is available in 1Km<sup>2</sup> grids for the years 2014, 2016, 2018, and 2019, and is estimated by the United Nations Office on Drugs and Crime (UNODC) using remote sensing methods. The illegality of the mines can be inferred after overlaying the EVOA raster with official geo-referenced data on mine titles.<sup>18</sup>

*3.2.5 Coca cultivation* To measure the size of illicit coca crops, the leaves of which are mixed with chemicals to produce cocaine and crack, we rely on the satellite-based annual estimation performed by the Integrated Monitoring System of Illicit Crops (SIMCI from its Spanish acronym) of UNODC. SIMCI uses satellite imagery to estimate coca production by the end of each calendar year with remote sensing tools, which are validated with high-definition photographs taken from a helicopter.<sup>19</sup> The data is produced in grids of 1Km<sup>2</sup> from 1999 to 2019. This allows us to compute the buffer-specific area covered with coca crops. Our dependent variable is the hyperbolic sine transformation of this measure.

The description of additional variables and data sources as well as the descriptive statistics are reported in Appendix A.

## 4 EMPIRICAL STRATEGY

**4.1 Staggered Difference-in-Differences** To study the local effects of demining before and after the end of the conflict with FARC, we exploit both the timing of demining events as well as their exact geo-referenced location. Our unit of analysis is the 5 Km-radius circle around the geographic location of the event.<sup>20</sup> Because demining activity takes place at different times along our sample period, we could estimate a staggered difference-in-differences specification of the form:

$$(4.1) \quad y_{it} = \alpha_i + \lambda_t + \beta \times Post_{it} + \varepsilon_{it},$$

where  $y_{it}$  are different measures of local activity measured within buffer  $i$  and in time  $t$ .  $Post_{it}$  is a dummy that takes the value one in buffer  $i$  after a demining event and zero

<sup>18</sup>Moreover, UNODC cross-checks the estimates with local environmental NGOs that often corroborate the existence of an alluvial gold mine. In 2019, 66% of the detected EVOA area corresponds to illegal exploitation, 27% corresponds to legal (titled) exploitation and the remaining 7% are mines in the process of being legalized UNODC (2020).

<sup>19</sup>SIMCI uses satellite images for a wide window around December 31st. In particular, about 70 percent of the images are obtained between mid-November of the year of the estimate and late February of the following year. Of the remaining 30 percent, roughly half is obtained from August to November of the year of the estimate, and the residual is obtained between March and April of the following year.

<sup>20</sup>In the Appendix, we show the robustness of our results to different radii, specifically 3, 4, 6, and 7Km around the location of the demining event.

otherwise.  $\alpha_i$  are event/buffer-level fixed effects and  $\lambda_t$  are year fixed effects.

A recent literature had documented that this type of *two-way fixed effects* (TWFE) model can suffer from a severe bias, which makes the estimated coefficient of interest ( $\beta$ ) different from the true average treatment effect on the treated (ATT). This is likely to occur when treatment effects are heterogeneous over time and across units. To assess the extent to which this is the case in our context, we start by performing the decomposition suggested by [Goodman-Bacon \(2021\)](#). We report the results in Appendix Table A1, and find that in the case of humanitarian demining, 5% of the TWFE estimate comes from the “forbidden comparison” (that uses the early treated units as controls for units treated later). The proportion is much larger for the case of military demining during conflict (11%) and after its termination (27%). In turn, these figures are consistent with the fact that the share of never-treated units is relatively low, especially in the case of post-conflict military demining events.<sup>21</sup> This implies that using TWFE in our context would most likely lead to biased estimates.

Given the results of these diagnostics, we follow the recent developments regarding the estimation of these types of models. In particular, we follow the [Callaway and Sant’Anna \(2020\)](#)’s procedure, which estimates group-time ( $g, t$ )-specific ATTs avoiding incorrect comparisons. These are then aggregated in ways that allow the presentation of both “event-study” figures and average estimates, using a range of potential weighting functions.<sup>22</sup> Importantly, in the Appendix, we corroborate that our results are robust to using alternative estimation methods, that also address the potential problems of the TWFE model, including those suggested by [Borusyak et al. \(2021\)](#), [De Chaisemartin and d’Haultfoeuille \(2020\)](#), and [Wooldridge \(2021\)](#).

One key feature of these types of models is the inclusion of a set of “never-treated” units, that however, *could have been* treated. To this end, we need to identify mined areas that were not demined during our sample period, but that could have been so. For the case of post-conflict (2013-2019) military demining, we use as never treated the demining events that occurred in 2020 and 2021, after the end of our sample period. For the case of humanitarian demining (which happened primarily in the post-conflict period), we use as never treated both the 2020-2021 demining as well as the areas confirmed to have mines but not yet demined due

<sup>21</sup>Since the estimated  $\beta$  is a weighted average of event-specific ATTs, we follow [De Chaisemartin and d’Haultfoeuille \(2020\)](#) to compute the share of ATTs that enter the computation with a negative weight. Consistent with [Goodman-Bacon \(2021\)](#)’s decomposition, we find that the share of negative weights is zero for humanitarian demining and 12% (27%) for demining in military operations carried out during the conflict (in the post-conflict period).

<sup>22</sup>We use the “simple” aggregation (as recommended by the authors) that uses as weight the size of the group-year cell. However, we also present the group-level aggregation in Table A2 of the Appendix, which first computes the ATT for each cohort  $g$  and then takes the average across them.

to the limited capacity of humanitarian demining NGOs. Finally, for the case of military demining during the course of the conflict with FARC (2004-2012), we use as never treated the military demining events carried out in 2013.<sup>23</sup>

Figure A6 reports, for each demining type, the number of treated units by year together with the never treated. It can be concluded that, while the number of never-treated units used for the analysis of humanitarian demining and military demining during conflict is fairly large, that available for post-conflict military demining is relatively small.<sup>24</sup> This suggests that using the “not-yet-treated” units as a complementary comparison group is important in our context.<sup>25</sup>

Finally, as suggested by Callaway and Sant’Anna (2020), we balance our estimation period around the event data, to avoid the estimates being confounded by changes in the weights driven by sample composition. Specifically, we use three years before and three years after the demining event.

*4.1.1 Identifying assumption* The main identifying assumption for the “not-yet-treated” version of the Callaway and Sant’Anna (2020) estimator is that the evolution in potential outcomes after the treatment is the same for treated cohorts  $g$  and never-treated (and/or soon-to-be-treated) units. We present the dynamic treatment effect version of the authors’ estimator in order to partially assess the validity of this assumption (Marcus and Sant’Anna, 2021). We also present the corrections for pre-testing bias and bias from a pre-demining linear trend following Roth (2021), as well as for the robustness of our results to moderate linear and non-linear deviations from the parallel trend assumption following Rambachan and Roth (2021).

## 5 MAIN RESULTS

This section discusses our estimated results. We start by summarizing our findings regarding the impact of post-conflict humanitarian demining efforts and then turn to that of demining in military operations. We then assess the validity of the main identifying assumption of our empirical strategy and summarize a battery of robustness tests that we report in Appendix B.

**5.1 Post-conflict humanitarian demining** Table 1 reports the main results concerning the effects of the humanitarian demining efforts that started after the end of the conflict. We do so in terms of six substantive outcomes, which are likely affected by demining given the

<sup>23</sup>This follows Callaway and Sant’Anna (2020)’s advise of using late cohorts as ‘never treated’.

<sup>24</sup>While it is impossible to know with the data at hand, it may also be the case that the never treated units were not mined at the start of the sample period.

<sup>25</sup>In the appendix, we show that results are similar if we estimate the baseline model using *only* either the “never-treated” or the “not-yet-treated” units as controls.



obstacles that landmines pose on mobility and agricultural investments, and the strategic use of landmines to protect illicit economies. The main outcomes are: nighttime lights (Column 1) and population density (Column 2), which are proxies of economic activity; the probability that students obtain a satisfactory grade in math (Column 3) and reading (Column 4) standardized national test scores; forest loss (Column 5) and the size of coca crops (Column 6).

Recall that we estimate the causal effect of all demining treatments using Callaway and Sant’Anna (2020)’s procedure, and report group-time aggregate ATTs together with their respective standard errors, clustered at the event (buffer) level.<sup>26</sup> We also present p-values that control for the family-wise error rate in multiple hypotheses testing following Romano and Wolf (2005). Table 1 reports the baseline results estimated for buffers of 5Km radius around the geo-located event, and for a three-year window around the event date.

The table includes four panels with the objective of exploring the robustness of the estimated impact of humanitarian demining. Panel A is the baseline specification with no controls. Panel B adds buffer-level geographic covariates in a doubly robust way, following Sant’Anna and Zhao (2020). This procedure allows the specification to be robust to either the misspecification of the kernel-based difference-in-differences estimator that includes covariates in a flexible way (Heckman et al., 1997), or misspecification of the inverse probability weighted estimator (Abadie, 2005).<sup>27</sup> Also following the doubly-robust procedure, Panel C includes municipal-level covariates, one of them being the number of previous victims that was used for prioritization.<sup>28</sup> Finally, Panel D residualizes the outcomes from municipality-specific linear trends. Following Borusyak et al. (2021), we estimate the municipality-level trends using the untreated observations. Our baseline estimates are robust to these alternative specifications in terms of both magnitude and statistical significance.

Regarding the effect of humanitarian demining on nighttime luminosity, Column 1 suggests

<sup>26</sup>Appendix Table A3 reports the robustness to using two alternative clustering levels, one at the level of sub-municipal hamlets (called “vereda”) and the other one at the level of 15x15km grids. Aside from somewhat more imprecise estimates for the effect of post-conflict demining on population density, the results are very similar. Note that, unfortunately, no Conley-type standard errors that account for spatial correlation have been developed for these types of models. However, in an additional robustness test, we follow Bauman et al. (2018) and include the Moran eigenvectors as covariates to remove the spatial autocorrelation from the residuals. The results are reported in Appendix Table A4.

<sup>27</sup>The set of characteristics includes buffer-level temperature, precipitation, altitude, distance to the closest river, and distance to the closest National Park.

<sup>28</sup>The set broader of municipal characteristics includes the number of landmine victims, population, a coca suitability index, distance to the country’s capital, a rurality index, exposure to FARC violence, and a poverty index, all of them measured at the beginning of our sample period.

that, relative to the control buffers that are not-yet-treated or never treated, nightlights increase by 12.4% on average, in the three years after a demining event.<sup>29</sup> This effect is about a third of the one found by [Chiovelli et al. \(2019\)](#) for the case of demining in Mozambique, namely a 37.3% increase in luminosity after a locality is cleared from landmines.<sup>30</sup> Consistent with the finding reported in Column 1, Column 2 shows that population density also increases after demining. It does so by 2.7% when compared to the sample average.

How does the increase in nighttime light density, triggered by humanitarian efforts to clear landmines, translate into more traditional metrics of economic performance? We answer this question by computing the share of the municipal area affected by 5Km-radius buffers around demining events in the median municipality and multiplying it by the estimated average surge in nighttime lights as reported in Column 1. We then take the product of the resulting number and the median elasticity of GDP to nighttime luminosity, as estimated by [Henderson et al. \(2011\)](#) ( $= 0.3$ ). This back-of-the-envelope calculation suggests that a humanitarian demining event increases the municipal GDP by 0.8% per year.

In turn, this figure can inform a cost-benefit analysis in which we compare the median municipal value-added to the cost of humanitarian demining per square meter and the size of the average demined area.<sup>31</sup> Following this procedure, we find that the benefit/cost ratio is 7.1. That is, humanitarian demining increased income by over seven dollars per invested dollar. This is likely a lower bound as it considers only the benefits realized the year after the humanitarian demining takes place, and does not take into account the learning gains that we document next.

The positive effects of humanitarian demining are not limited to economic activity. For instance, we also find that students' performance in national standardized tests improves after demining events that take place in the vicinity of the school. In particular, we find an increase in the share of students with satisfactory performance in the math (reading) test of 6.7 (8.1) percentage points (Columns 3 and 4 of Table 1, respectively). This increase is

<sup>29</sup>As suggested by [Bellemare and Wichman \(2020\)](#), we compute the percentage change in the outcomes subject to a hyperbolic sine transformation as  $e^{\hat{\beta}} - 1$ .

<sup>30</sup>One potential driver of this increase in nighttime lights can be an increase in electrification after humanitarian demining. Using municipality-level data, we estimate a difference-in-differences model for municipalities with humanitarian demining before and after 2013. We find a positive non-statistically significant increase in the number of electricity subscribers per capita (see Appendix Table A5). We use this point estimate, 0.005, and the relationship between nighttime lights and electricity subscribers, 1.293, and find that the increase in subscribers can only explain at most 5.1% ( $= 0.005 * 1.293 / 0.127$ ) of the estimated effect of humanitarian demining on nighttime lights.

<sup>31</sup>We use the municipal value-added since Colombia has no official GDP statistics at the municipality level. However, the correlation between these two variables at the department level (the smallest administrative unit for which GDP figures are available) is 0.81, and it is strongly significant. We obtained the median cost of demining per square meter –COP 66,700 (= \$ 18)–from [Mutual-Co \(2021\)](#). The actual cost, however, varies substantially depending on how isolated are the areas.

statistically significant and its magnitude is large: it translates to a 32% (36%) increase in the probability of getting a satisfactory grade in math (reading) relative to the sample mean. In Table A6 of the Appendix, we present the results of the effects of demining on grade-specific test scores. Interestingly, the magnitude of the effect is larger for younger students, especially for the math test. In the next section, we explore the potential mechanisms of the positive effect of humanitarian demining on students' performance. In particular, we study the channels related both to changes in the composition of the student population and to learning.

Humanitarian demining also reduced forest loss, albeit by a small and non-significant magnitude (Column 5). In contrast, the reduction that it caused in the size of illegal coca crops, the first activity in the chain of cocaine traffic to the US and other consumption destinations, is large and significant. As shown in Column 6, after a demining event in the vicinity, the area cultivated with coca decreased by 10.4%.

**5.2 Demining in post-conflict military operations** Table 2 follows the same structure as Table 1 to study the effect of demining events that result from military operations carried out during the post-conflict period on the same set of outcomes. This demining treatment has no robust effect on nighttime light density (Column 1). It is positive and significant only when covariates are added in a doubly robust way in Panels B and C (Sant'Anna and Zhao, 2020). When no covariates are added or when the outcome is residualized from municipal-specific trends, the point estimates are nearly zero. We conclude that demining in military operations after the end of conflict does not affect nighttime luminosity in a robust way. In contrast, it does significantly decrease population density by 2.9% when compared with the sample average (Column 2).

Demining activity in military operations during the post-conflict period has no robust effect on students' performance in math test scores (Column 3) and seems to reduce the performance in reading test scores by a small magnitude (Column 4). However, this is not robust to the inclusion of buffer-specific geographic covariates in a doubly-robust fashion (Panel B). Nonetheless, this demining treatment does increase deforestation in treated buffers in a magnitude equivalent to 29.4% (Column 5), and it also reduces coca crops by 9.2% (Column 6).

The different effects of humanitarian and military demining can be explained by the different nature of the treatments (described in section 2), or by differences in the areas targeted by one type of treatment or the other. In an attempt to clean the latter (selection) channel and thus make the two treatments more comparable, we perform two tests aimed at increasing

the overlap in the characteristics of the areas subject to humanitarian and military demining. First, we estimate a LASSO model following Belloni et al. (2014) where the dependent variable takes the value of one if the area was subject to humanitarian demining.<sup>32</sup> Once we uncover the variables that best predict the treatment, we estimate the propensity score and follow Crump et al. (2009) to truncate the sample to increase overlap and re-estimate our baseline model on the common support. Second, we select groups composed by three to ten covariates based on all possible combinations to construct the propensity score, then truncate the sample and re-estimate the model within each underlying common support, thus obtaining a distribution of treatment effects.

In Table A7, we present the estimates from the first approach. We find that humanitarian demining increases nighttime lights in a magnitude that is twice as large as the increase experienced after military demining.<sup>33</sup> Similar patterns emerge for population density and students' performance, but for deforestation and coca cultivation, the effects are larger in the case of military demining. Figure A7 reports the results obtained with the second approach, which leads to similar conclusions.

**5.3 Demining in military operations during conflict** Finally, Table 3 repeats the same analysis to study the effect of demining events that result from military operations carried out over the course of the conflict. For this period (2004-2012), we have no data on the performance of students in standardized test scores, so we focus on the other four outcomes. The first finding, which is robust in terms of magnitude and significance to the addition of controls (Panels B and C) and municipality-specific trends (Panel D), is that demining events during conflict *decrease* economic activity, as measured both in terms of nighttime light density and population density (Columns 1 and 2, respectively). In terms of magnitudes, demining events that take place during the conflict decrease average nighttime luminosity by 1.3% and population density by 1.8% relative to the sample mean. The second finding is that demining in military operations during conflict increases deforestation by 10.2%, but has no impact on coca-growing.

In short, we find that post-conflict humanitarian demining increases –in the small buffers around the events–both economic activity and the performance of students in standardized test scores. It also reduces coca-growing but has no effect on forest loss. In contrast, demining in military operations that took place during the post-conflict period did increase deforestation with no effect on either nighttime lights or students' performance. It also decreased

<sup>32</sup>The list of right-hand-side variables includes: a poverty index, the logarithm of population, the distance to the closest department's capital, a coca suitability index, the distance to the country's capital, the distance to the closest national park, a rurality index, elevation, precipitation, and temperature.

<sup>33</sup>The standard errors of the common support sub-sample are, however, somewhat large and thus the difference is not statistically significant at conventional levels.

coca-growing. Finally, demining in military operations during the conflict period decreased economic activity and increased forest loss. Importantly, the credibility of these estimates depends to a large extent on the validity of the methodology’s identifying assumption. We discuss this in the next subsection.

**5.4 Main identifying assumption** The main assumption of the validity of Callaway and Sant’Anna (2020)’s approach to identifying causal effects is that, in the absence of the treatment, the evolution of the potential outcomes would be the same for the treated cohort ( $g$ ) and the never-treated or soon-to-be-treated units. To partially assess the validity of this assumption, we present the event-study version of the estimated ATTs, aggregated according to the relative time to the demining event.

Figure 2 reports Callaway and Sant’Anna (2020)’s event study of the effect of humanitarian demining on the outcomes of interest.<sup>34</sup> We find that, before the treatment, the coefficients tend to move around zero and show no discernible differential pre-treatment trend. This is particularly so for nighttime lights (Panel A) and students’ performance (Panels C and D).<sup>35</sup> There seems to be a drop in (relative) year -1 in the differential level of coca-growing (Panel F), which leads us to interpret the findings associated to this outcome with caution. Likewise, Figure 3 suggests that, for the case of demining in military operations after the end of the conflict, most of the outcomes lack differential pre-trends. The exception is forest loss, for which the differential trend is slightly decreasing prior to the demining event. Finally, Figure 4 shows that in the case of demining in military operations carried out over the course of the conflict, there are no differential pre-treatment patterns for any of the outcomes.

We complement the event study figures with a formal test of whether the pre-treatment trends are parallel. Following Roth (2021), we use the precision of our estimates in the pre-treatment period to compute the pre-trend that has a 50% power of being detected, as well as the adjusted pre-trend that takes into account the pre-testing bias that arises from the fact that the reported analysis is conditional on passing a pre-test. We report the average biases in Table A8 of the Appendix.<sup>36</sup> The results suggest that, for the case of the humanitarian demining treatment, the size of bias is similar to the magnitude of the estimates for population density and coca cultivation (see Panel A). This suggests that the finding that

<sup>34</sup>Note that for the estimator suggested by Callaway and Sant’Anna (2020), there is no need for omitting relative year -1 as in the usual difference-in-differences regression with two-way fixed effects. This is because the point estimates of the pre-treatment period are weighted averages of differences between year  $t$  and  $t - 1$  across treated units and their relevant comparison group.

<sup>35</sup>The immediate surge in students’ performance cannot only be explained by the effect of humanitarian demining of students’ learning. In the next section, we posit that mine clearance generates migration dynamics that likely explain this effect through changes in the composition of students in treated schools.

<sup>36</sup>This is the average of the hypothesized trend that goes from (relative) year 0 to year 3, as well as the average of the pre-testing bias-adjusted trend.

humanitarian demining increases population density and decreases coca cultivation should be interpreted with caution.

Finally, we follow [Rambachan and Roth \(2021\)](#) and estimate the 90% confidence set for our parameters of interest after allowing for linear and non-linear deviations from the parallel trends assumption. We estimate such confidence set for the reported coefficient of the year after the demining event. In the case of non-linear deviations, we allow the change in the trend from consecutive periods to be as large as the size of the pre-trend that has a 50% power of being detected given the precision of the estimates in the pre-treatment period (as in [Roth, 2021](#)). Figures [A3](#) to [A5](#) report the confidence sets resulting for the three treatments. In most of the cases, we find significant results even after allowing for a linear deviation of the parallel trends assumption ( $M = 0$ ). When we allow for non-linear deviations –i.e., the trend can change size and sign for consecutive periods ( $M > 0$ )–we find that the increase in nighttime lights and students’ test performance after humanitarian demining is robust. In the case of demining resulting from post-conflict military operations, we find that significant baseline estimates are robust to both linear and non-linear violations, with the exception of coca cultivation which is only robust to the former. Finally, for demining resulting from military operations during the conflict, the reported increase in forest loss and population density are robust to both linear and moderate non-linear violations.

**5.5 Robustness** Our results are robust to a battery of additional tests, that we present and thoroughly discuss in [Appendix B](#). These include using other estimation methods such as those proposed by [Borusyak et al. \(2021\)](#), [De Chaisemartin and d’Haultfoeuille \(2020\)](#), and [Wooldridge \(2021\)](#); allowing for spillover effects; accounting for potential anticipation effects by excluding from the comparison group the observations that occur one year prior to the demining events; using alternative comparison groups such as the ‘not-yet-treated’ or the ‘never treated’ only; using alternative buffer sizes of 3, 4, 6, and 7Km; excluding demining events in the proximity of areas of interest; and removing outliers.

## 6 MECHANISMS

In this section, we explore the potential mechanisms behind the heterogeneous local effects of different demining treatments during conflict and in the post-conflict period. In particular, we study: i) the role of local road connectivity and market access in promoting economic activity and school performance after humanitarian demining events; ii) the role of composition and learning channels in explaining the positive effect of humanitarian demining on students’ performance; iii) the effect of demining on the dynamics of conflict and how it may reduce economic activity after military demining prior to the end of the conflict; iv) the differential effects of demining on deforestation in areas with different types of soil suitability;

v) the complementarity of the demining efforts with other policies that seek to promote rural development, particularly with a recent illegal crops substitution program.

**6.1 Road connectivity and market access** We start by exploring potential heterogeneous effects of the documented impact of humanitarian demining efforts on nighttime light density and on students' performance. The potential economic benefits of demining are likely exacerbated if the clearance takes place in areas that are more connected to local markets through a network of roads. Once the mobility restrictions that landmines impose are lifted, better access to inputs, markets for goods and services, and labor opportunities, result in a faster and higher pick up of key economic and social activities. We explore this hypothesis with two different but complementary approaches. First, we use a reduced-form procedure that exploits a rich network of geo-located paved and unpaved roads, available for the entire country and measured in 2012. Second, we leverage on a market access general equilibrium framework following the original contribution of [Donaldson and Hornbeck \(2016\)](#) as well to the application of [Chiovelli et al. \(2019\)](#) to a setting similar to ours. This allows us to estimate the economy-wide effects of humanitarian demining under the assumption that it lifts all mobility restrictions across the previously mined roads.

For the reduced-form approach, we follow the strategy proposed by [Marcus and Sant'Anna \(2021\)](#) to estimate heterogeneous effects in settings of difference-in-differences with staggered adoption. We thus re-estimate our baseline specification on two mutually exclusive samples of demining events: those that occurred in more connected areas and those that took place in less connected places. To that end, we use two different measures. The first one exploits the extensive margin of connectivity and looks at demining events in areas with at least one (paved or unpaved) road that passes within 100 meters of the demined area. The second one exploits the intensive margin, as parametrized by the length of all roads that cross within 100 meters of the demined area. In this case, we use the median of the empirical road length distribution to separate places with high and low connectivity.

Table 4 reports the results from this exercise. Based on the extensive margin of connectivity, we find that the increase in nighttime lights following a demining event is 18.5% in areas with at least one road close to the centroid (Panel A, Column 1). In contrast, the effect of demining of nighttime light density is less than half in areas with no road nearby (Panel A, Column 2). This difference is statistically significant at the 5% level. A similar pattern is found for students' performance, with the positive effect of demining being larger in more connected areas, even though the difference between the two samples is not statistically significant at conventional levels (Columns 1 and 2 of Panels B and C for math and reading test scores respectively).

The results are quite similar when we exploit the intensive margin of road connectivity. The effect of demining on both nighttime lights and students' performance is larger in areas better connected to markets (Columns 4 and 5 of Table 4). In this case, the effects of demining on the performance of students are significantly larger in relatively more connected areas, which is consistent with the interpretation that the risk of a landmines explosion prevents children from going to school. Indeed, the available anecdotal evidence suggests that, in several parts of Colombia, landmines are an important obstacle for accessing schools (CNMH, 2017).<sup>37</sup>

Regarding the market access framework, we estimate Donaldson and Hornbeck (2016)'s log-linear reduced-form relationship between aggregate welfare and market access.<sup>38</sup> The purpose of implementing this model is to understand whether there are general equilibrium effects from humanitarian demining that increase connectivity and potential trade across regions. Appendix C summarizes our empirical approach.

Table 5 summarizes the results. Columns 1 to 3 include the log of market access as the main regressor, while columns 4 to 6 include the cell-specific cumulative number of demining events instead. This allows us to study the direct effects of demining. Columns 7 to 9 include both variables. Columns, 1, 4, and 7 only control for cell and year fixed effects. Columns 2, 5, and 8 add the department-year fixed effects. Finally, columns 3, 6, and 9 control for spatial correlation that changes over time by adding a cubic polynomial in latitude and longitude interacted with the time fixed effects (Donaldson and Hornbeck, 2016). Our preferred specification (column 3) implies a luminosity-to-market-access elasticity of 0.25. That is, a one-standard-deviation increase in market access is associated with a 25 percent increase in nightlights density. Importantly, the magnitude of the estimates elasticity is quite similar to the 0.3 found by Donaldson and Hornbeck (2016) and equal to the 0.25 of Chiovelli et al. (2019).

When adding cumulative demining as our treatment variable, we find results that are consistent to those obtained from our main specification, that uses as a comparison only places that are known to host landmines. Specifically, and focusing on column 6, we find that moving from no demining to the mean of the cumulative mine clearance (3.5) increases nighttime lights by 7 percent. Finally, when we add both the log market access and cumulative demining as dependent variables, we find that the elasticity of nighttime lights to market access decreases to 0.21 (column 9). However, the estimate of the effect of cumulative demining remains largely unchanged. This suggests that humanitarian demining efforts yield both

<sup>37</sup>In Appendix Figure A8, we show that more connected buffers tend to be similar to less connected ones. Thus, alleviating concerns that this heterogeneity analysis is capturing other characteristics such as population density or agricultural suitability.

<sup>38</sup>See Donaldson and Hornbeck (2016) for details about the model's assumptions and the derivation of the reduced-form relationship.



direct and spillover effects on economic activity.

Finally, we leverage on the estimated elasticity of nightlights to market access to conduct a counterfactual exercise similar to that proposed by [Chiovelli et al. \(2019\)](#). Specifically, we compare Colombia's current GDP (with the demining policy actually implemented) with the figure that would result from a counterfactual scenario with no humanitarian demining. For the latter, we use a market access measure based on the 2012 road network (that includes all planted landmines) and we keep fixed the 2012 distribution of nighttime lights to attribute their 2019 counterpart. We then compare the resulting counterfactual values with the actual 2019 market access values that were shaped by the observed demining. Specifically, for this comparison, we sum across localities the exponential difference of the factual and counterfactual market access values multiplied by our estimated nightlight-to-access elasticity of 0.25 (see column 3 of Table 5). We obtain that, absent any humanitarian effort, nighttime luminosity would have been 17 percent lower between 2013 and 2019. Moreover, using the GDP to nighttime lights elasticity of 0.3 (computed by [Henderson et al., 2011](#)), this implies that, over that period, Colombia grew an additional 5.1 percent thanks to humanitarian demining. This equates to a yearly average additional growth of 0.7 percent. Interestingly, this estimate is remarkably similar to that suggested by [Chiovelli et al. \(2019\)](#) for Mozambique (which lies in the range of 0.7 to 1 percent per year).

**6.2 Students' composition and learning** The positive effect of humanitarian demining on students' performance in math and reading tests, which takes place across different school grades is in line with recent findings that document the effect of Colombia's recent peace agreement on educational outcomes, which improved more than proportionally in areas that experienced landmine explosions before the permanent ceasefire that preceded the agreement ([Prem et al., 2021b](#)). In our setting, the progress of students' performance could potentially be accounted by mechanisms pertaining to both school's composition and students' learning. In particular, because improvements in learning likely take time to materialize into better performance, the immediate effect of humanitarian demining on the performance of students in math and reading tests (reported in Panels C and D of Figure 2) is presumably also driven by a school composition effect. Consistent with this idea, recall that humanitarian demining increases population density (see Table 2, Column 2), which implies that the surges in safety that demined areas experience attract more people. This is in turn probably driven by formerly displaced households who return to their land, as evidenced by the fact that cleared areas experience an increase in land restitution requests followed by the Land Restitution Unit.<sup>39</sup>

<sup>39</sup>Results available upon request.

The incoming migration changed schools' composition. For instance, in Table 6 we document that humanitarian demining increased school enrollment by 34% (Column 1).<sup>40</sup> This translates into a differential increase in the number of students taking the standardized test (Column 2). Moreover, humanitarian demining also increased by 4.4 percentage points the probability that a new school opened in the treated buffer (Column 4).

Using the estimated increase in enrollment, we can see how much of the estimated change in the share of students with satisfactory performance is driven by a composition effect. If we assume that all the newly enrolled students pass the test, the increase in enrollment would explain 49% and 37% of the effects that we find in math and reading tests.

Regarding the effect of demining on the production function of students' learning, we find that, because of the arrival to additional teachers, the students-to-teachers ratio decreased, effectively reducing class size (Column 3 of Table 6). Moreover, the share of approved students in the standardized national tests also increased differentially (Column 5). Finally, the fact that the increased performance of students is larger in the demined areas that are closer to the road network is consistent with a reduction in school absenteeism, which also likely generates learning improvements.

Figure A9 shows the event-study counterparts of the outcomes reported in Table 6. The immediate (contemporaneous to the demining year) increase in enrollment and decrease in the students-to-teachers ratio are consistent with the complementarity of composition and learning mechanisms explaining the effects of humanitarian demining on students performance.

**6.3 Conflict and demining** While the availability and the density of the road network can, at least partially, account for the positive effects of post-conflict humanitarian demining, the negative effect of demining in military operations during the conflict in terms of declining nighttime lights and population density is perhaps more puzzling. We posit that, when the conflict was fully active, demining during military operations likely exacerbated violent dynamics of territorial contestation by illegal armed groups. Due to the absence of geo-referenced data on the incidence and intensity of violence, we explore this channel indirectly by studying the municipal-level correlation between the different types of demining and variables related to the incidence of conflict-related violence and forced internal displacement. To this end, we estimate a municipal panel specification of the effect of demining on the

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<sup>40</sup>Though not reported, the additional enrollment is driven by that taking place in the (mandatory) elementary and middle school grades, which include the grades in which the standardized national tests studied in this paper are implemented (third, fifth, and ninth). Instead, we see no differential increase in the enrollment in neither preschool nor high-school.

number of violent attacks and victims of forced displacement,  $y$ . That is, we estimate:

$$(6.1) \quad y_{mdt} = \beta_1 \times \text{Conflict}_t \times \text{Military Demining}_{mt} + \beta_2 \times \text{Military Demining}_{mt} \\ + \beta_3 \times \text{Humanitarian Demining}_{mt} + \alpha_m + \alpha_{dt} + \sum_{c \in \mathbf{X}_m} \gamma'(c \times \delta_t) + \epsilon_{mdt},$$

where  $m$ ,  $d$ , and  $t$  stand for municipality, department, and time (year), respectively. We estimate this regression model over the entire sample period (2004-2019), and define  $\text{Conflict}_t$  as a time dummy that takes the value one from the beginning of the sample period until just before the start of the peace negotiation with FARC (2012). Both our violence-related dependent variables and the right-hand-side military and humanitarian demining treatments are hyperbolic sine transformations of the variables in levels. We include municipality and department-year fixed effects as well as flexible trends parametrized by municipality characteristics measured before the beginning of our sample period.<sup>41</sup> The parameters  $\beta_2$  and  $\beta_3$  pick up the correlation between, respectively, military demining and humanitarian demining, on the incidence of violence during the post-conflict period (2013 onward). Instead,  $\beta_1$  picks up the differential correlation between military demining and violence during the conflict period.<sup>42</sup>

Table A9 reports the results from estimating equation (6.1). We find a positive correlation between demining in military operations and violence during the post-conflict period. However, the magnitude of this relationship more than doubles during the conflict period. In contrast, the relationship between (post-conflict only) humanitarian demining and violence is *negative* (Columns 1 and 2). Moreover, we find that there is a negative association between demining in military operations and attacks by both FARC and paramilitary groups during the post-conflict period, but that this correlation turns positive and large during the conflict period. The relationship between humanitarian demining and bellicose activity in the form of attacks is negative, though only statistically significant for the case of attacks perpetrated by FARC (Columns 3 and 4).

We interpret these results as aligned with the idea that (mostly illegal) armed groups use landmines to prevent the territorial advancement of enemies (Fundación Seguridad y Democracia, 2006), and therefore demining amidst the conflict triggers violent territorial contestation between illegal groups, including the victimization of civilians thought to collaborate with the enemy (CNMH, 2017; Procuraduría, 2011). Importantly, by highlighting a potential demining-driven violence surge, these results are very much consistent with the documented

<sup>41</sup>These include total population, a coca suitability index, distance to country's capital, a rurality index, and a poverty index.

<sup>42</sup>Note that in equation (6.1) we do not interact *Humanitarian Demining* with the conflict period dummy since by and large this type of demining took place during the post-conflict period.

decrease in nighttime luminosity and population density that partially demined areas experience along the conflict sample period.

**6.4 Deforestation, soil suitability, and extractive activities** To shed light on relevant underlying mechanisms behind our results regarding the effects of demining on deforestation, we exploit rich information about soil suitability at the level of 30m<sup>2</sup> grids. Based on that input, we build average suitability measures within the 5Km buffer around all the demining events of our sample. We then split the sample of demining events into those that took place in areas highly suitable for specific land use, as well as those that took place in low-suitability areas. We do so based on the empirical distribution of buffer-specific average suitability.

With this input, we explore the potential heterogeneous effects that demining in military operations have on deforestation according to the extent to which the soil is suitable to specific extractive activities such as oil palm crops, cattle herding, and rubber crops (Indepaz, 2008; Indepaz, 2020). Table 7 reports the results from this analysis. The results are compelling in suggesting that the effect of demining on deforestation is driven by its occurrence in areas highly suitable to extractive activities. For instance, Panel A suggests that the documented post-military demining forest loss (both during conflict and during peace) is larger in areas more suitable for oil palm. This increase is 21% (60%) in high suitable areas during conflict (peace) with no effects for low suitable areas. Similar figures are found when looking at the effect of military demining on forest loss in areas suitable to cattle herding (Panels B and C): military demining causes a large forest loss in high cattle-suitable areas, with the effect being around 17% (56%) for military demining during conflict (peace). We find similar stories for rubber and forestry suitability.<sup>43</sup> Importantly, the suitability of the land to extractive activities does not exacerbate or attenuate the effects of humanitarian demining on deforestation (Columns 1 and 2), and we do not find any differential forest loss patterns after military demining in areas suitable to non-extractive traditional crops such as rice, maize, and potato (Panel F).

To complement the idea that the change in forest loss can be related to an increase in extractive agricultural activities after military demining, we explore the effect of demining on the incidence of wildfires.<sup>44</sup> Anecdotal evidence has shown that, in the Colombian context, fires

<sup>43</sup>The empirical relevance of this mechanism is consistent with the recent findings of Prem et al. (2020), who document a differential increase in deforestation –most likely related to extractive activities–after the start of the ceasefire in municipalities previously exposed to FARC violence.

<sup>44</sup>This is similar to the approach followed by Harding et al. (2021) to study how deforestation is likely to be related to extractive illegal activities.

are used to clear forests for cattle ranching and other land-intensive agricultural activities.<sup>45</sup> Table A10 documents that demining in military operations during the conflict (post-conflict) period caused fires to increase by 5% (3.5%). This is consistent with the increase in forest loss that we documented in our baseline analysis (see Figure A10 for the event study estimates).

We also explore the effect of demining on illegal gold mining, a highly profitable extractive activity that has been widely used by illegal armed actors in Colombia to finance their operation (Idrobo et al., 2014). Since UNODC estimates of illegal gold mining are only available since 2014, we can only do the analysis for demining events during the post-conflict period. Table A11 shows that humanitarian demining events have no effect on illegal gold mining, neither in the extensive nor in the intensive margin (Columns 1 and 2). In contrast, demining activity in the context of military operations carried out during this period increased both the incidence of illegal gold mining and the extension of this activity (Columns 3 and 4). See Figure A11 for the event study estimates. Overall, the results highlighted in this subsection are consistent with the extensive qualitative evidence suggesting that in the Colombian context, the new rural investors that demining attracts may be associated with illegal armed groups, and particularly paramilitary militias.

**6.5 Government programs and coca cultivation** Finally, we exploit one of the main milestones of the 2016 peace agreement with FARC. The implementation of an ambitious illegal crops substitution program (PNIS from its Spanish acronym). This program was created in May 2017, 6 months after the peace agreement was ratified by the Colombian Congress. By 2018, it had reached almost 99,000 farmers in 56 municipalities, and two-thirds of them had received payments (Garzón et al., 2019) as a reward for having successfully eradicated illegal coca crops and replaced them with a legal alternative. Given the relevance of this program for the rural development prospects of the main cocaine exporter of the world, we study whether the documented effect of demining during peace on the level of coca cultivation can be at least partially explained by the national roll out of the PNIS crop substitution program. To this end, we split our sample for those cohorts treated after 2017 (when PNIS was launched) into the areas where PNIS was present and those with no active crop-substitution policy. Table A12 summarizes the results. We find that peace-time demining decreases coca crops, especially on PNIS-targeted areas. Moreover, the complementarity effect between demining and PNIS is over three times larger for humanitarian demining than for demining in military operations.

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<sup>45</sup>See for example, <https://news.mongabay.com/2019/09/as-the-amazon-burns-colombias-forests-decimated-for-cattle-and-coca/> and <https://theecologist.org/2020/aug/17/deforestation-colombia> (last accessed 10/9/2021).

## 7 CONCLUSIONS

In spite of the tens of millions of planted anti-personnel landmines that persist today worldwide, the enormous stock of manufactured but not yet planted landmines, and the thousands of landmine victims every year, the literature on the economic costs of conflict has surprisingly relegated the study of the long term economic and social consequences of landmines, as well as that of the potential benefits of demining campaigns. While recent efforts have highlighted that comprehensive landmine clearance operations result in increased economic activity, we know little about its impact on other socio-economic and political outcomes and about the effects of other types of demining, especially the kind that occurs as a result of military operations, both over the course of a conflict and during the post-conflict period. This paper contributes to filling these gaps.

We study the case of Colombia, the country with the highest number of casualties from improvised handmade landmines, and that has engaged in a range of demining activities since before the start of the peace process with FARC, Colombia's largest guerrilla group. Moreover, we focus on the local effects of demining by taking advantage of a unique dataset that includes the coordinates of both humanitarian demining campaigns and demining events resulting from military operations. Based on recent methodologies developed for difference-in-differences settings with staggered adoption, and exploiting the longitudinal variation of all demining events that took place from 2004 to 2019 in Colombia, we estimate the causal effect of demining on a range of outcomes, including nighttime light density, students' performance in standardized tests, population density, deforestation, and coca-growing.

Consistent with the previous literature, we find that comprehensive humanitarian demining events that take place after the end of the conflict increase economic activity. Moreover, we find that they also increase other variables associated with higher welfare, such as population density and students' performance in standardized tests. Importantly, all these effects are significantly larger in areas that are more connected to inputs, goods and services, and labor markets through a denser road network.

Quantitatively, based on both reduced-form and general equilibrium structural estimates, we find that humanitarian demining increases the annual GDP growth rate by around 0.7 percent, and that, consequently, each dollar invested in this policy yields about \$7 in return. Moreover, [Perilla et al. \(2021\)](#) estimate that, by saving lives, humanitarian demining yields an additional \$2.6 per dollar invested in targeted municipalities.

However, unlike any previous literature, we document that demining events that occur in the context of military operations are likely to backfire, especially if they take place while the conflict is still ongoing. Indeed, we find that demining in military operations increases

violent territorial contestation and, as a result, decreases economic activity and population density. It also increases deforestation rates, especially in areas that are suitable for extractive economic activities such as cattle ranching. This highlights the potential environmental costs of demining.

Altogether, our results highlight the fact that demining can, perhaps surprisingly, backfire. This suggests that in order to trigger beneficial economic and social dynamics demining campaigns should be both comprehensive (in terms of mines' clearance) and complemented with other state-building efforts and local investments.

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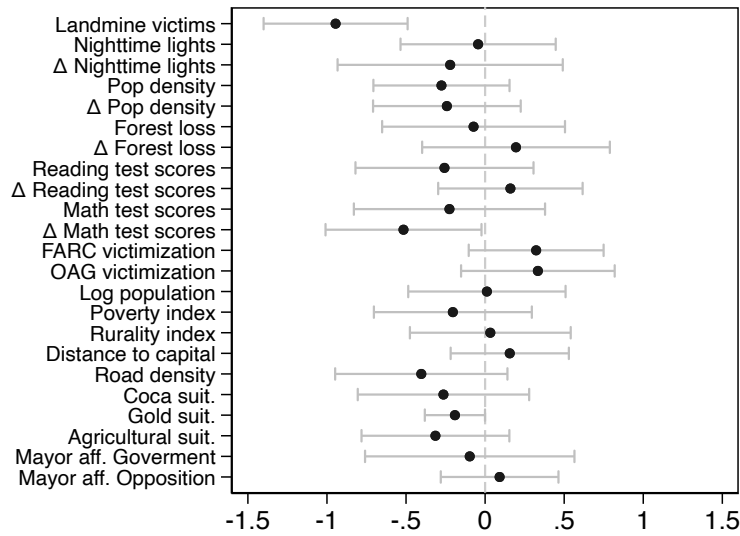


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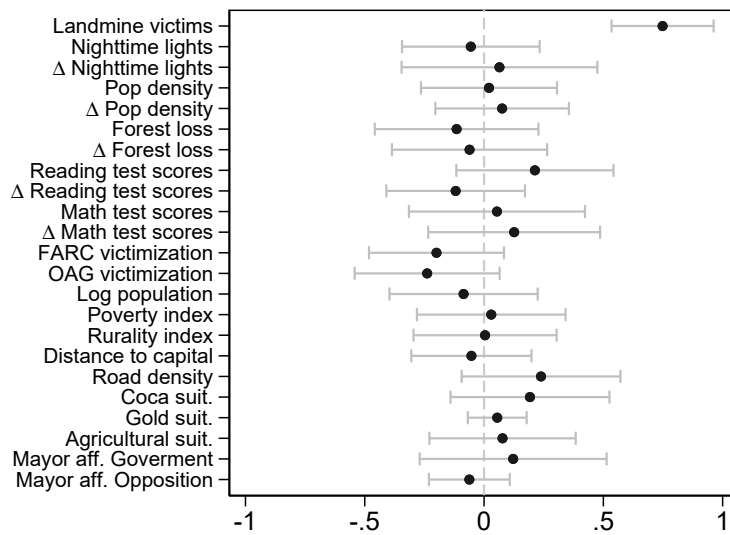
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FIGURE 1. Differential characteristics by timing and intensity of the treatment



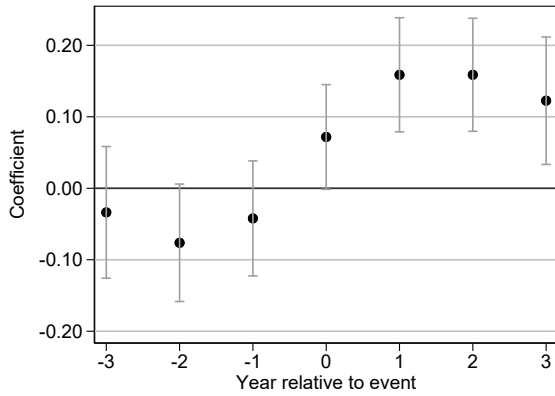
A. Humanitarian: Timing



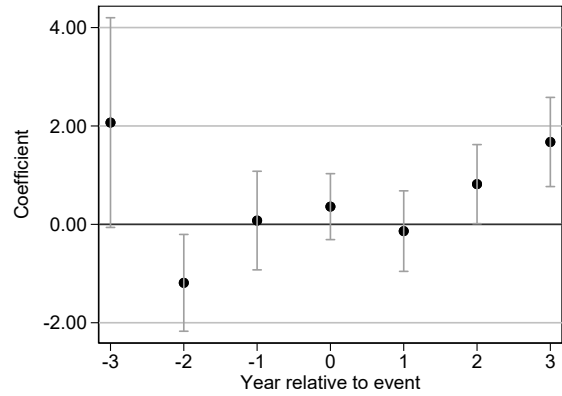
B. Humanitarian: Intensity

**Notes:** This figure presents the standardized differences by treatment timing and intensity. The sample is limited to municipalities that experienced any humanitarian demining since 2013. In Panel A, we present a point estimates and confidence intervals for a regression of the first year that a municipality experienced humanitarian demining on municipality level characteristics measured before 2013. In Panel B, we do the same but in this case the dependent variable is the logarithm of the total areas assigned for humanitarian demining in the municipality. Variables with a  $\Delta$  are first differences of the variable taking an average of two years before 2013.

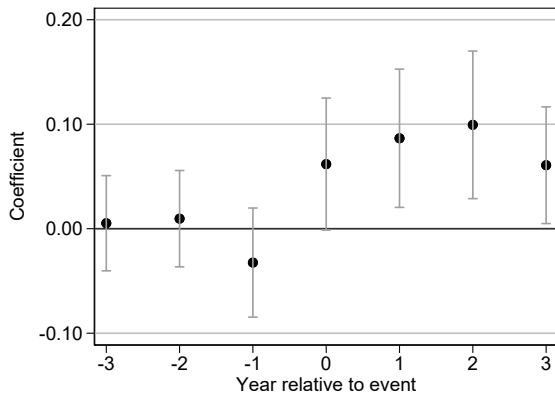
FIGURE 2. Humanitarian demining during peace and local activity



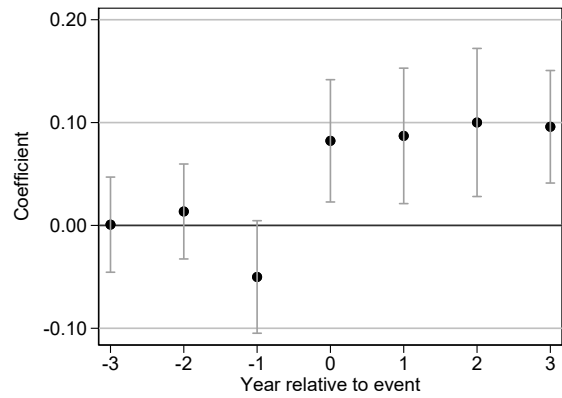
A. Nighttime lights



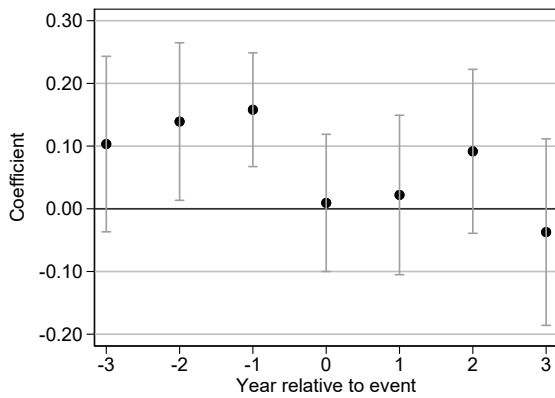
B. Population density



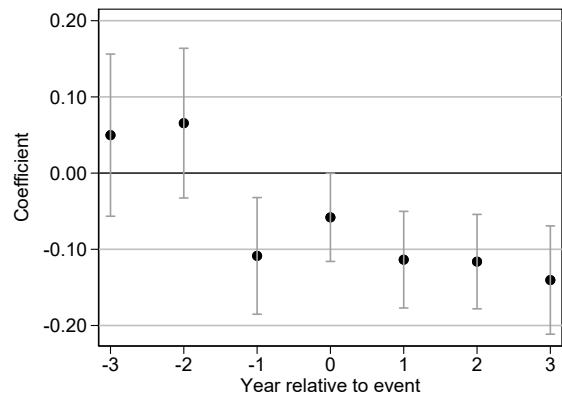
C. Test scores: Math



D. Test scores: Reading



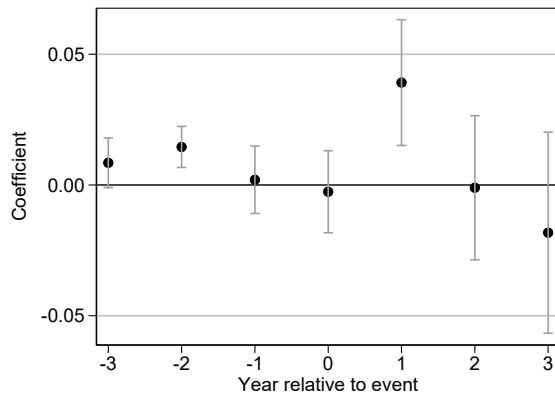
E. Forest loss



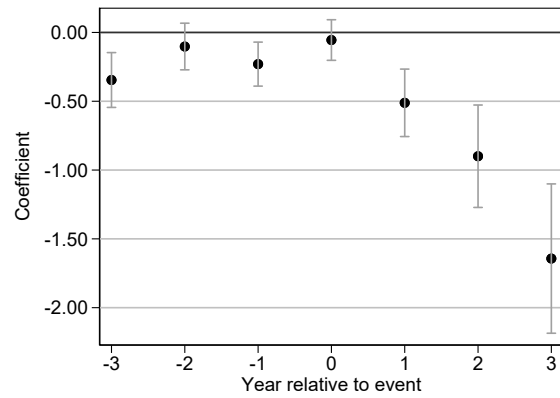
F. Coca

Notes: This figure presents the event study coefficients following Callaway and Sant'Anna (2020) for the treatment of humanitarian demining. We present the point estimates as well as the 95% confidence interval. Standard errors clustered at the event level. The outcomes were computed using a radius of 5km around the demining.

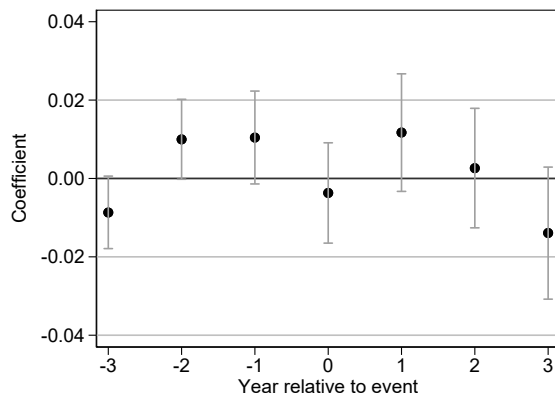
FIGURE 3. Military demining during peace and local activity



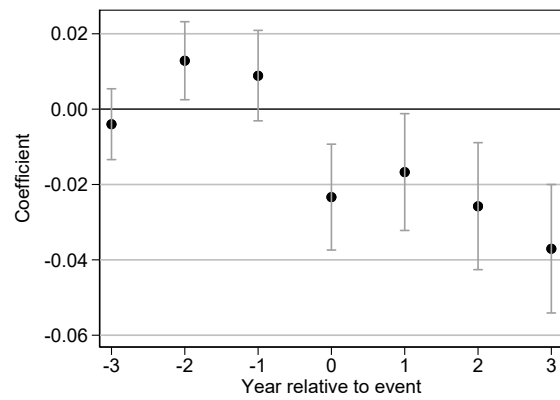
A. Nighttime lights



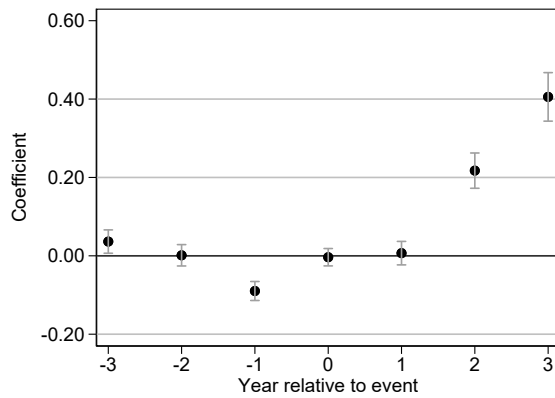
B. Population density



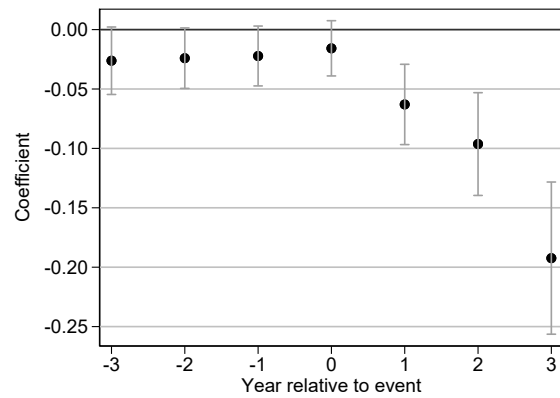
C. Test scores: Math



D. Test scores: Reading



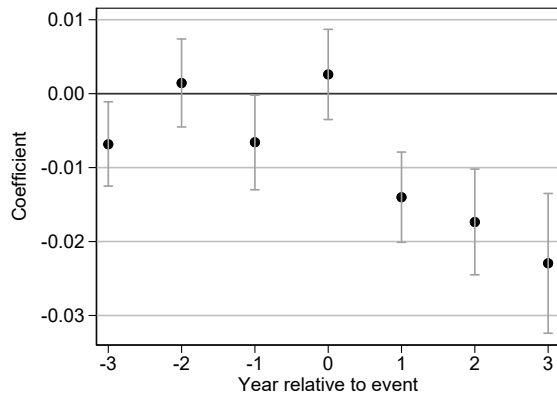
E. Forest loss



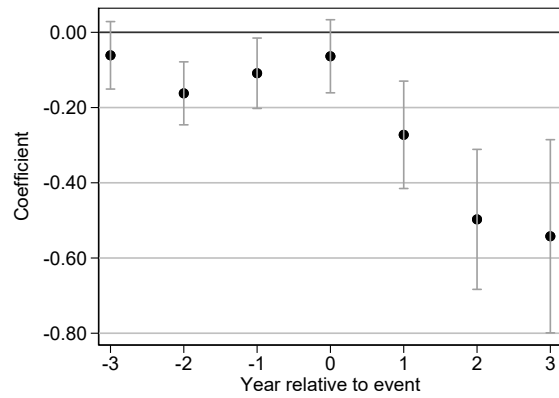
F. Coca

Notes: This figure presents the event study coefficients following Callaway and Sant'Anna (2020) for the treatment of demining during peace. We present the point estimates as well as the 95% confidence interval. Standard errors clustered at the event level. The outcomes were computed using a radius of 5km around the demining.

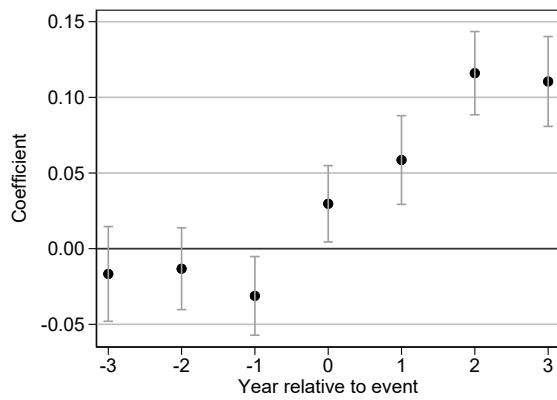
FIGURE 4. Military demining during conflict and local activity



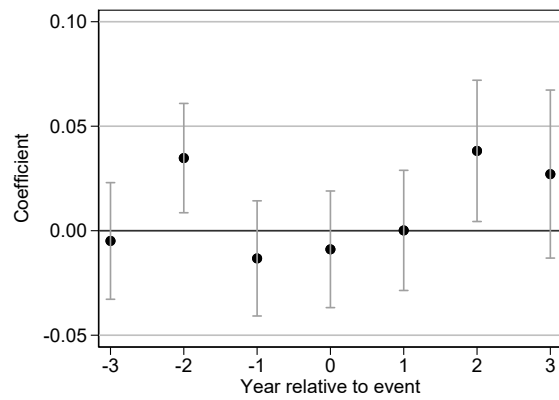
A. Nighttime lights



B. Population density



C. Forest loss



D. Coca

Notes: This figure presents the event study coefficients following Callaway and Sant'Anna (2020) for the treatment of demining during peace. We present the point estimates as well as the 95% confidence interval. Standard errors clustered at the event level. The outcomes were computed using a radius of 5km around the demining.

TABLE 1. The local effects of humanitarian demining during peace

	(1)	(2)	(3)	(4)	(5)	(6)
			Test scores:			
Dep. variable:	Nighttime Lights	Population density	Math	Reading	Forest loss	Coca
<b>Panel A: Baseline specification</b>						
Post demining	0.117*** (0.036)	0.938*** (0.363)	0.067*** (0.021)	0.081*** (0.021)	-0.031 (0.054)	-0.110*** (0.035)
<b>Panel B: Adds geographic covariates</b>						
Post demining	0.094** (0.046)	0.868** (0.352)	0.040* (0.021)	0.042** (0.020)	-0.078 (0.049)	-0.094** (0.037)
<b>Panel C: Adds municipality covariates</b>						
Post demining	0.106*** (0.040)	0.995*** (0.372)	0.051** (0.022)	0.065** (0.026)	-0.044 (0.067)	-0.063 (0.050)
<b>Panel D: Adds municipality linear trends</b>						
Post demining	0.108*** (0.036)	0.932*** (0.350)	0.067*** (0.022)	0.081*** (0.021)	-0.031 (0.052)	-0.109*** (0.036)
Observations	5983	7460	6960	6960	7460	7460
Treated	291	294	283	283	294	294
Never treated	449	452	413	413	452	452
Average dep var	1.675	34.407	0.207	0.222	3.290	0.636
MHT p-value	0.001	0.002	0.002	0.001	0.418	0.002

**Notes:** This table presents the overall ATT following Callaway and Sant'Anna (2020) for the treatment of humanitarian demining during peace. *Post demining* is the weighted average of all group-time average treatment effects with weights proportional to group size. The set of controls include not yet treated and never treated. The outcomes were computed using a radius of 5km around the demining. In Panels B and C, we use a doubly robust estimator following Sant'Anna and Zhao (2020). In Panel B, we use geographic covariates to predict the outcome. The set of covariates includes temperature, precipitation, altitude, distance to the closest river, and to the closest national park. In Panel C, we add a set of municipality characteristics as covariates that includes the logarithm of population, a coca suitability index, distance to the country's capital, a rurality index, a poverty index, number of landmine victims over population, and a dummy for FARC presence during 2007 - 2012. In Panel D, we residualized the outcome from municipality linear trends, that are computed using untreated observations. Bootstrap standard errors clustered at the event level. We also present p-values that control for the family-wise error rate in multiple hypotheses testing following Romano and Wolf (2005). \* is significant at the 10% level, \*\* is significant at the 5% level, \*\*\* is significant at the 1% level.



TABLE 2. The local effects of military demining during peace

	(1)	(2)	(3)	(4)	(5)	(6)
			Test scores:			
Dep. variable:	Nighttime Lights	Population density	Math	Reading	Forest loss	Coca
<b>Panel A: Baseline specification</b>						
Post demining	0.009 (0.013)	-0.983*** (0.220)	-0.001 (0.006)	-0.016** (0.006)	0.258*** (0.024)	-0.097*** (0.025)
<b>Panel B: Adds geographic covariates</b>						
Post demining	0.067*** (0.016)	-0.907** (0.375)	0.010 (0.009)	-0.011 (0.009)	0.217*** (0.026)	-0.093*** (0.030)
<b>Panel C: Adds municipality covariates</b>						
Post demining	0.067*** (0.019)	-1.054*** (0.379)	-0.003 (0.009)	-0.025*** (0.009)	0.080*** (0.024)	-0.117*** (0.032)
<b>Panel D: Adds municipality linear trends</b>						
Post demining	0.009 (0.013)	-0.979*** (0.227)	-0.001 (0.006)	-0.016*** (0.006)	0.258*** (0.024)	-0.094*** (0.025)
Observations	90504	100560	69340	69370	100560	100560
Treated	9630	9630	6641	6641	9630	9630
Never treated	426	426	293	296	426	426
Average dep var	0.973	33.381	0.149	0.168	3.721	2.392
MHT p-value	0.540	0.001	0.751	0.001	0.001	0.001

**Notes:** This table presents the overall ATT following Callaway and Sant'Anna (2020) for the treatment of military demining during peace. *Post demining* is the weighted average of all group-time average treatment effects with weights proportional to group size. The set of controls include not yet treated and never treated. The outcomes were computed using a radius of 5km around the demining. In Panels B and C, we use a doubly robust estimator following Sant'Anna and Zhao (2020). In Panel B, we use geographic covariates to predict the outcome. The set of covariates includes temperature, precipitation, altitude, distance to the closest river, and to the closest national park. In Panel C, we add a set of municipality characteristics as covariates that includes the logarithm of population, a coca suitability index, distance to the country's capital, a rurality index, a poverty index, number of landmine victims over population, and a dummy for FARC presence during 2007 - 2012. In Panel D, we residualized the outcome from municipality linear trends, that are computed using untreated observations. Bootstrap standard errors clustered at the event level. We also present p-values that control for the family-wise error rate in multiple hypotheses testing following Romano and Wolf (2005). \* is significant at the 10% level, \*\* is significant at the 5% level, \*\*\* is significant at the 1% level.

TABLE 3. The local effects of demining during conflict

	(1)	(2)	(3)	(4)
Dep. variable:	Nighttime Lights	Population density	Forest Loss	Coca
<b>Panel A: Baseline specification</b>				
Post demining	-0.013*** (0.003)	-0.478*** (0.089)	0.098*** (0.011)	-0.009 (0.013)
<b>Panel B: Add geographic covariates</b>				
Post demining	-0.013*** (0.003)	-0.463*** (0.096)	0.090*** (0.010)	-0.027** (0.013)
<b>Panel C: Adds municipality covariates</b>				
Post demining	-0.017*** (0.003)	-0.433*** (0.090)	0.110*** (0.010)	-0.021 (0.014)
<b>Panel D: Adds municipality linear trends</b>				
Post demining	-0.017*** (0.003)	-0.388*** (0.087)	0.103*** (0.011)	-0.009 (0.013)
Observations	213000	213000	213000	213000
Treated	15150	15150	15150	15150
Never treated	2600	2600	2600	2600
Average dep var	0.435	26.349	3.403	1.728
MHT p-value	0.001	0.001	0.001	0.361

**Notes:** This table presents the overall ATT following Callaway and Sant'Anna (2020) for the treatment of demining during conflict. *Post demining* is the weighted average of all group-time average treatment effects with weights proportional to group size. The set of controls include not yet treated and never treated. The outcomes were computed using a radius of 5km around the demining. In Panels B and C, we use a doubly robust estimator following Sant'Anna and Zhao (2020). In Panel B, we use geographic covariates to predict the outcome. The set of covariates includes temperature, precipitation, altitude, distance to the closest river, and to the closest national park. In Panel C, we add a set of municipality characteristics as covariates that includes the logarithm of population, a coca suitability index, distance to the country's capital, a rurality index, a poverty index, number of landmine victims over population, and number of guerilla attacks and clashes the year before demining. In Panel D, we residualized the outcome from municipality linear trends, that are computed using untreated observations. Bootstrap standard errors clustered at the event level. We also present p-values that control for the family-wise error rate in multiple hypotheses testing following Romano and Wolf (2005). \* is significant at the 10% level, \*\* is significant at the 5% level, \*\*\* is significant at the 1% level.

TABLE 4. Heterogeneous effects for economic and social outcomes by road connectivity – Humanitarian demining

	(1)	(2)	(3)	(4)	(5)	(6)
	Any road			Road's density		
	Paved or Unpaved	No roads	p-value diff.	High	Low	p-value diff.
<b>Panel A: Dep. var.: Nighttime Lights</b>						
Post demining	0.172*** (0.051)	0.071 (0.046)	0.045	0.179** (0.070)	0.084** (0.042)	0.173
<b>Panel B: Dep. var.: Test scores - Math</b>						
Post demining	0.091*** (0.032)	0.046 (0.030)	0.163	0.152*** (0.040)	0.037 (0.026)	0.003
<b>Panel C: Dep. var.: Test scores - Reading</b>						
Post demining	0.095*** (0.032)	0.070** (0.030)	0.436	0.142*** (0.037)	0.058** (0.025)	0.023
Observations (Panel A)	3360	4100		1860	5560	
Observations (Panel B)	3240	3720		1790	5170	
Observations (Panel C)	3240	3720		1790	5170	
Treated (Panel A)	139	155		88	206	
Treated (Panel B)	136	147		86	197	
Treated (Panel C)	136	147		86	197	
Never treated (Panel A)	197	255		98	354	
Never treated (Panel B)	188	225		93	320	
Never treated (Panel C)	188	225		93	320	

**Notes:** This table presents the overall ATT following Callaway and Sant'Anna (2020) for nighttime lights (Panel A) and students' performance (Panels B and C) after an humanitarian demining event. *Post demining* is the weighted average of all group-time average treatment effects with weights proportional to group size. The set of controls include not yet treated and never treated. The outcomes were computed using a radius of 5km around the demining. For each type of treatment, we divide the treated and never treated into events with any paved or unpaved road that crosses as close as 1km from the demined area and those with no close roads (columns 1 to 3). In columns 4 to 6, we follow a similar strategy but we divide the events into those that have a higher area of roads, measured by the length of the road that crosses as close as 1km from the demined area. We use the median of the empirical distribution to separate those with high and low connectivity. \* is significant at the 10% level, \*\* is significant at the 5% level, \*\*\* is significant at the 1% level.

TABLE 5. The impact of market access on nighttime lights

	Dependent variable: <i>Nighttime lights</i>								
	(1)	(2)	(3)	(4)	(5)	(6)	(7)	(8)	(9)
Log market access	0.199*** (0.064) [0.000]	0.238*** (0.084) [0.000]	0.250*** (0.074) [0.000]				0.145** (0.060) [0.001]	0.203** (0.093) [0.001]	0.212** (0.084) [0.000]
Cumulative hum demining				0.027*** (0.007) [0.000]	0.019*** (0.006) [0.000]	0.021*** (0.005) [0.000]	0.027*** (0.007) [0.000]	0.018*** (0.005) [0.000]	0.020*** (0.005) [0.000]
Observations	35,296	35,296	35,296	35,296	35,296	35,296	35,296	35,296	35,296
R-squared	0.890	0.915	0.918	0.890	0.915	0.918	0.890	0.915	0.918
Grid FE	Yes	Yes	Yes	Yes	Yes	Yes	Yes	Yes	Yes
Year FE	Yes	Yes	Yes	Yes	Yes	Yes	Yes	Yes	Yes
Dept-Year FE	No	Yes	Yes	No	Yes	Yes	No	Yes	Yes
Spatial controls	No	No	Yes	No	No	Yes	No	No	Yes
Mean dep variable	0.974	0.974	0.974	0.974	0.974	0.974	0.974	0.974	0.974

**Notes:** This table presents the impact of market access derived from humanitarian demining following the model suggested by Donaldson and Hornbeck (2016). The unit of analysis is a grid of size 15x15Km. The dependent variable is the hyperbolic transformation of nighttime lights. *Log market access* is the hyperbolic sine transformation of the nighttime lights-weighted market access. Spatial controls include the latitude and longitude and its square interacted with year fixed effects. Clustered standard errors at the municipality level are presented in parenthesis. In square brackets, we present the p-values for standard errors control for spatial and first-order time correlation (see Conley, 1999, Conley, 2016). We allow spatial correlation to extend to up to 416 km from each grid's centroid, which is the average distance from one municipality to all the rest. \* is significant at the 10% level, \*\* is significant at the 5% level, \*\*\* is significant at the 1% level.

TABLE 6. The local effects of humanitarian demining on other educational outcomes

	(1)	(2)	(3)	(4)	(5)
Dep. variable:	Enrollment	Share Test Taker	Ratio Student Teacher	School Entry	Share Approved
Post demining	0.295*** (0.081)	0.076** (0.033)	-0.292*** (0.075)	0.045* (0.025)	0.034** (0.015)
Observations	6453	4513	5913	6714	6714
Treated	288	287	288	294	294
Never treated	429	420	427	452	452
Average dep var	5.389	0.165	2.358	0.155	0.755

**Notes:** This table presents the overall ATT following [Callaway and Sant'Anna \(2020\)](#) for humanitarian demining during peace. *Post demining* is the weighted average of all group-time average treatment effects with weights proportional to group size. The set of controls include not yet treated and never treated. The outcomes were computed using a radius of 5km around the demining. Bootstrap standard errors clustered at the municipality level. \* is significant at the 10% level, \*\* is significant at the 5% level, \*\*\* is significant at the 1% level.

TABLE 7. Heterogeneous effects for forest loss by soil suitability

	(1)	(2)	(3)	(4)	(5)	(6)	(7)	(8)	(9)
	<b>Dependent variable: <i>Forest loss</i></b>								
During:	Peace						Conflict		
Demining:	Humanitarian			Military			Military		
Suitability:	Low	High	p-value diff.	Low	High	p-value diff.	Low	High	p-value diff.
<b>Panel A: Oil palm</b>									
Post demining	-0.072 (0.084)	0.037 (0.061)	0.294	-0.015 (0.033)	0.472*** (0.031)	0.000	0.021 (0.013)	0.189*** (0.017)	0.000
<b>Panel B: Cattle</b>									
Post demining	-0.020 (0.078)	-0.015 (0.067)	0.961	0.082*** (0.030)	0.450*** (0.035)	0.000	0.019 (0.014)	0.163*** (0.015)	0.000
<b>Panel C: Grass</b>									
Post demining	0.032 (0.081)	0.028 (0.065)	0.969	0.112*** (0.027)	0.364*** (0.034)	0.000	0.058*** (0.013)	0.156*** (0.016)	0.000
<b>Panel D: Rubber</b>									
Post demining	-0.016 (0.087)	0.014 (0.063)	0.780	0.084*** (0.032)	0.357*** (0.030)	0.000	0.046*** (0.014)	0.163*** (0.017)	0.000
<b>Panel E: Forestry</b>									
Post demining	0.001 (0.083)	0.052 (0.071)	0.641	0.157*** (0.032)	0.285*** (0.032)	0.005	0.108*** (0.013)	0.084*** (0.016)	0.244
<b>Panel F: Non-extractive traditional crops</b>									
Post demining	-0.091 (0.072)	0.024 (0.084)	0.299	0.295*** (0.030)	0.239*** (0.040)	0.263	0.086*** (0.014)	0.109*** (0.015)	0.262
Observations	3680	3780		44660	55900		119364	93636	
Treated	125	169		4309	5321		8640	6510	
Never treated	243	209		157	269		1307	1293	
Average dep var	2.811	3.757		3.147	4.179		2.904	4.039	

**Notes:** This table presents the overall ATT following Callaway and Sant'Anna (2020) for forest loss using the three demining treatments. *Post demining* is the weighted average of all group-time average treatment effects with weights proportional to group size. The set of controls include not yet treated and never treated. The outcomes were computed using a radius of 5km around the demining. For each type of treatment we divide the treated and never treated into high and low suitability. We do this by constructing the share of the area around the event with suitability for each activity and then define as high (low) the ones with suitability above (below) the median. The non-extractive traditional crops is a z-score index that includes rice, maize, potato, onion, and pepper. \* is significant at the 10% level, \*\* is significant at the 5% level, \*\*\* is significant at the 1% level.

## APPENDIX (FOR ONLINE PUBLICATION)

## A ADDITIONAL VARIABLES AND DESCRIPTIVE STATISTICS

**A.1 Soil suitability** To study heterogeneous effects based on soil suitability, we build a novel cross-section database by rasterizing –at a resolution of approximately  $30\text{m} \times 30\text{m}$  the suitability zoning shapefiles provided by the Agricultural Rural Planning Unit (UPRA) of the Colombian Ministry of Agriculture. Each pixel contains estimated information on different degrees of suitability for a wide range of activities, based on physical, socio-ecosystemic, and technical factors; socioeconomic and legal criteria; and regulatory guidelines that affect the delimitation of areas according to national level planning and regulation.<sup>46</sup> We construct information for key activities largely related to deforestation, such as palm oil growing, cattle herding, growing grass used for cattle, rubber crops, banana crops, and forestry. Based on this information, we can create soil suitability at the buffer level using the proportion of the buffer area with suitability for each type of land use.

**A.2 Road network** We also use detailed information of the location of the entire roads network of Colombia, including all road types from primary (highways) to tertiary (intra municipal non-paved) roads. These data were obtained from the *Instituto Geográfico Agustín Codazzi (IGAC)* for the 2012 cross-section.

**A.3 Geographic and municipality characteristics** We complement the variables described above, used either as the main outcomes or to test potential mechanisms, with weather and geographic characteristics that are geo-referenced at the buffer level. These include temperature and rainfall measures, altitude, distance to rivers and national parks, and the terrain’s ruggedness (Nunn and Puga, 2012).<sup>47</sup> Finally, we also add municipality characteristics from the CEDE municipal panel, compiled by Acevedo et al. (2014).

**A.4 Summary statistics** We start by plotting the intensity of demining across the country, aggregated for the whole period. Figure A12 plots the spatial distribution of demining events between 2004 and 2019. During this period, there was at least one demining event in 438 (38%) of the municipalities in Colombia. In Table A13, we present summary statistics

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<sup>46</sup>The physical component considers temperature, precipitation, climatic index, adequate depth, soil moisture, nutrient availability, textural class, degree of erosion, slope, landslide susceptibility, flood susceptibility, and volcanic hazards. The socio-ecosystem component includes ecological integrity, land cover, fire hazards, strategic ecosystems, and deforestation. Finally, the socioeconomic component considers institutional framework, security, labor market, living conditions, land size distribution, infrastructure and logistics, cost of rural land, and municipal economic indicators. For a detailed discussion about the weights of each criterion and the construction process, see <https://www.upra.gov.co/uso-y-adequacion-de-tierras/evaluacion-de-tierras/zonificacion>.

<sup>47</sup>The temperature, rainfall, and altitude data were constructed from Fick and Hijmans (2017), and the distance to rivers and parks was computed based on IDEAM national shapefiles of rivers and national parks.

for each outcome variable as measured within the sample buffers used to identify the effect of each of the three demining types. Overall, we find that areas, where humanitarian demining took place, tend to have more intense nighttime luminosity, better students' test performance, and lower levels of coca crops. Population density is larger for areas demined during the post-conflict period (regardless of whether demining took place in military operations or carried out by humanitarian agencies) and forest loss is similar across the three treatments.

In Figure [A13](#), we show that there are no substantial differences in grid characteristics across demined and not-demined areas within the same municipalities. Likewise, we find no substantial differences regarding the timing of demining events.



## B ROBUSTNESS EXERCISES

**B.1 Other estimation methods** In addition to Callaway and Sant’Anna (2020), other econometric procedures have been recently proposed to estimate causal effects in difference-in-differences settings with staggered adoption. Our results are robust to using three of them. First, we estimate the effects of demining using Borusyak et al. (2021)’s approach in which the ATT is a weighted average of individual treatment effects. The individual treatment effects are in turn estimated with an imputation technique that recovers the missing (non-treated) potential outcome of the treated units. This counterfactual is constructed using a linear model for untreated observations.<sup>48</sup> Figures A14 to A16 report the dynamic specification resulting from this estimation procedure and Panel A of Tables A14 to A16 report the overall treatment effects. Reassuringly, most outcomes follow the same pre-treatment dynamics (with perhaps a more pronounced decreasing pre-trend for the case of population density in military demining). Moreover, in terms of the ATTs, the effects are similar to the baseline estimates reported in Tables 1 to 3 for the three treatments.<sup>49</sup>

The second alternative approach that we explore is the one suggested by De Chaisemartin and d’Haultfoeuille (2020). In this case, the authors compute an ATT that measures the instantaneous treatment effect of moving from being untreated to becoming treated. Again, this model yields no substantial differences in terms of pre-treatment dynamics (see Figures A17 to A19 in the Appendix). Panel B of Appendix Tables A14 to A16 report the overall ATTs derived from this model. In general, the estimates are of similar magnitude and significance for humanitarian demining, except for the increase in population density which has half the size found in the baseline estimate. In the case of both treatments pertaining to demining in military operations, the results are also similar to the baseline estimates except for the changes in forest loss, the magnitude of which is half the reported for the Callaway and Sant’Anna (2020) model.

Finally, we estimate a model proposed by Wooldridge (2021), where the estimated ATT is a weighted average of the post-treatment group-year dummies that are estimated with a linear regression over the full sample of three years around the event date. We aggregate the group-year dummies in the same way we do it for the baseline specification and find treatment effects that are similar to those of Borusyak et al. (2021) (see Panel C of Appendix Tables A14 to A16).

<sup>48</sup>We use the balanced version of their estimate to avoid results arising from changes in sample composition.

<sup>49</sup>The magnitudes of the estimates are somewhat smaller for the effect of humanitarian demining on population density and for that of both military demining treatments on forest loss. In turn, they are larger for the effect of humanitarian demining on forest loss.

**B.2 Accounting for spillover effects** We now explore the presence of spillover effects in our baseline results. Usually, demining is not an isolated event, especially in the case of demining events in military operations, which typically clear entire strategic corridors to allow the ground mobilization of equipment and soldiers from one area to the other. This implies that some of our not-yet-demined controls may soon become demined and thus may contaminate our overall comparison group. We explore the extent of this potential threat to the internal validity of our results with two different but complementary strategies. First, we keep the whole sample and add as a covariate (in a doubly-robust way Sant’Anna and Zhao, 2020) a dummy that identifies whether there was a (one year) prior demining event in the 3, 5, or 7Km buffer around the current demining event.<sup>50</sup> Tables A17 to A19 of the Appendix report the results, which prove quite similar to those coming from the baseline specification that accounts for no potential geographic spillover.<sup>51</sup>

Second, we exclude from our estimating sample any demining event that experienced a (one year) prior demining within a 3, 5, or 7Km buffer. Tables A20 to A22 of the Appendix also suggest that this alternative approach to account for potential spillovers yields similar results relative to our baseline specification.<sup>52</sup>

**B.3 Alternative comparison groups** Recall that our baseline specification uses as a comparison group both the “never-treated” and the “not-yet-treated” mined areas. This section explores the robustness of our results to only using either of the two components. First, Table A23 reports the average ATTs of the three demining treatments on all the outcomes using only the “never-treated” as the comparison group. The effects of both treatments related to demining in military operations become noisier, which is consistent with this treatment having fewer “never-treated” controls. Second, Table A24 reports the average ATTs of the demining treatments using only the “not-yet-treated” as the comparison group. Again, the estimates are of similar magnitude and significance but the effects of humanitarian demining on nighttime lights and population density become more imprecisely estimated.

**B.4 Accounting for anticipation effects** We allow for potential anticipation of the demining treatments by excluding from the comparison group the observations that occur

<sup>50</sup>We also control for an indicator of whether there was a demining event that intersects with the buffer around never-treated controls during the year prior to the current demining event.

<sup>51</sup>Two exceptions are a smaller and more imprecise estimate of the increase in students’ performance in math test scores after humanitarian demining, and a smaller and more imprecise estimate of the decrease of population density after demining in post-conflict military operations.

<sup>52</sup>Two exceptions are a smaller and more imprecise estimate of the decrease in population density after a demining event in a post-conflict military operation, and a smaller increase in forest loss after a demining event in a military operation during the conflict.

one year prior to the demining events.<sup>53</sup> This is particularly important for humanitarian demining efforts as it may take several months for a targeted area to be fully cleared. Table A25 shows that the baseline results are robust to allowing for one-year anticipation. The only exception is that now forest loss increases after humanitarian demining.

**B.5 Alternative buffer sizes** Our results are also robust to changing the size of the buffers that we draw around the geo-referenced demining events to delimit the area within which we study the local effects of demining. Appendix Tables A26 to A28 and Figures A20 to A22 report respectively the average estimates and event studies obtained when defining buffers of radii 3, 4, 6, and 7Km. The results are remarkably similar both in terms of size and significance.

**B.6 Excluding events close to re-integration areas for ex-combatants** One potential concern is that the positive effects we find for humanitarian demining are driven by other public policies that affect these areas. We address this concern by excluding the demining events that happened in the surrounding of the 24 zones that were designated to host FARC ex-combatants after the signing of the peace agreement. In these zones, the government provided job-training, support for productive activities, and security (Presidencia de la República de Colombia, 2018). In Table A29, we show that the effects are robust to excluding these areas.

**B.7 Excluding events close to more populated areas** Our results are not driven by the few demining events that are close to relatively more populated areas.<sup>54</sup> Appendix Figures A23 to A25 report the effect of each demining treatment on each outcome of interest, excluding events that take place within 250m, 500m, 750m, or 1km from the centroid of a populated area.

**B.8 Outliers** In Table A30, we show that the results are robust to removing outliers by winsorizing the dependent variables at the 1 and 3% levels of the empirical distribution.

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<sup>53</sup>By doing so, we change the comparison year of the ATT estimates from one to two years before the demining event.

<sup>54</sup>These areas called *centros poblados* by the Colombian Statistics Bureau and are the urban centers of the municipalities, where the city hall and other institutional supply is located.

## C MARKET ACCESS FRAMEWORK

Following Donaldson and Hornbeck (2016), we measure welfare using nighttime luminosity and construct a market access measure with the following equation:

$$(A1) \quad MA_{ot} \approx \sum_d \tau_{odt}^{-\theta} L_d$$

where  $o$ ,  $d$ , and  $t$  stand for origin, destination, and year;  $L_d$  reflects the average luminosity in the destination;  $\tau_{odt}$  is a measure of transportation costs and  $\theta$  is a trade elasticity parameter that we set to 3.88 following the authors.

We compute the above formula using the 2012 road network, that was in place before the beginning of the humanitarian demining efforts. Consistently, we use the 2012 nightlight intensity at the destination.  $\tau_{odt}$  is computed as the shortest path between  $o$  and  $d$  based on such a network and using Dijkstra’s algorithm.<sup>55</sup> Finally, the network’s yearly variation is derived from the demining events that take place on year  $t$  within 100 meters of one of the network roads. That is, we abstract from changes in market access due to other reasons, such as the construction of new roads or to changes in economic activity (i.e. nighttime lights) at the destination.

As in Donaldson and Hornbeck (2016), we then employ the estimated market access of each location as the independent variable of the following linear specification:

$$(A2) \quad \ln(L_{ot}) = \alpha_o + \alpha_{st} + \beta \ln(MA_{ot}) + \epsilon_{ost}$$

where  $s$  stands for department. Here, a unit of observation is each 15x15km grid cell of the country, and we use the cell centroid’s closest road to compute the shortest path.<sup>56</sup>  $\alpha_o$  and  $\alpha_{st}$  are cell-specific and department-year fixed effects.  $\epsilon_{ost}$  is an error term clustered at the municipality level –and we also report, in brackets, p-values that account for potential spatial as well as time correlation (Conley, 1999, 2016). Our parameter of interest,  $\beta$ , measures the elasticity between market access and nighttime luminosity.<sup>57</sup>

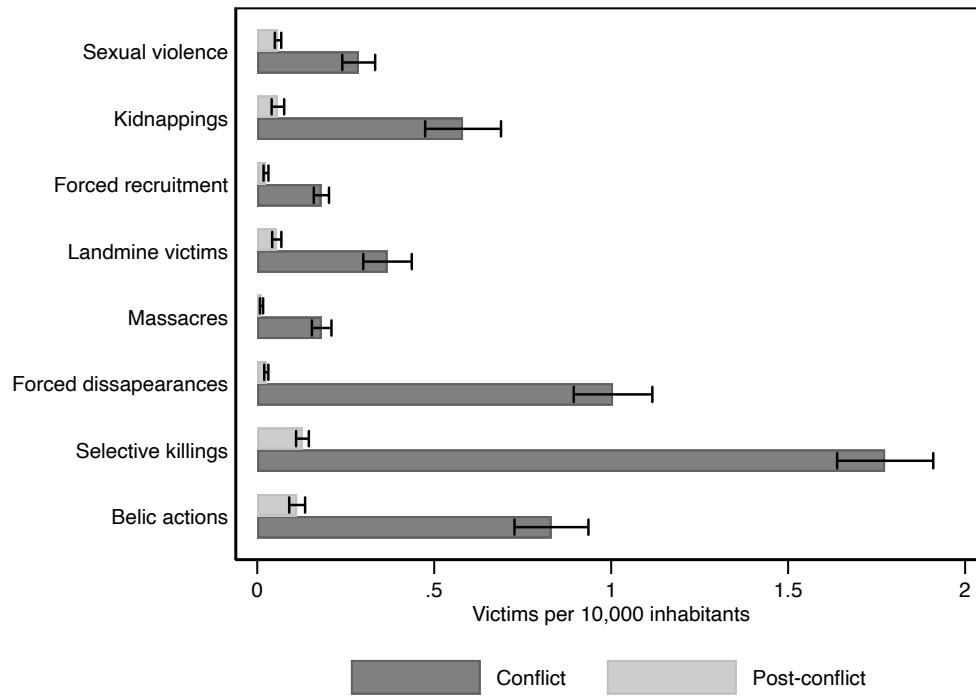
<sup>55</sup>Note that alternative off-road transportation infrastructure such as rails and rivers are very rarely used in Colombia (Champin et al., 2016; DNP, 2021).

<sup>56</sup>Results are robust to using larger (20x20km) grid cells (results available upon request).

<sup>57</sup>As in Chiovelli et al. (2019), we standardize  $\ln(MA_{ot})$  to facilitate the coefficient’s interpretation.

D ADDITIONAL TABLES AND FIGURES

FIGURE A1. Change in conflict since peace-negotiation



**Notes:** This figure presents the average incidence of conflict for different measures. We present the average across municipalities for the years between 2002 and 2012 (conflict) and between 2013 and 2019 (post-conflict). We present confidence intervals at the 95%.

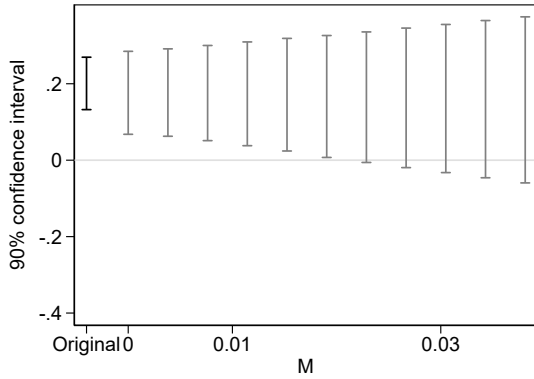
## FIGURE A2. Plan “Renacer” by FARC

Camaradas del Secretariado. Mi saludo

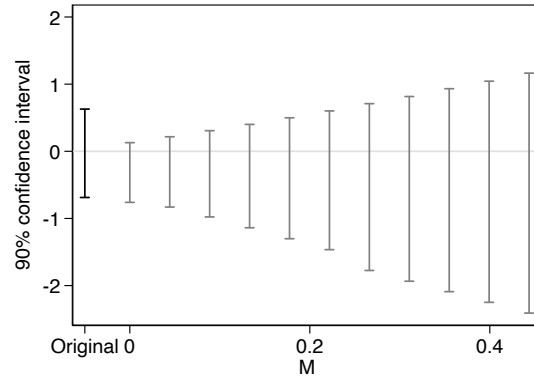
1. Importante las relaciones del camarada TIMO. Con los amigos colaboradores del presidente CHAVEZ. Vale la pena darles a conocer el plan estratégico, así como se le presento a su jefe, a su asesor y amigo CHACIN. Igual de importante reforzar en los encuentros con los ELENOS propiciados por el gobierno, la necesidad de crear la fusión en algunas regiones de dominio primordial de las FARC EP y buscar el apoyo de los asistentes a estas reuniones. A la senadora PIEDAD, hablarle sobre la necesidad de crear un partido del pueblo y buscar su alianza al movimiento Bolivariano.
2. Ya todos conocemos los cambios en la situación política del país y al mismo tiempo la situación interna de nuestra organización guerrillera, por eso es tiempo de realizar algunos cambios temporales y pasar nuevamente a la táctica de GUERRA DE GUERRILLAS, plan propuesto como “RENACER REVOLUCIONARIO DE LAS MASAS” es allí donde se encuentra la estrategia y el éxito de la guerra de guerrillas con el desarrollo del PLAN PATRIOTA y la mal llamada POLITICA DE SEGURIDAD DEMOCRATICA, el enemigo ha ganado espacio geográfico y por mal utilización de nuestros recursos sociales también hemos visto afectado el espacio político social. Situación un poco distinta a la manejada por el camarada SANTRICH y MATIAS con las células del Cauca, Valle y Nariño, estructuras que dejaron fortalecidas antes de trasladarse al área del Bloque Caribe. Por esto dentro del desarrollo de este plan propongo adelantar algunas actividades y otras ponerlas en consideración para su posterior ejecución.
3. Desarrollar por lo menos, antes de terminar el presente año, cursos de misiones especiales, programa desarrollado por el Comando Conjunto Central y que ha dado resultados positivos en corto tiempo luego de terminar el entrenamiento de las unidades.
4. Disponer de 5 a 6 millones de dólares del fondo del Secretariado, para adquirir intendencia, material de guerra y comunicaciones. Necesario para fortalecer la capacidad de lucha de los guerrilleros urbanos y milicias. Del manejo de este dinero se encargara el Bloque Oriental y cada bloque aportara entre 1 y 2 millones según condiciones para este fin.
5. Aumentar los visos defensivos y de movilidad con minados para detener el avance de las operaciones enemigas, ya conocemos que las minas son el único factor que los detiene y los intimida, por esto aumentar los cursos de explosivistas para lograr un nivel de conocimiento en explosivos, generalizados dentro de la guerrillerada e iniciar igualmente el entrenamiento del personal del MB y de milicias, haciendo énfasis en que no se debe de manipular los mismos con excesiva confianza los que lleva a accidentes.
6. El Comando Conjunto ya con capacidades en este ámbito, ejecutara algunas operaciones, para mantener el nombre de nuestra organización y evitar así crear un ambiente de derrota progresiva a las FARC EP.
7. En la medida que se vayan ejecutando los entrenamientos, como ejercicios finales se deben de colocar objetivos reales, que propicien golpes al enemigo.
8. Con el uso de minas y explosivos se equilibran las cargas frente a un enemigo numeroso, bastante equipado y con gran poder de fuego.
9. Los resultados logrados en el Guayabero, son una muestra de la necesidad de entrenar bien militarmente a las milicias y miembros del MB, aun cuando se trata de un poder invaluable y necesario, solo se encuentran proporcionando inteligencia y logística, situación que se dificulta cuando hay controles enemigos sobre las rutas o medios, ejemplo claro de esto es la situación presentada con Cesar. Hay que pensar en un mecanismo para reforzar ese mismo mecanismo sin exponer la seguridad y brindar más resultados al enemigo.
10. Es difícil para el enemigo mantener el despliegue de personal, material sobre un área permanente, por esto que al retomar la táctica de guerrillas móviles aunado con los golpes que pueden propinar las milicias y el MB, fortalecerá la presencia nuestra en áreas.
11. La táctica de francotiradores ya tratada desde la Octava Conferencia, se debe desarrollar con los recursos destinados dentro de la ejecución de este plan, adquirir el material necesario, fusiles y munición especializada por Bloque, el efecto de la ejecución de esta maniobra tendrá iguales resultados que los minados.
12. Los grupos encargados de la tarea telefónica se debe incrementar en todas las áreas de operaciones enemigas, está comprobado que estando lo bastando cerca de ellos arroja buenos resultados para IC.
13. Alistar por bloque unidades de confianza y que tengan el servicio militar para que se presenten como soldados profesionales y utilizarlos para IC. Como se esta trabajando en el Oriental y el Bloque Sur.
14. En la historia de las guerras de guerrillas, se ha demostrado que lo que ha creado un paralelo de negociación obligatorio entre la parte más fuerte y el apoyo aéreo, que termina por causar gran daño a la contraparte, pero también es claro que si se logra golpear este par, los resultados en la balanza se inclinan a favor, es por esto que se hace de extrema necesidad lograr la negociación de misiles que nos permitan propinar golpes contundentes al poderío aéreo del enemigo. Las tareas de destrucción de aeronaves mediante la infiltración como lo ha hecho el Oriental nos ha demostrado que el precio es alto y se cometen errores.

Es todo y espero sus opiniones. Alfonso

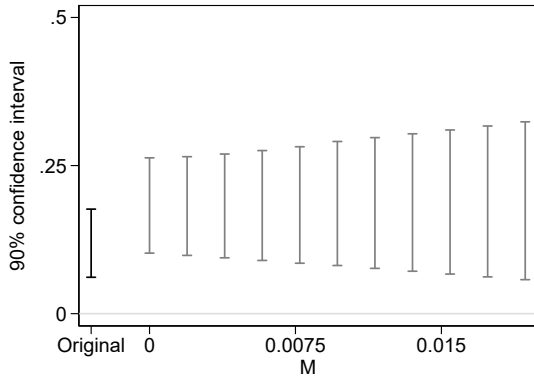
FIGURE A3. Violations to parallel trends assumption: Humanitarian demining during peace



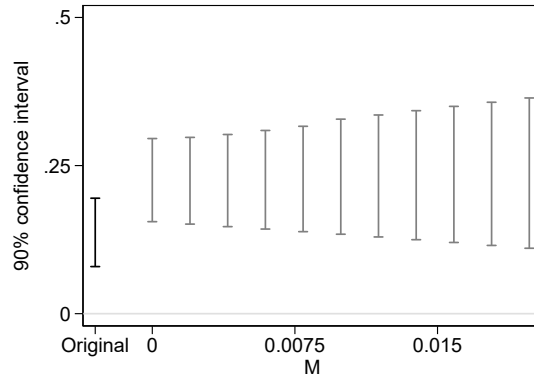
**A.** Nighttime lights



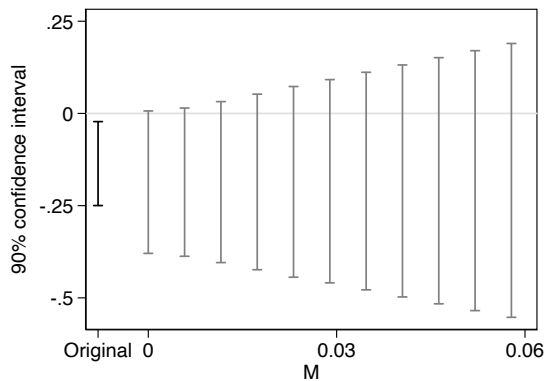
**B.** Population density



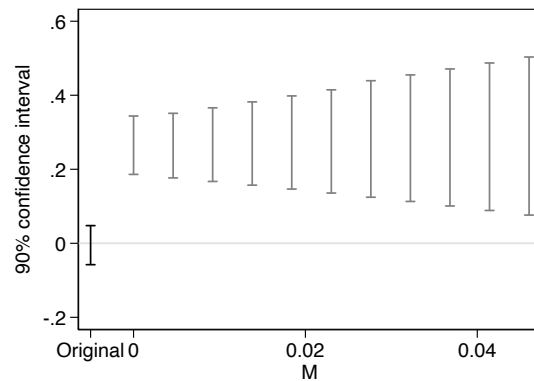
**C.** Test scores: Math



**D.** Test scores: Reading



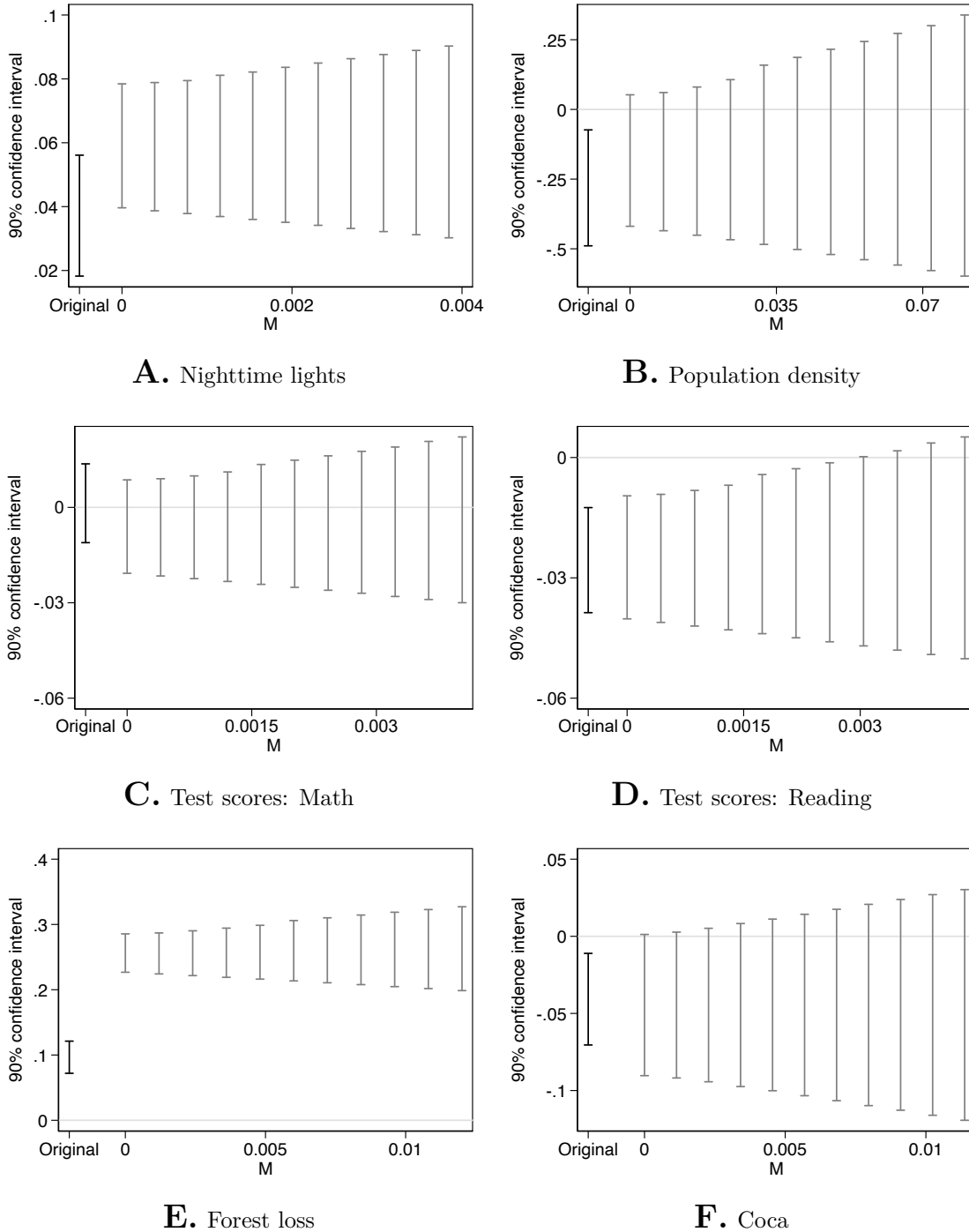
**E.** Forest loss



**F.** Coca

**Notes:** This figure presents the confidence set at 90% for linear and non-linear violation of the parallel trends assumption (Rambachan and Roth, 2021). The figure is shown for the coefficient the year after the demining event. The treatment is humanitarian demining during peace.  $M$  measures the size of the change in the trend between consecutive periods. Thus  $M = 0$  is a linear violation of the parallel trend assumption. The maximum value of  $M$  is equal to the trend that has a 50% power of being detected given the precision of the estimates in the pre-period (Roth, 2021).

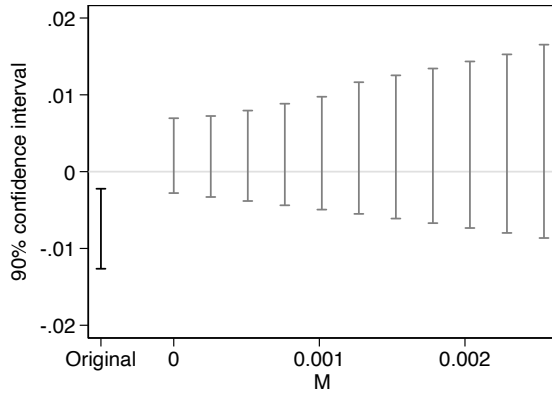
FIGURE A4. Violations to parallel trends assumption: Military demining during peace



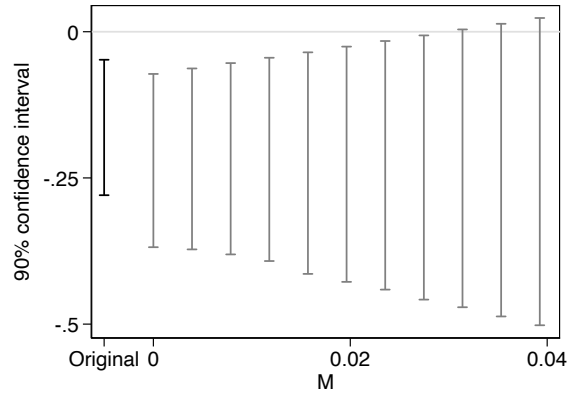
**Notes:** This figure presents the confidence set at 90% for linear and non-linear violation of the parallel trends assumption (Rambachan and Roth, 2021). The figure is shown for the coefficient the year after the demining event. The treatment is military demining during peace.  $M$  measures the size of the change in the trend between consecutive periods. Thus  $M = 0$  is a linear violation of the parallel trend assumption. The maximum value of  $M$  is equal to the trend that has a 50% power of being detected given the precision of the estimates in the pre-period (Roth, 2021).



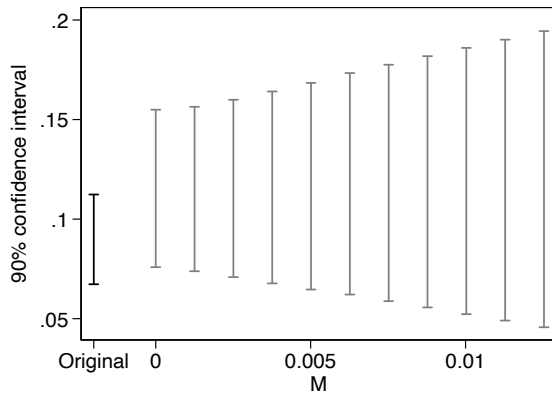
FIGURE A5. Violations to parallel trends assumption: Military demining during conflict



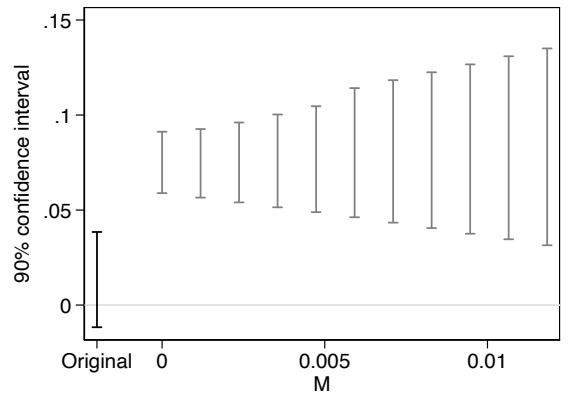
A. Nighttime lights



B. Population density



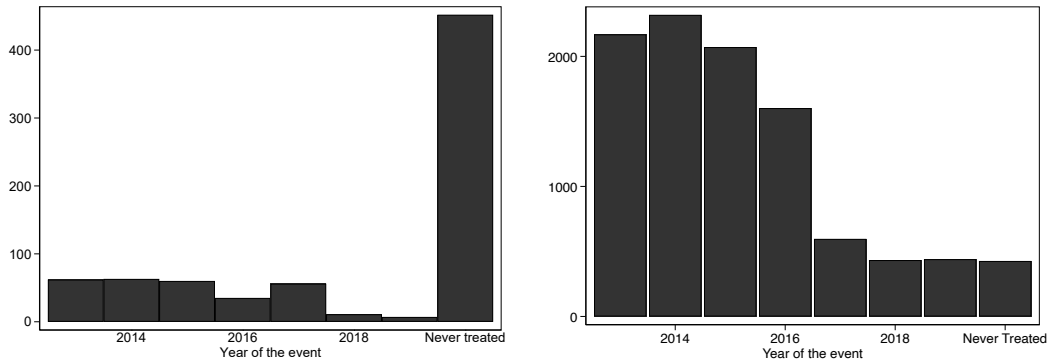
C. Forest loss



D. Coca

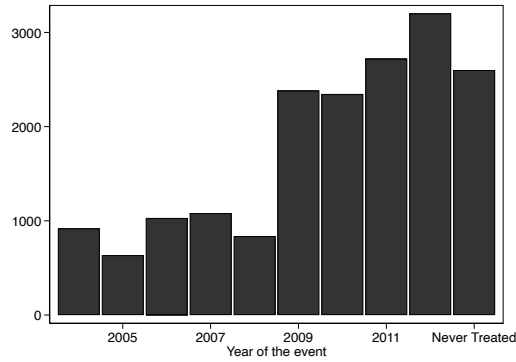
**Notes:** This figure presents the confidence set at 90% for linear and non-linear violation of the parallel trends assumption (Rambachan and Roth, 2021). The figure is shown for the coefficient the year after the demining event. The treatment is military demining during conflict.  $M$  measures the size of the change in the trend between consecutive periods. Thus  $M = 0$  is a linear violation of the parallel trend assumption. The maximum value of  $M$  is equal to the trend that has a 50% power of being detected given the precision of the estimates in the pre-period (Roth, 2021).

FIGURE A6. Events by year of occurrence



A. Humanitarian demining during peace

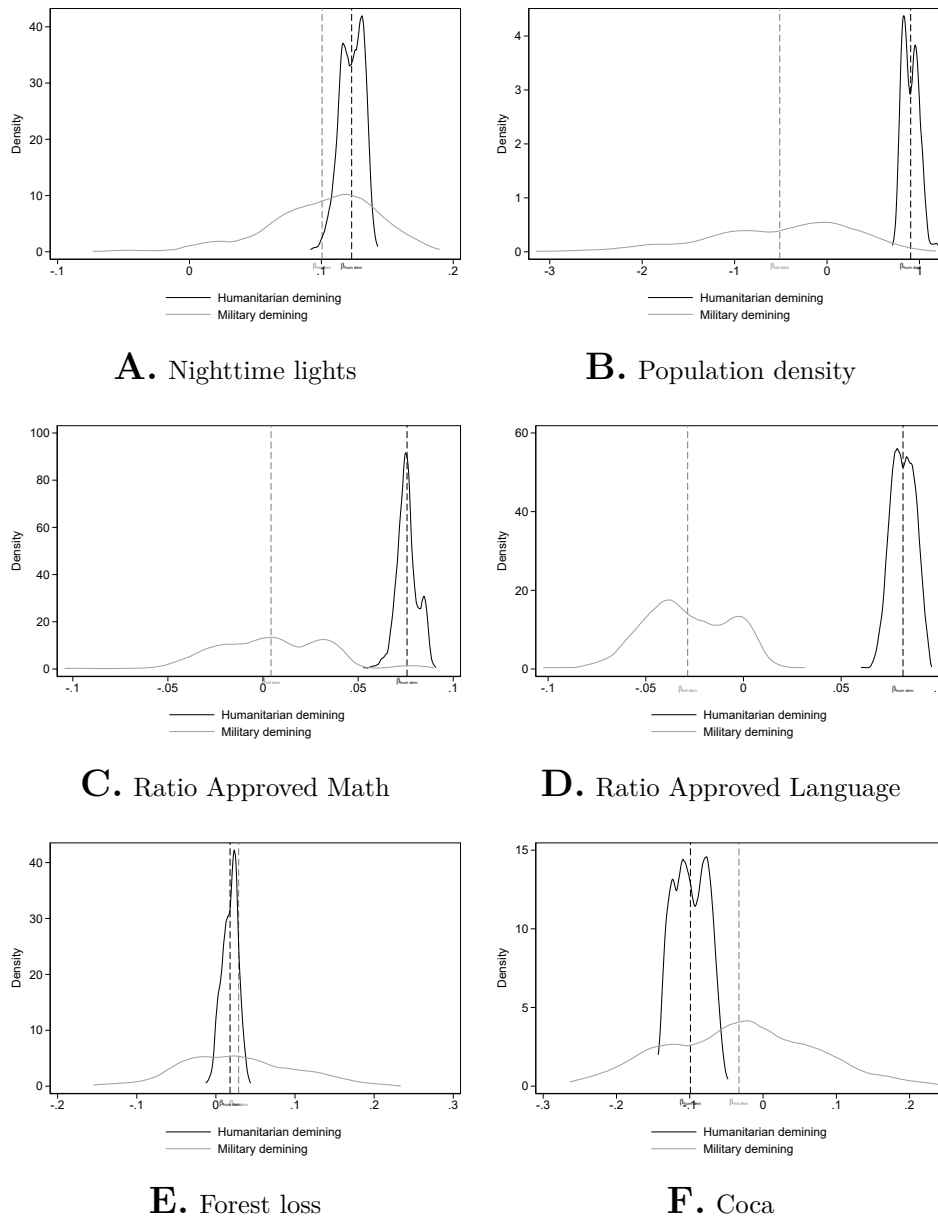
B. Military demining during peace



C. Military demining during conflict

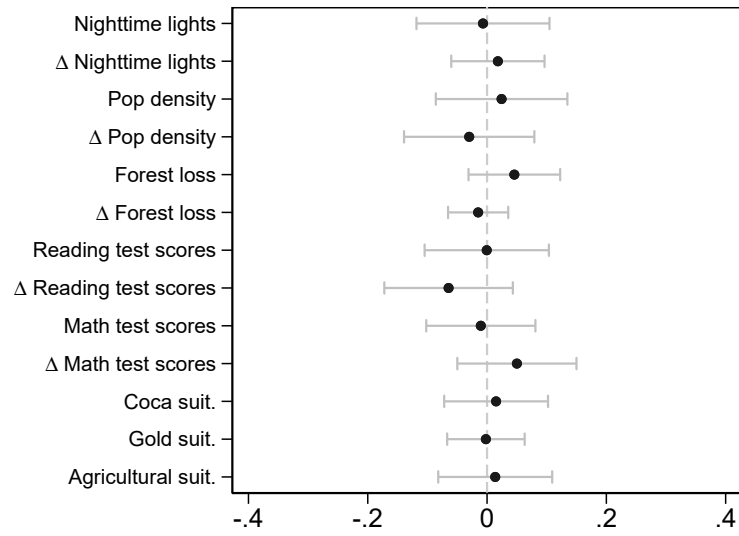
**Notes:** This figure presents the number of treated units by cohort and the never treated units for the three treatments, humanitarian demining during peace (panel A), military demining during peace (panel B), and demining during conflict (panel C).

FIGURE A7. Distribution of treatment effects based on samples with higher overlap based on different pcores

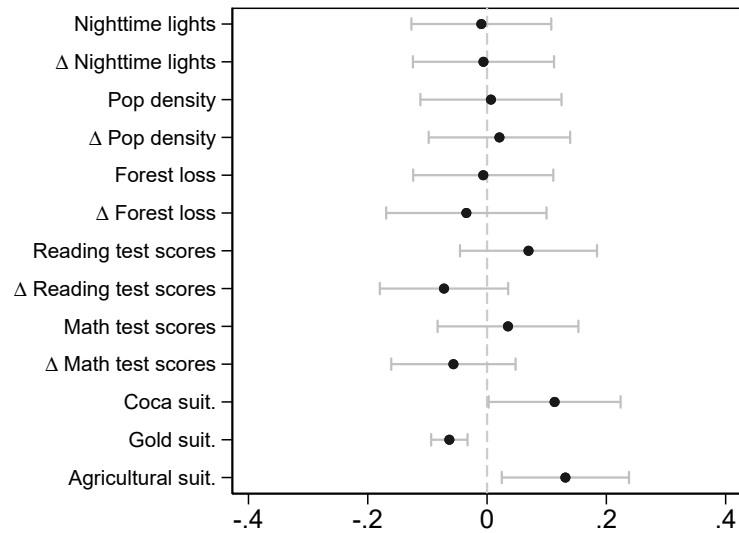


**Notes:** This figure presents the distribution of placebo ATT following Callaway and Sant’Anna (2020) for the treatments of humanitarian demining and military demining during peace. The set of controls include not yet treated and never treated. The outcomes were computed using a radius of 5km around the demining. For each iteration, we restricted the sample to the optimal selection rule from Crump et al. (2009) over different propensity score definitions on the probability of being a humanitarian demining. In each propensity score definition, we selected the covariates from the set of all possible combinations of variables, and we started with at least three covariates in each group. The covariates used to predict the probability were selected from the following group of variables; a poverty index, the logarithm of population, a rurality index, the distance to the closest department’s capital, the distance to the country’s capital, the distance to the closest national park, a coca suitability index, elevation, precipitation, and temperature.

FIGURE A8. Differential characteristics by road connectivity



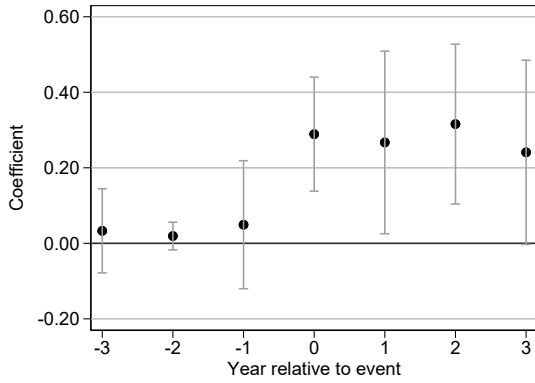
A. Any Road



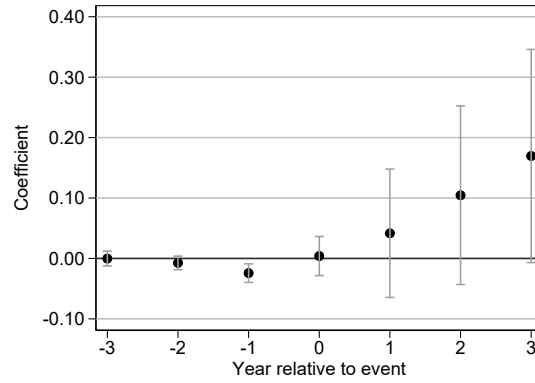
B. Road's Density

**Notes:** This figure presents the standardized differences by road connectivity. The sample is limited to municipalities that experienced any humanitarian demining since 2013. In Panel A, we present point estimates and confidence intervals for a regression of the indicator of any paved or unpaved road that crosses as close as 1km the demined area in the municipality. In Panel B, we do the same but in this case the dependent variable is an indicator for higher area of roads, measured by the length of the road that crosses as close as 1km from the demined area in the municipality. Variables with a  $\Delta$  are first differences of the variable taking an average of two years before 2013.

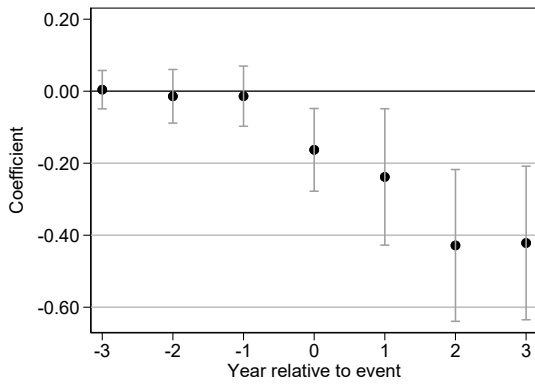
FIGURE A9. Humanitarian demining during peace and education



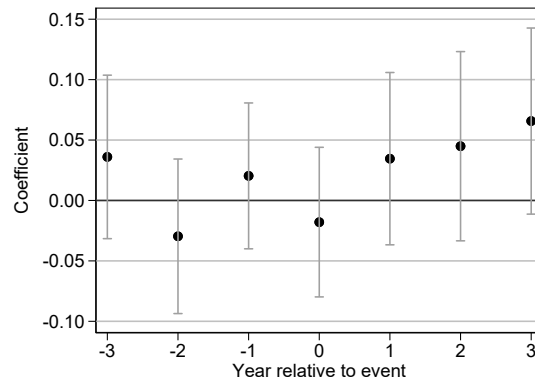
A. Enrollment - All



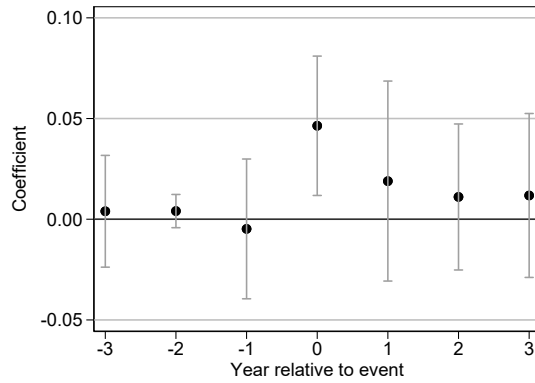
B. Share Test Taker



C. Ratio Student Teacher



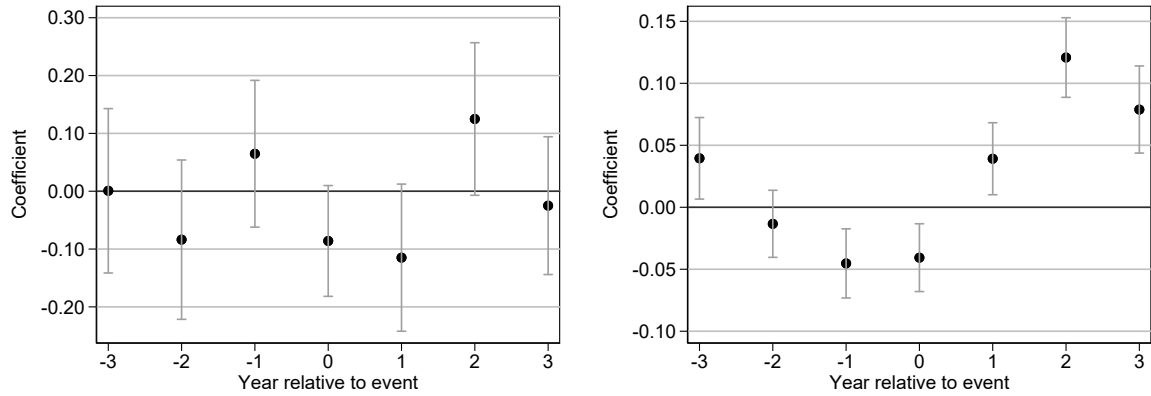
D. School Entry



E. Share Approved

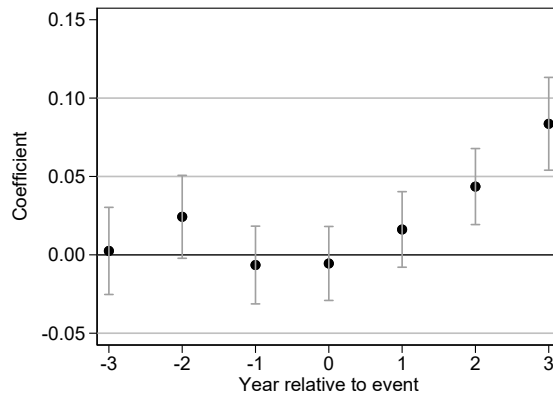
**Notes:** This figure presents the event study coefficients following Callaway and Sant'Anna (2020) for the treatment of demining during peace. We present the point estimates as well as the 95% confidence interval. Standard errors clustered at the event level. The outcomes were computed using a radius of 5km around the demining.

FIGURE A10. Demining and fires



**A.** Humanitarian demining during peace lights

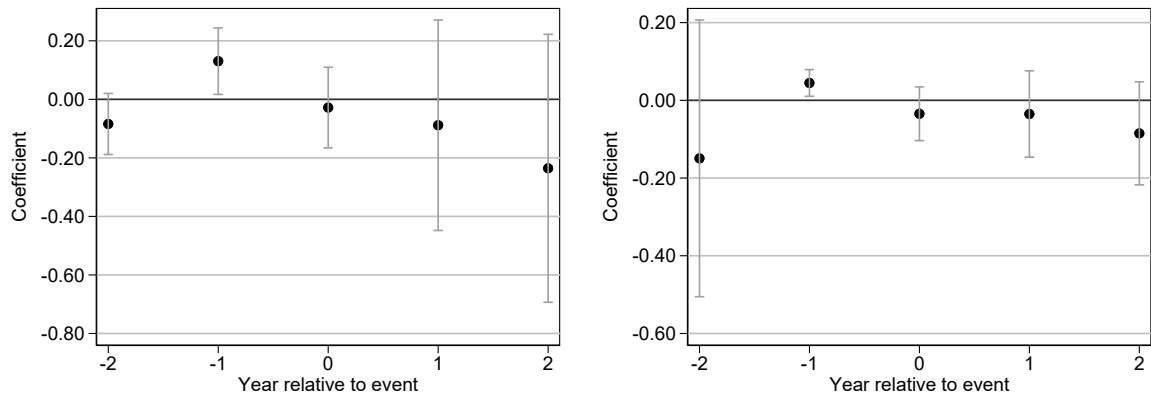
**B.** Military demining during peace



**C.** Military demining during conflict

**Notes:** This figure presents the event study coefficients following Callaway and Sant’Anna (2020) for the treatment of demining during peace. We present the point estimates as well as the 95% confidence interval. Standard errors clustered at the event level. The outcomes were computed using a radius of 5km around the demining.

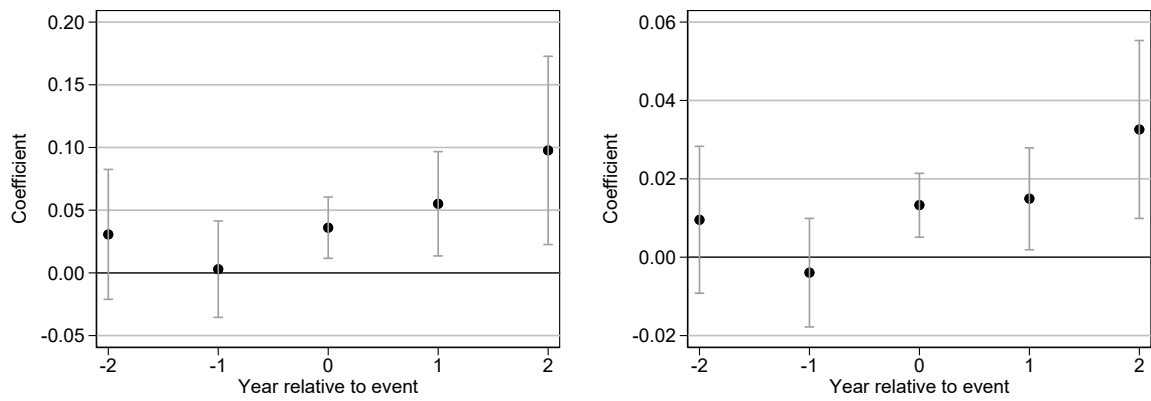
FIGURE A11. Demining and illegal mining  
Humanitarian demining during peace



**A.** Area illegal gold mining

**B.** Illegal gold mining

Military demining during peace

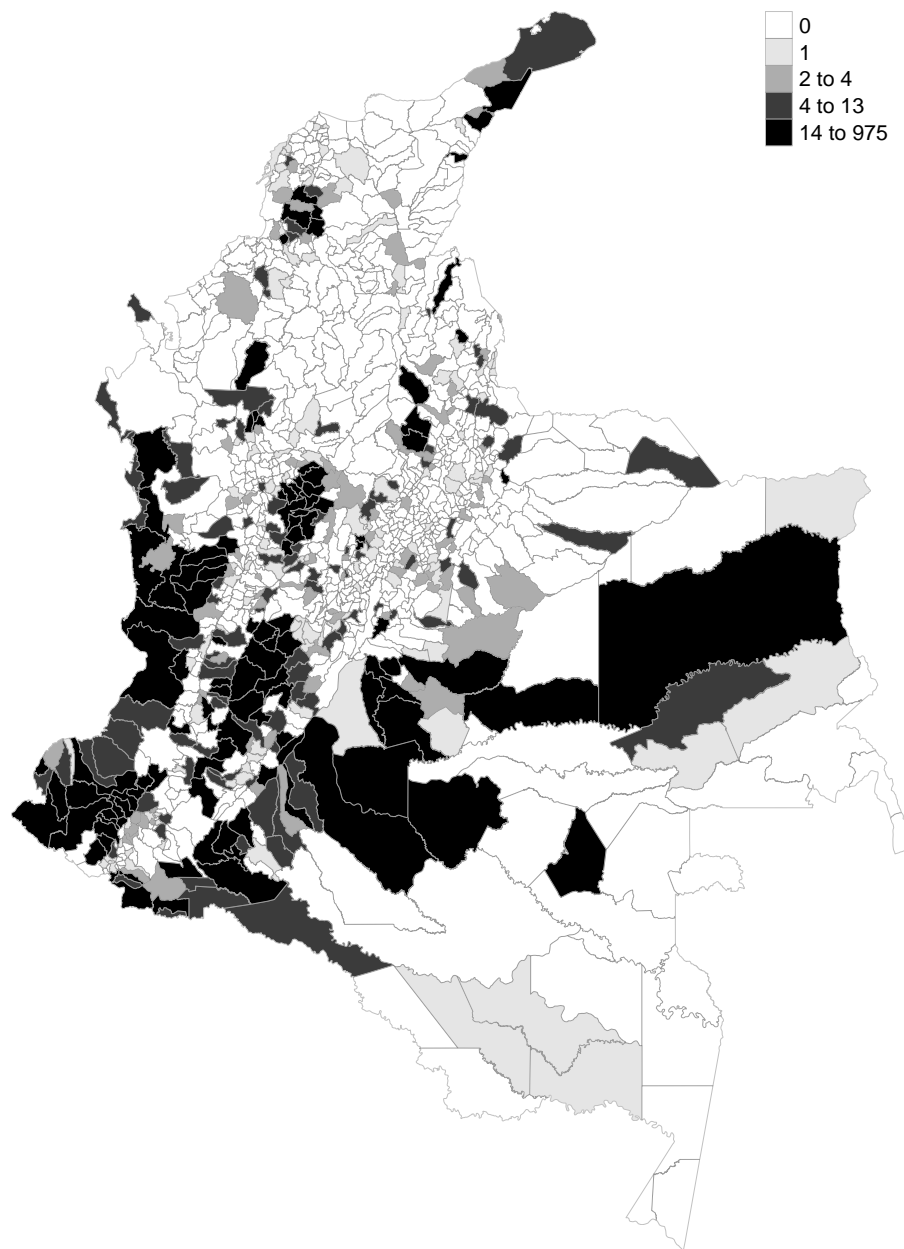


**C.** Area illegal gold mining

**D.** Illegal gold mining

**Notes:** This figure presents the event study coefficients following Callaway and Sant’Anna (2020) for the treatment of demining during peace. We present the point estimates as well as the 95% confidence interval. Standard errors clustered at the event level. The outcomes were computed using a radius of 5km around the demining.

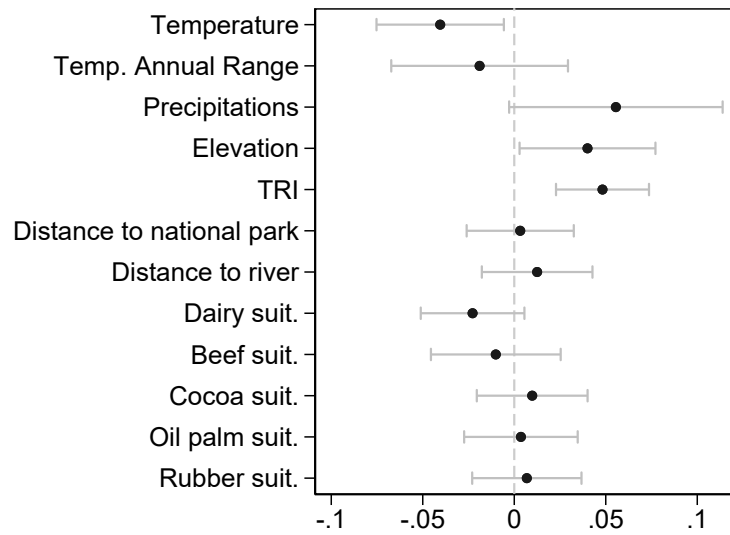
FIGURE A12. Number of demining events (2004-2019)



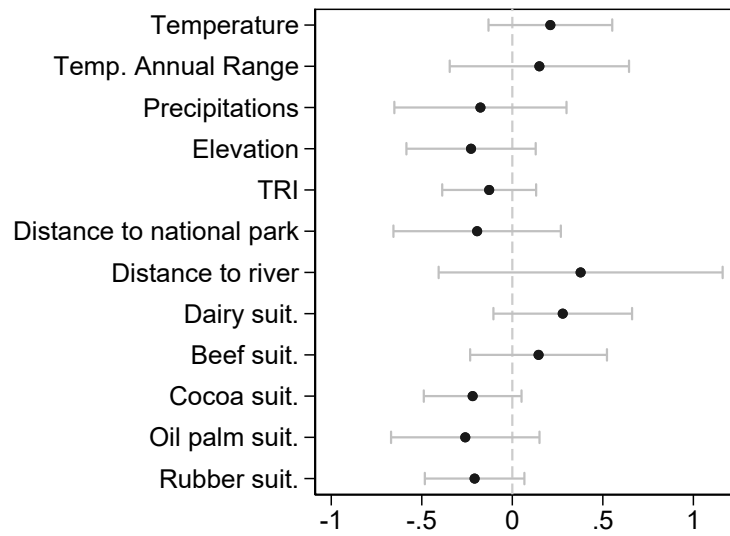
Notes: This map presents the spatial distribution of humanitarian and military demining events from 2004 to 2019.



FIGURE A13. Differential characteristics by treatment status and timing of treatment



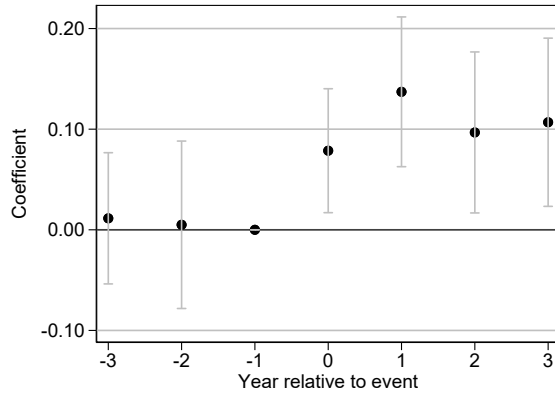
A. All: Treated versus non-treated



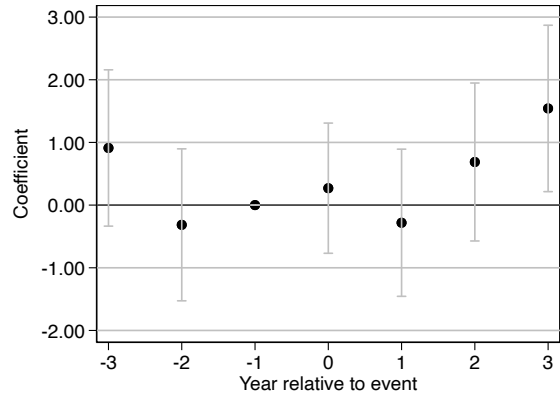
B. All: Timing of treatment

**Notes:** This figure presents the standardized differences by treatment status and treatment timing. In Panel A, we compare grids of 10x10km that were demined during 2004-2019 versus grids that were not demined within the same municipality. In Panel B, we compare within demined grids and look at the year of the first demining event. All characteristics are computed at the 10x10km grid.

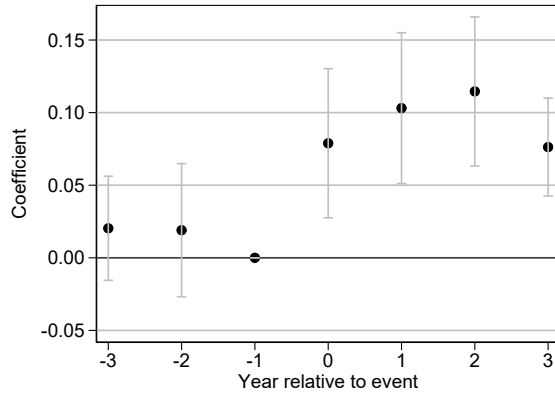
FIGURE A14. Humanitarian demining during peace and local activity: Borusyak et al. (2021)



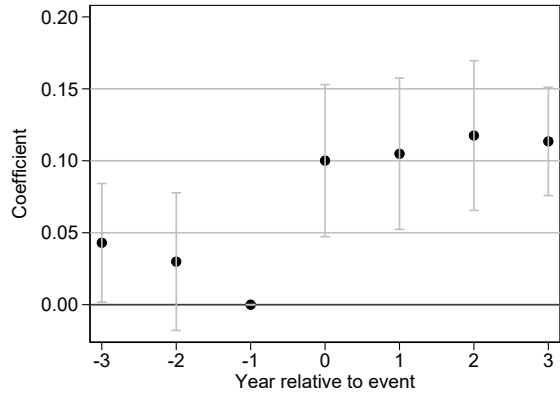
A. Nighttime lights



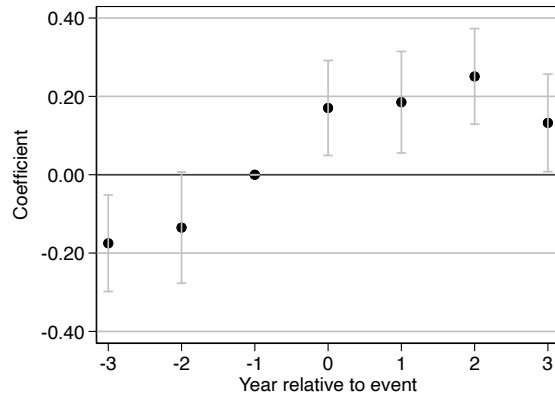
B. Population Density



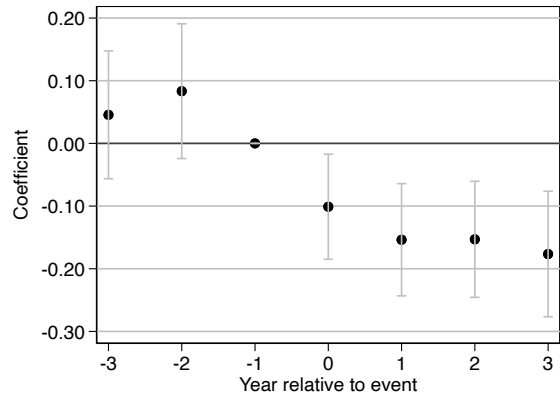
C. Test scores: Math



D. Test scores: Reading



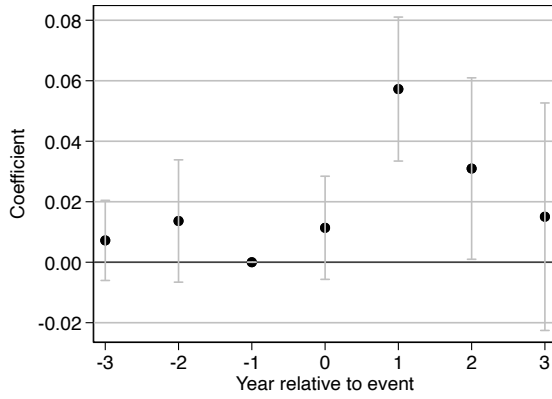
E. Forest loss



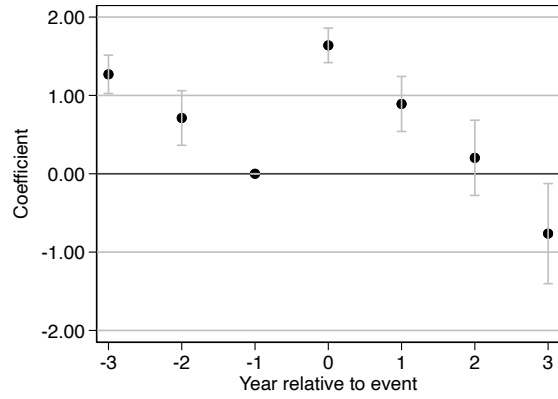
F. Coca

**Notes:** This figure presents the event study coefficients following Borusyak et al. (2021) for the treatment of humanitarian demining. We present the point estimates as well as the 95% confidence interval. Standard errors clustered at the event level. The outcomes were computed using a radius of 5km around the demining.

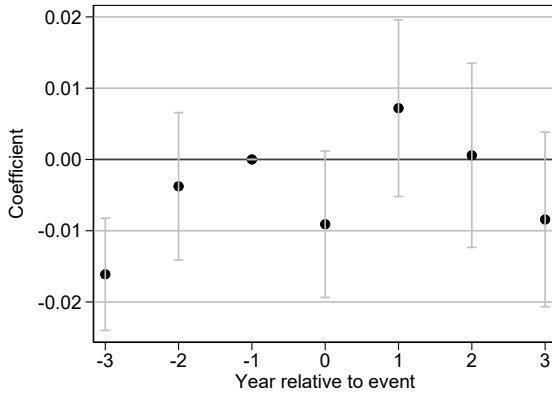
FIGURE A15. Military demining during peace and local activity: [Borusyak et al. \(2021\)](#)



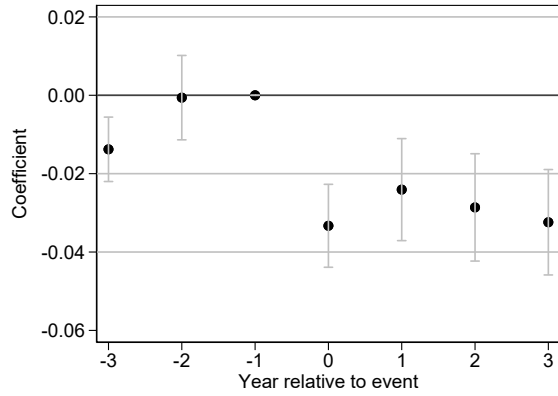
**A.** Nighttime lights



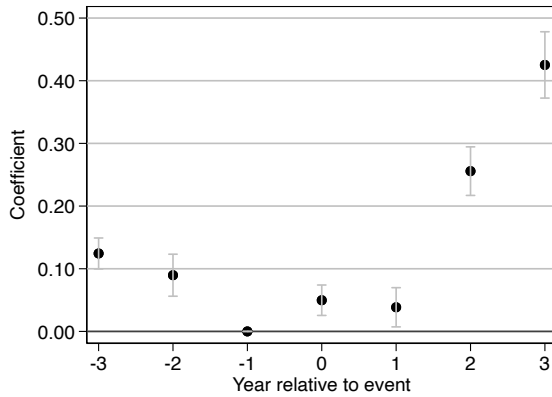
**B.** Population Density



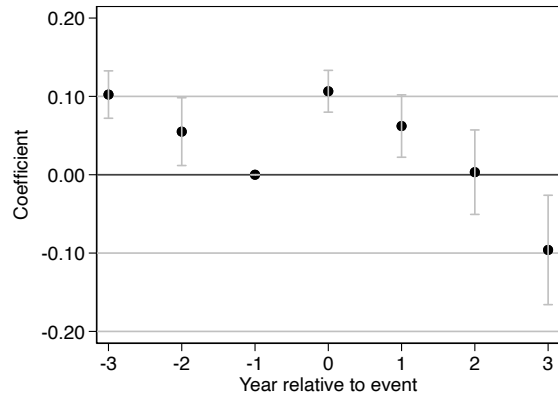
**C.** Test scores: Math



**D.** Test scores: Reading



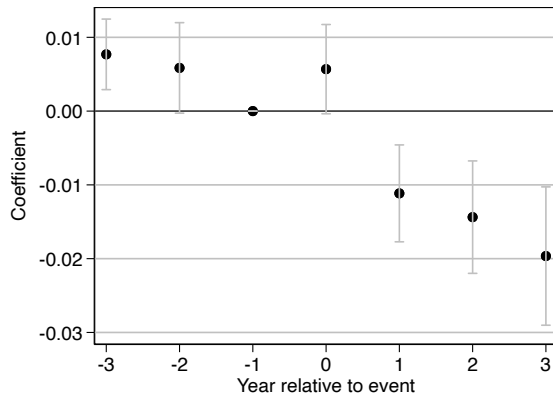
**E.** Forest loss



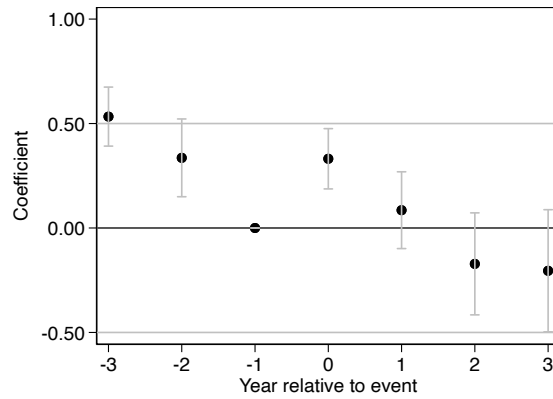
**F.** Coca

**Notes:** This figure presents the event study coefficients following [Borusyak et al. \(2021\)](#) for the treatment of demining during peace. We present the point estimates as well as the 95% confidence interval. Standard errors clustered at the event level. The outcomes were computed using a radius of 5km around the demining.

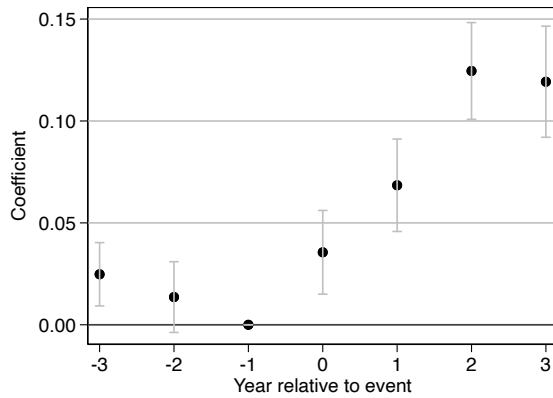
FIGURE A16. Military demining during conflict and local activity: [Borusyak et al. \(2021\)](#)



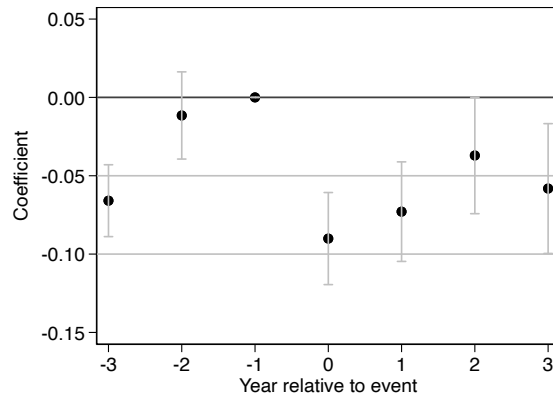
A. Nighttime lights



B. Population Density



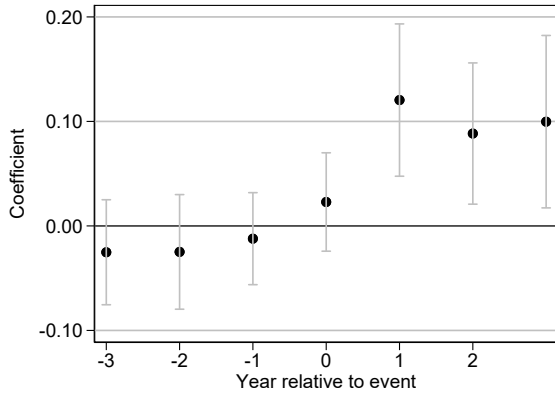
C. Forest loss



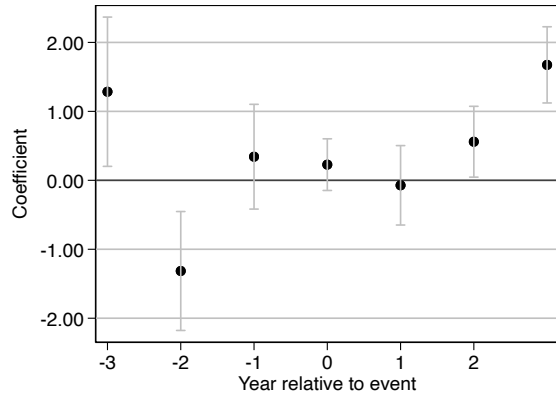
D. Coca

**Notes:** This figure presents the event study coefficients following [Borusyak et al. \(2021\)](#) for the treatment of demining during conflict. We present the point estimates as well as the 95% confidence interval. Standard errors clustered at the event level. The outcomes were computed using a radius of 5km around the demining.

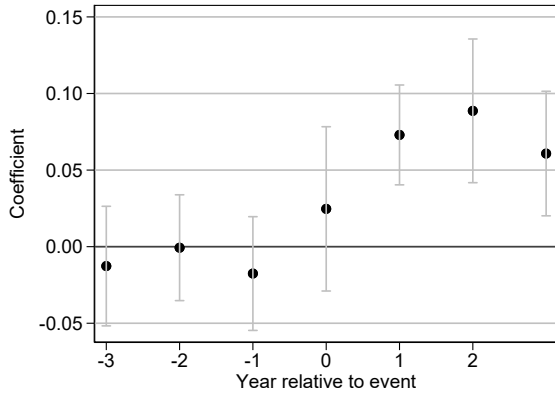
FIGURE A17. Humanitarian demining during peace and local activity: De Chaisemartin and d’Haultfoeuille (2020)



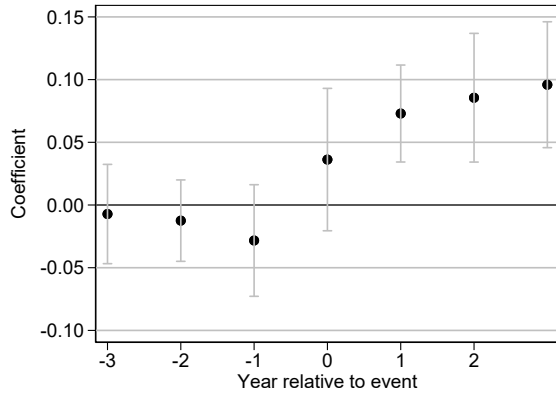
A. Nighttime lights



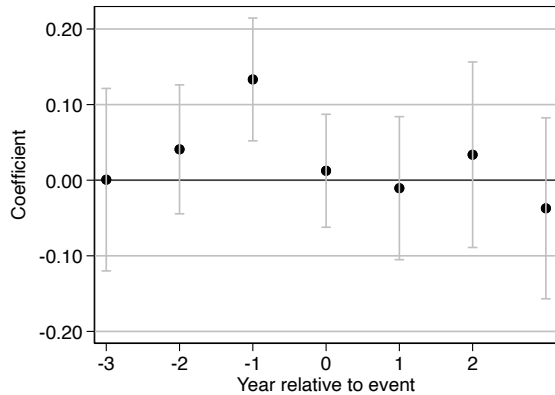
B. Population Density



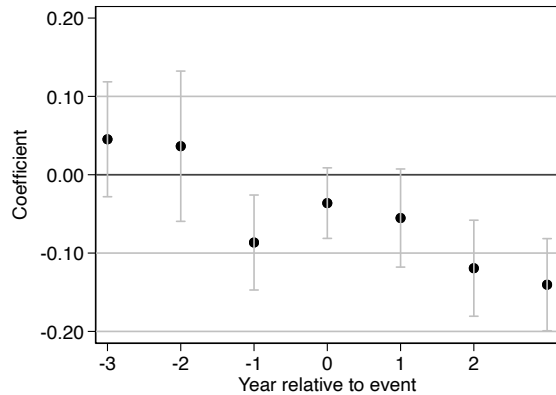
C. Test scores: Math



D. Test scores: Reading



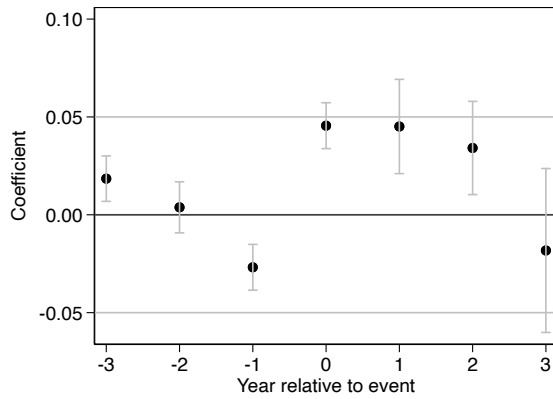
E. Forest loss



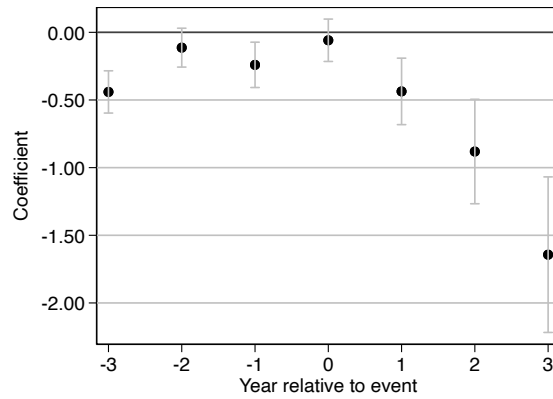
F. Coca

Notes: This figure presents the event study coefficients following De Chaisemartin and d’Haultfoeuille (2020) for humanitarian demining during peace. We present the point estimates as well as the 95% confidence interval. Standard errors clustered at the event level. The outcomes were computed using a radius of 5km around the demining.

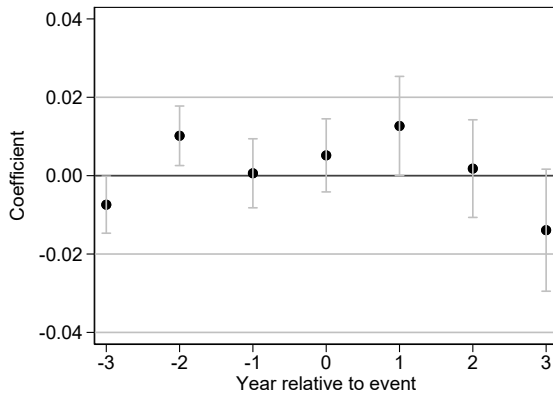
FIGURE A18. Military demining during peace and local activity: De Chaisemartin and d’Haultfoeuille (2020)



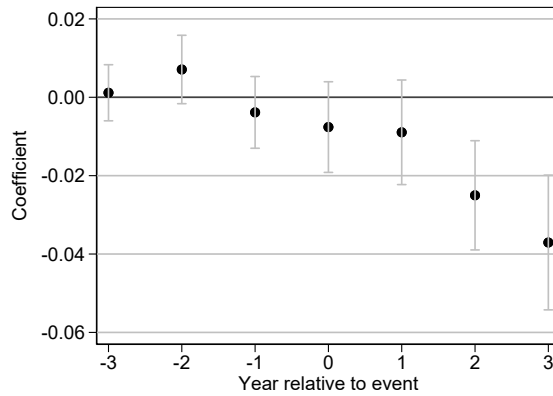
**A.** Nighttime lights



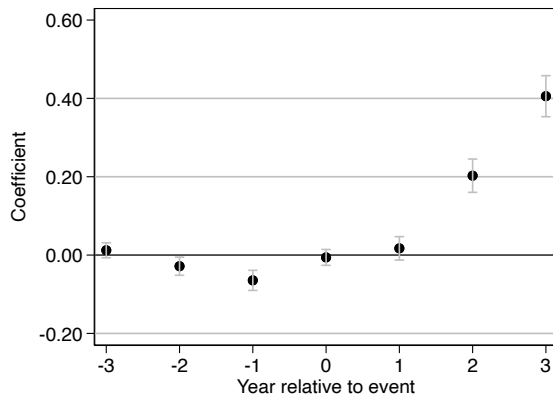
**B.** Population Density



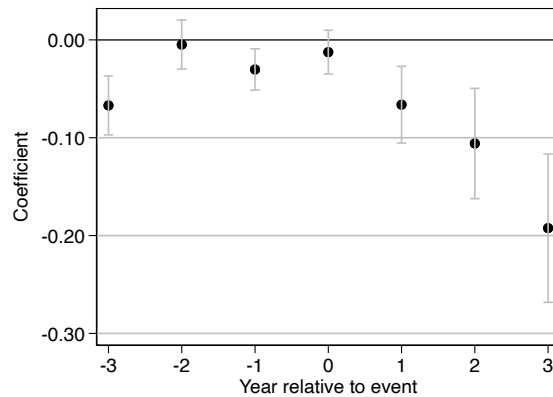
**C.** Test Scores: Math



**D.** Ratio Approved Reading



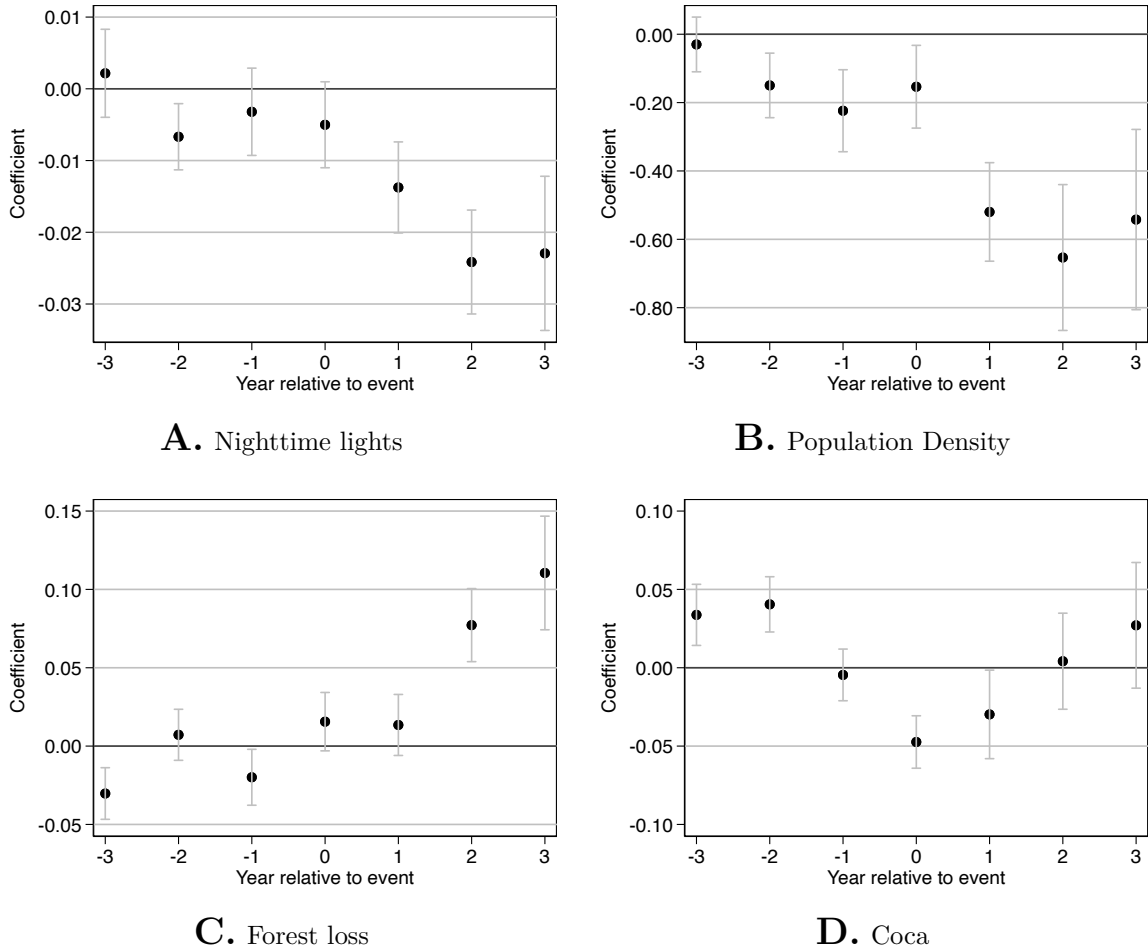
**E.** Forest loss



**F.** Coca

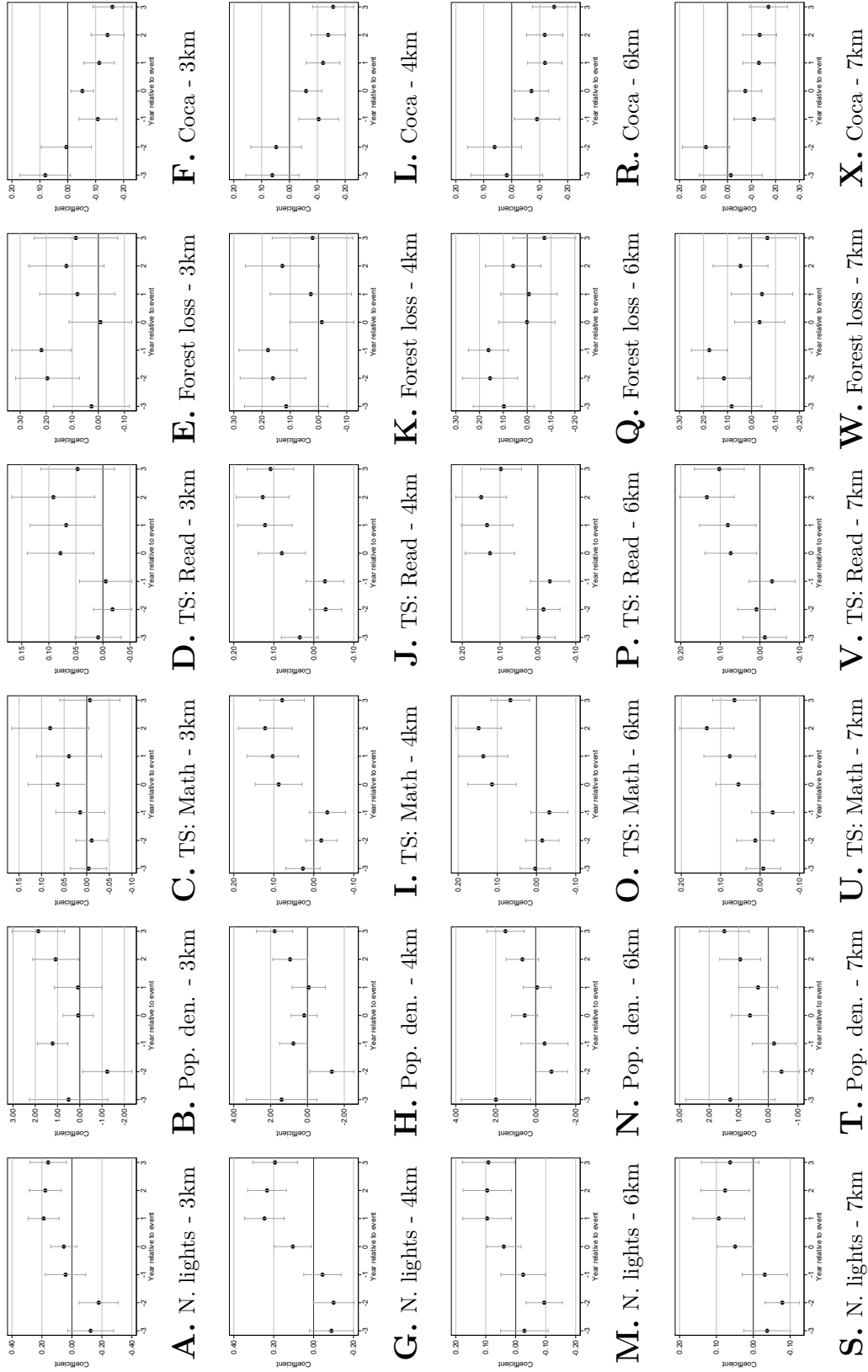
**Notes:** This figure presents the event study coefficients following De Chaisemartin and d’Haultfoeuille (2020) for military demining during peace. We present the point estimates as well as the 95% confidence interval. Standard errors clustered at the event level. The outcomes were computed using a radius of 5km around the demining.

FIGURE A19. Military demining during conflict and local activity: De Chaisemartin and d’Haultfoeuille (2020)



**Notes:** This figure presents the event study coefficients following De Chaisemartin and d’Haultfoeuille (2020) for military demining during conflict. We present the point estimates as well as the 95% confidence interval. Standard errors clustered at the event level. The outcomes were computed using a radius of 5km around the demining.

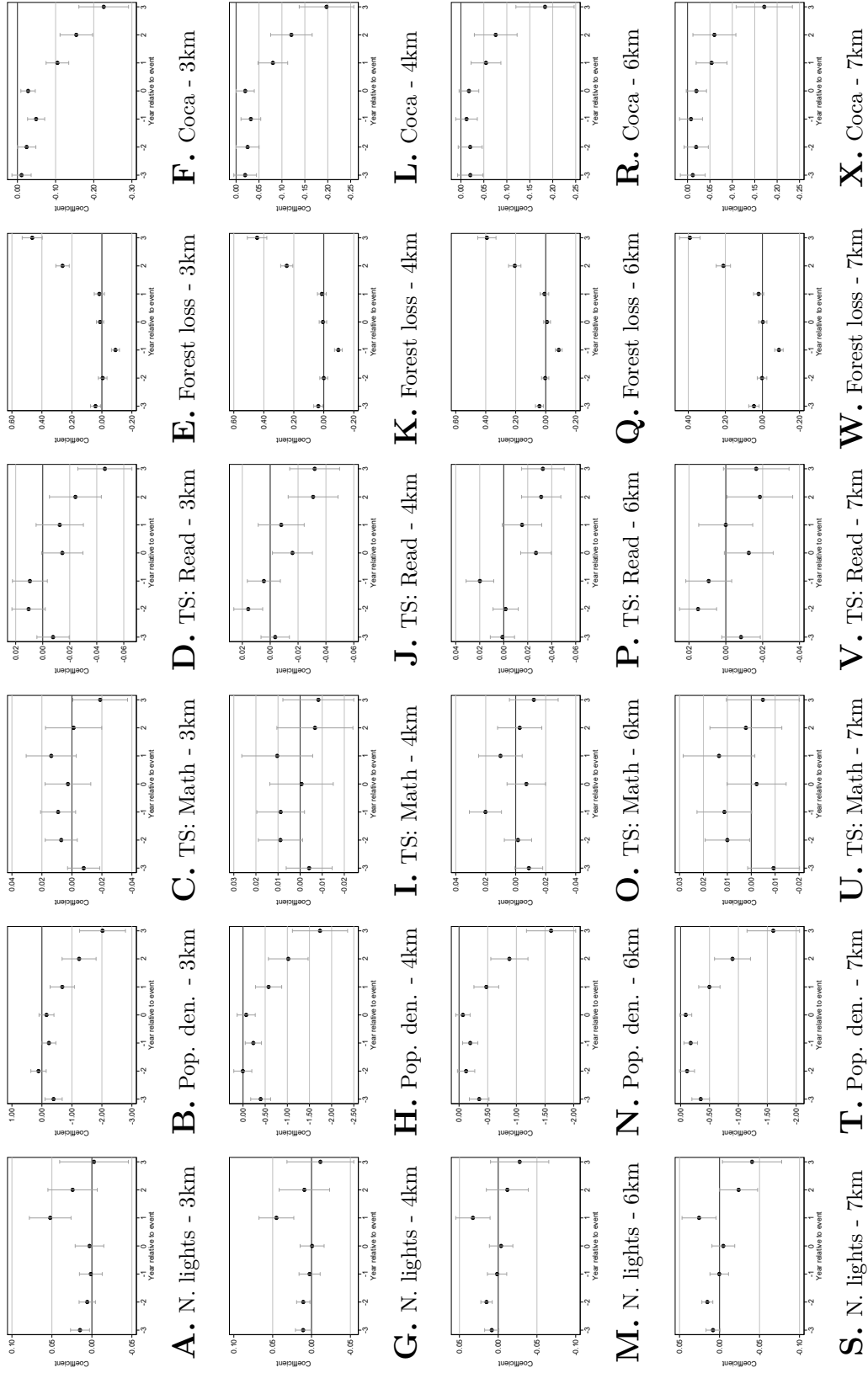
FIGURE A20. Humanitarian demining during peace over different radii



**Notes:** This figure presents the event study coefficients following Callaway and Sant'Anna (2020) for the treatment of humanitarian demining during peace. We present the point estimates as well as the 95% confidence interval. The outcomes were computed using radii of 3, 4, 6, and 7km around the demining event.

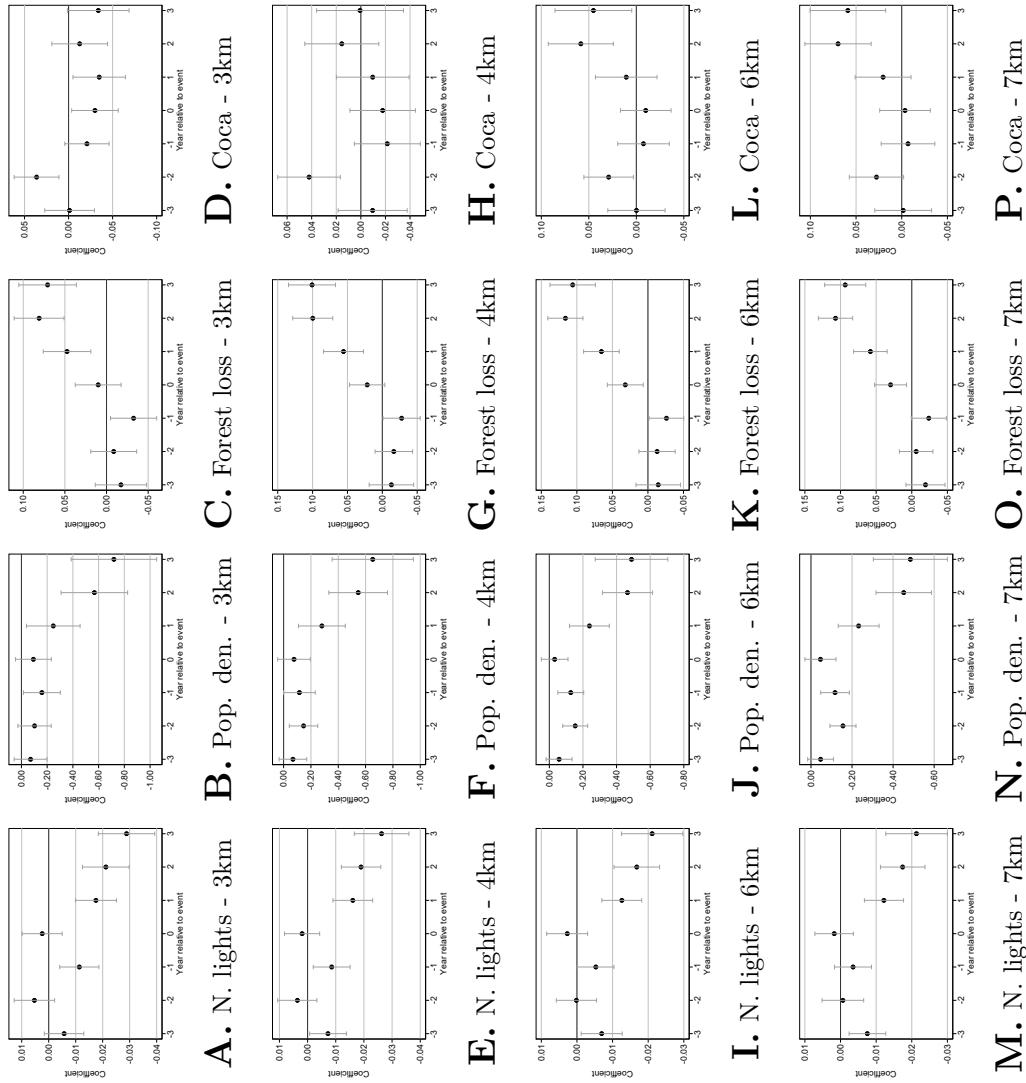


FIGURE A21. Military demining during peace over different radii



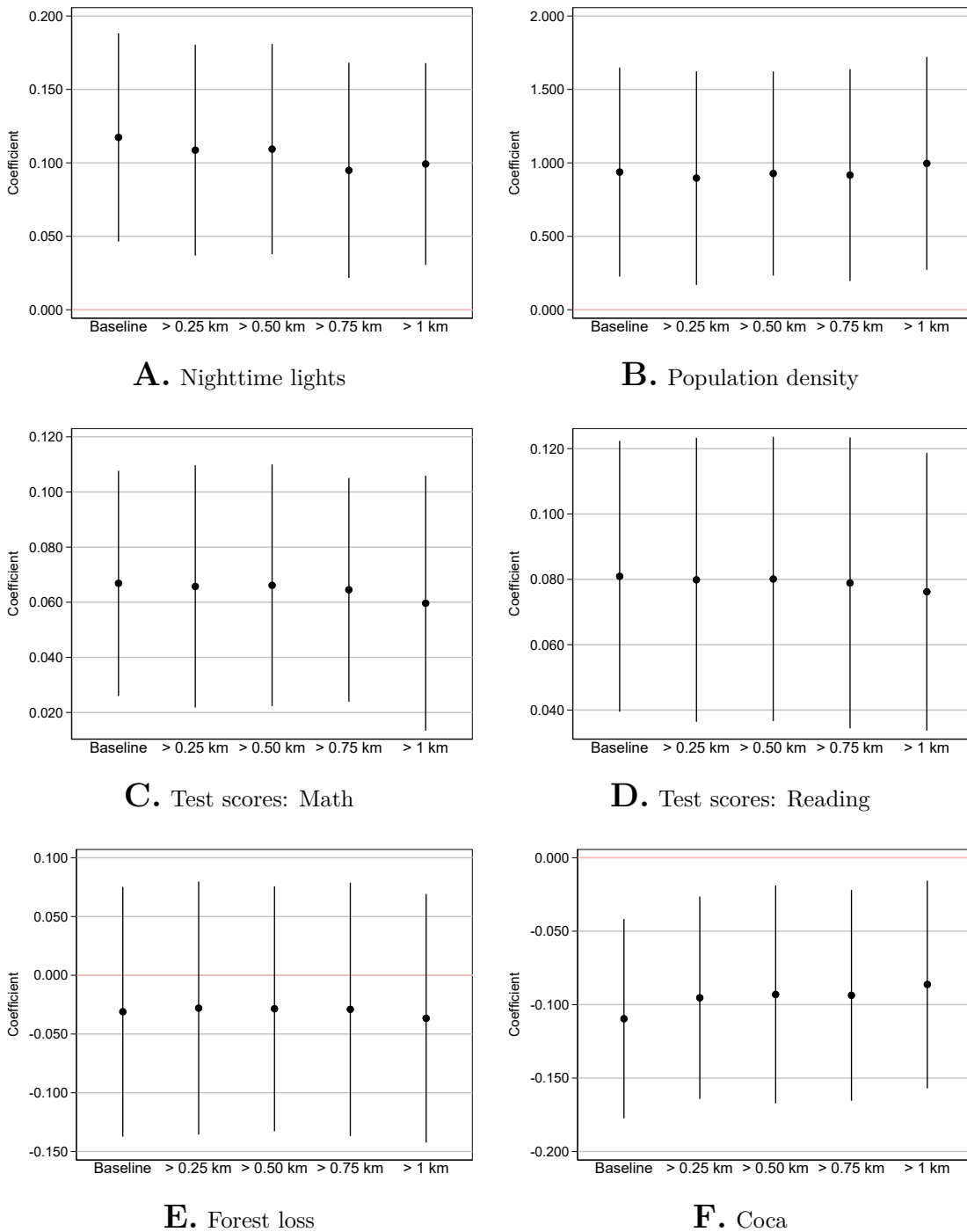
**Notes:** This figure presents the event study coefficients following Callaway and Sant'Anna (2020) for the treatment of military demining during peace. We present the point estimates as well as the 95% confidence interval. The outcomes were computed using radii of 3, 4, 6, and 7km around the demining event.

FIGURE A22. Military demining during conflict over different radii



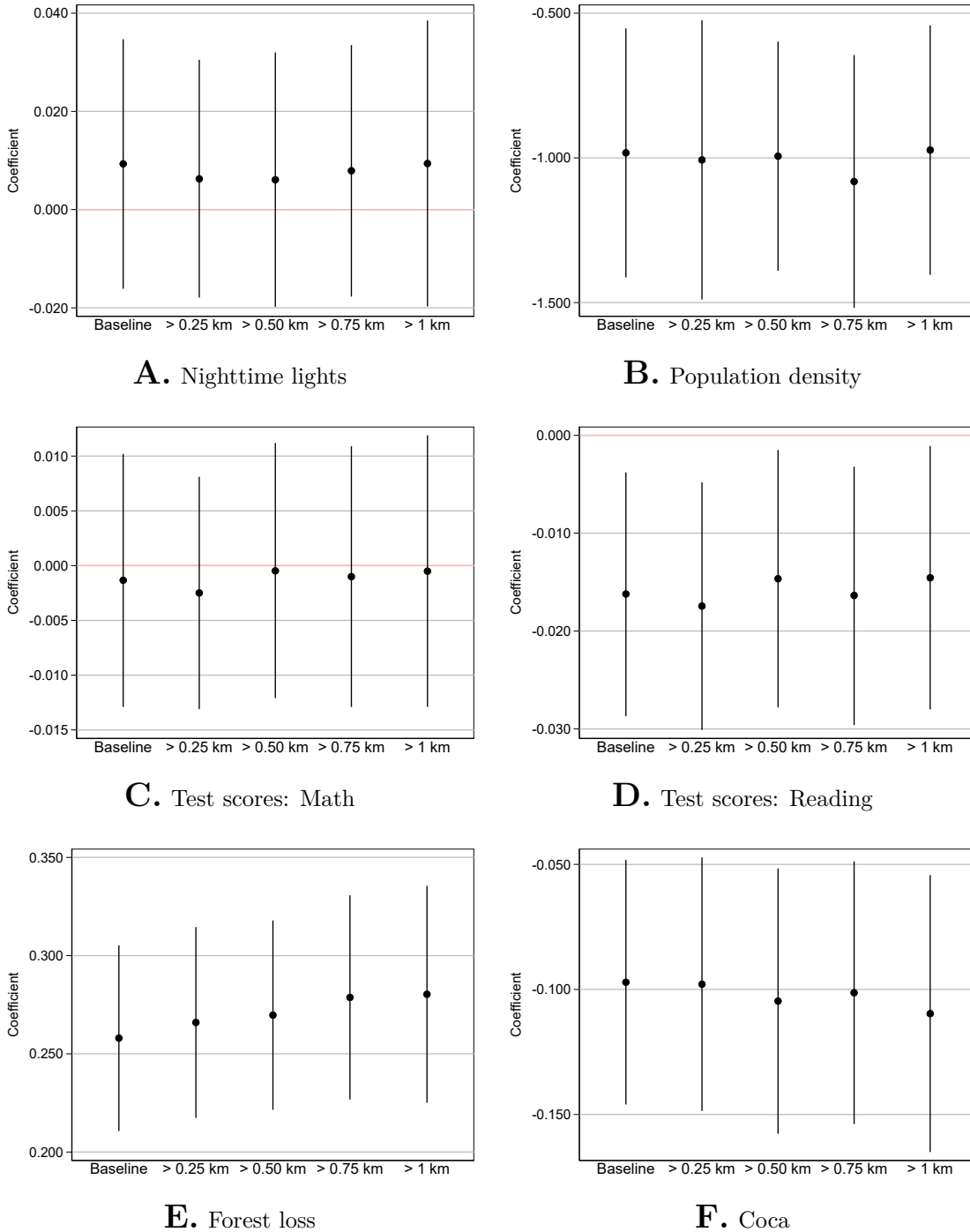
**Notes:** This figure presents the event study coefficients following Callaway and Sant'Anna (2020) for the treatment of military demining during peace. We present the point estimates as well as the 95% confidence interval. The outcomes were computed using radii of 3, 4, 6, and 7km around the demining event.

FIGURE A23. Humanitarian demining during peace: Excluding events to closest village



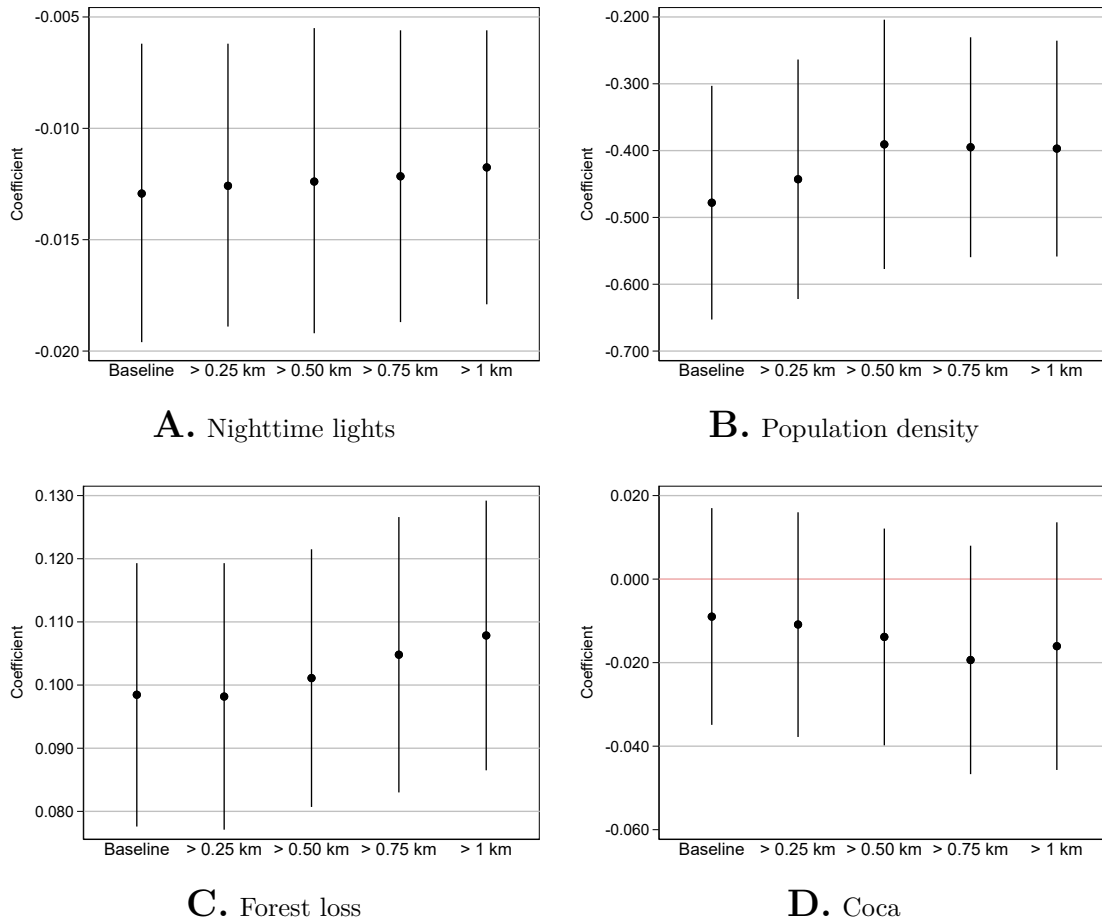
**Notes:** This figure presents the overall ATT following Callaway and Sant’Anna (2020) for the treatment of humanitarian demining during peace. We present the overall ATT from our baseline specification in Panel A from Table 1 and excluding events closest to villages at 250 m, 500 m, 750 m, and 1 km. We present the point estimates as well as the 95% confidence interval. The outcomes were computed using a radius of 5km around the demining

FIGURE A24. Military demining during peace: Excluding events to closest village



**Notes:** This figure presents the overall ATT following Callaway and Sant’Anna (2020) for the treatment of military demining during peace. We present the overall ATT from our baseline specification in Panel A from Table 2 and excluding events closest to villages at 250 m, 500 m, 750 m, and 1 km. We present the point estimates as well as the 95% confidence interval. The outcomes were computed using a radius of 5km around the demining

FIGURE A25. Military demining during peace: Excluding events to closest village



**Notes:** This figure presents the overall ATT following Callaway and Sant’Anna (2020) for the treatment of military demining during conflict. We present the overall ATT from our baseline specification in Panel A from Table 3 and excluding events closest to villages at 250 m, 500 m, 750 m, and 1 km. We present the point estimates as well as the 95% confidence interval. The outcomes were computed using a radius of 5km around the demining

TABLE A1. Two-way fixed effects decomposition and weights

During:	(1)	(2)	(3)
	Peace		Conflict
	Humanitarian	Military	Military
<b>Panel A: Bacon decomposition</b>			
Treated (T) vs Never treated (C)	0.847	0.267	0.550
Early treated (T) vs Late treated (C)	0.100	0.461	0.339
Late treated (T) vs Early treated (C)	0.053	0.272	0.111
<b>Panel B: Negative weights</b>			
Share of negative weights	0.000	0.272	0.124
Share of sum of negative weights	0.000	0.166	0.061

**Notes:** This table presents the decomposition of the two-way fixed effects model from equation (4.1) for humanitarian demining during peace (column 1), military demining during peace (column 2), and military demining during conflict (column 3). In panel A, we present the [Goodman-Bacon \(2021\)](#) decomposition, (T) stand for treated units and (C) for comparison units. In panel B, we present the share of negative weights and the relevance of them following [De Chaisemartin and d'Haultfoeuille \(2020\)](#).

TABLE A2. Robustness to “group” overall ATT

	(1)	(2)	(3)	(4)	(5)	(6)
			Test scores:			
Dep. variable:	Nighttime Lights	Population density	Math	Reading	Forest loss	Coca
<b>Panel A: Humanitarian demining during peace</b>						
Post demining	0.077*** (0.031)	0.759** (0.295)	0.051** (0.020)	0.062*** (0.021)	-0.033 (0.059)	-0.073** (0.035)
<b>Panel B: Military demining during peace</b>						
Post demining	0.017 (0.013)	-0.977*** (0.205)	-0.003 (0.006)	-0.018*** (0.006)	0.269*** (0.025)	-0.107*** (0.023)
<b>Panel C: Military demining during conflict</b>						
Post demining	-0.014*** (0.003)	-0.463*** (0.096)	—	—	0.045*** (0.009)	-0.025** (0.012)
Observations (Panel A)	5983	7460	6960	6960	7460	7460
Observations (Panel B)	90504	100560	69340	69370	100560	100560
Observations (Panel C)	213000	213000	—	—	213000	213000
Treated (Panel A)	291	294	283	283	294	294
Treated (Panel B)	9630	9630	6641	6641	9630	9630
Treated (Panel C)	15150	15150	—	—	15150	15150
Never treated (Panel A)	449	452	413	413	452	452
Never treated (Panel B)	426	426	293	296	426	426
Never treated (Panel C)	2600	2600	—	—	2600	2600
Average dep var (Panel A)	1.675	34.407	0.207	0.222	3.290	0.636
Average dep var (Panel B)	0.973	33.381	0.149	0.168	3.721	2.392
Average dep var (Panel C)	0.435	26.349	—	—	3.403	1.728

**Notes:** This table presents the overall ATT following Callaway and Sant’Anna (2020). Panel A presents the results for humanitarian demining during peace, panel B for military demining during peace, and panel C for military demining during conflict. *Post demining* is the weighted average across cohorts of the cohort average treatment effects with weights proportional to the cohort size. The set of controls include not yet treated and never treated. The outcomes were computed using a radius of 5km around the demining. Panel Bootstrap standard errors clustered at the event level. \* is significant at the 10% level, \*\* is significant at the 5% level, \*\*\* is significant at the 1% level.

TABLE A3. Robustness to different level clustered standard errors

	(1)	(2)	(3)	(4)	(5)	(6)
			Test scores:			
Dep. variable:	Nighttime Lights	Population density	Math	Reading	Forest loss	Coca
<b>Humanitarian demining during peace</b>						
<b>Panel A: Vereda level</b>						
Post demining	0.117* (0.060)	0.938* (0.539)	0.067** (0.026)	0.081*** (0.028)	-0.031 (0.080)	-0.110** (0.047)
<b>Panel B: Grids of 15x15Km</b>						
Post demining	0.117* (0.064)	0.938 (0.663)	0.067* (0.038)	0.081** (0.039)	-0.031 (0.093)	-0.110* (0.063)
<b>Military demining during peace</b>						
<b>Panel C: Vereda level</b>						
Post demining	0.009 (0.116)	-0.983 (0.655)	0.005 (0.011)	-0.011 (0.012)	0.258 (0.215)	-0.097** (0.049)
<b>Panel D: Grids of 15x15Km</b>						
Post demining	0.009 (0.064)	-0.983 (0.600)	0.005 (0.014)	-0.011 (0.016)	0.258** (0.105)	-0.097 (0.074)
<b>Military demining during conflict</b>						
<b>Panel E: Vereda level</b>						
Post demining	-0.013** (0.005)	-0.478*** (0.151)	—	—	0.098*** (0.019)	-0.009 (0.026)
<b>Panel F: Grids of 15x15Km</b>						
Post demining	-0.013** (0.007)	-0.478** (0.208)	—	—	0.098*** (0.034)	-0.009 (0.036)
Observations (Panel A & B)	5983	7460	6960	6960	7460	7460
Observations (Panel C & D)	90504	100560	69340	69370	100560	100560
Observations (Panel E & F)	213000	213000	—	—	213000	213000
Treated (Panel A & B)	291	294	283	283	294	294
Treated (Panel C & D)	9630	9630	6641	6641	9630	9630
Treated (Panel E & F)	15150	15150	—	—	15150	15150
Never treated (Panel A & B)	449	452	413	413	452	452
Never treated (Panel C & D)	426	426	293	296	426	426
Never treated (Panel E & F)	2600	2600	—	—	2600	2600
Average dep var (Panel A & B)	1.675	34.407	0.207	0.222	3.290	0.636
Average dep var (Panel C & D)	0.973	33.381	0.149	0.168	3.721	2.392
Average dep var (Panel E & F)	0.435	26.349	—	—	3.403	1.728

**Notes:** This table presents the overall ATT following Callaway and Sant'Anna (2020). Panels A and B present the results for humanitarian demining during peace, panels C and D for military demining during peace, and panels E and F for military demining during conflict. *Post demining* is the weighted average across cohorts of the cohort average treatment effects with weights proportional to the cohort size. The set of controls include not yet treated and never treated. The outcomes were computed using a radius of 5km around the demining. Panels A, C, E present bootstrap standard errors clustered at the vereda level, and in panels B, D, F present bootstrap standard errors clustered at grids of size 15x15Km. \* is significant at the 10% level, \*\* is significant at the 5% level, \*\*\* is significant at the 1% level.



TABLE A4. Robustness to Spatial Autocorrelation

	(1)	(2)	(3)	(4)	(5)	(6)
			Test scores:			
Dep. variable:	Nighttime Lights	Population density	Math	Reading	Forest loss	Coca
<b>Panel A: Humanitarian demining during peace</b>						
Post demining	0.109*** (0.037)	0.930*** (0.351)	0.067*** (0.022)	0.081*** (0.021)	-0.031 (0.052)	-0.109*** (0.035)
<b>Panel B: Military demining during peace</b>						
Post demining	0.009 (0.013)	-0.977*** (0.215)	0.005 (0.006)	-0.011* (0.006)	0.258*** (0.026)	-0.095*** (0.026)
<b>Panel C: Military demining during conflict</b>						
Post demining	-0.013*** (0.003)	-0.479*** (0.087)	—	—	0.099*** (0.011)	-0.009 (0.014)
Observations (Panel A)	5983	7460	6960	6960	7460	7460
Observations (Panel B)	90504	100560	69340	69370	100560	100560
Observations (Panel C)	213000	213000	—	—	213000	213000
Treated (Panel A)	291	294	283	283	294	294
Treated (Panel B)	9630	9630	6641	6641	9630	9630
Treated (Panel C)	15150	15150	—	—	15150	15150
Never treated (Panel A)	449	452	413	413	452	452
Never treated (Panel B)	426	426	293	296	426	426
Never treated (Panel C)	2600	2600	—	—	2600	2600
Average dep var (Panel A)	1.675	34.407	0.207	0.222	3.290	0.636
Average dep var (Panel B)	0.973	33.381	0.149	0.168	3.721	2.392
Average dep var (Panel C)	0.435	26.349	—	—	3.403	1.728

**Notes:** This table presents the overall ATT following [Callaway and Sant'Anna \(2020\)](#). Panel A presents the results for humanitarian demining during peace, panel B for military demining during peace, and panel C for military demining during conflict. *Post demining* is the weighted average across cohorts of the cohort average treatment effects with weights proportional to the cohort size. The set of controls include not yet treated and never treated. The outcomes were computed using a radius of 5km around the demining. We residualized our outcomes from Moran's I eigenvectors, which capture spatial autocorrelation among our buffers ([Bauman et al., 2018](#)). Panel Bootstrap standard errors clustered at the event level. \* is significant at the 10% level, \*\* is significant at the 5% level, \*\*\* is significant at the 1% level.

TABLE A5. Electricity subscribers and nighttime lights

Dep. variable:	(1) Electricity subscribers per capita	(2) Nighttime lights
Humanitarian demining $\times$ Post	0.005 (0.005)	
Electricity subscribers per capita		1.293*** (0.232)
Observations	10,244	9,220
R-squared	0.938	0.677
Municipality FE	Yes	No
Department-Year FE	Yes	Yes
Baseline controls	Yes	Yes

**Notes:** This table presents the effect of humanitarian demining on electricity subscribers, as well as the relationship between electricity subscribers and nighttime lights. *Humanitarian demining* is a dummy that takes the value one for municipalities where humanitarian demining took place. *Post* is a dummy that takes the value one from 2013 onward. *Electricity subscribers per capita* is the number of electricity subscribers per inhabitant. The source of these data is Information System of Utilities (SUI from its Spanish acronym). *Nighttime lights* is the hyperbolic sine transformation of the average nighttime luminosity. Standard errors clustered at the municipality level. \* is significant at the 10% level, \*\* is significant at the 5% level, \*\*\* is significant at the 1% level.

TABLE A6. The local effects of humanitarian demining on test scores by school degree

	(1)	(2)	(3)	(4)	(5)	(6)
	Test scores:					
	Math			Reading		
Dep. variable:	3°	5°	9°	3°	5°	9°
Post demining	0.075*** (0.019)	0.055*** (0.019)	0.031** (0.013)	0.084*** (0.018)	0.082*** (0.020)	0.029*** (0.011)
Observations	6880	6900	4940	6890	6940	4940
Treated	282	281	224	282	283	223
Never treated	406	409	270	407	411	271
Average dep var	0.177	0.147	0.052	0.172	0.183	0.060

**Notes:** This table presents the overall ATT following [Callaway and Sant'Anna \(2020\)](#) for humanitarian demining during peace. *Post demining* is the weighted average of all group-time average treatment effects with weights proportional to group size. The set of controls include not yet treated and never treated. The outcomes were computed using a radius of 5km around the demining. Bootstrap standard errors clustered at the municipality level. \* is significant at the 10% level, \*\* is significant at the 5% level, \*\*\* is significant at the 1% level.

TABLE A7. Treatment effects based on a sample with higher overlap

	(1)	(2)	(3)	(4)	(5)	(6)
			Test scores:			
Dep. variable:	Nighttime Lights	Population density	Math	Reading	Forest loss	Coca
<b>Panel A: Humanitarian demining</b>						
Post demining	0.123*** (0.038)	1.254*** (0.364)	0.069*** (0.022)	0.073*** (0.022)	-0.006 (0.057)	-0.074** (0.034)
<b>Panel B: Military demining</b>						
Post demining	0.067*** (0.023)	0.703* (0.376)	-0.015 (0.013)	-0.026** (0.013)	-0.087** (0.038)	-0.313*** (0.054)
Observations (Panel A)	5291	6470	6170	6170	6470	6470
Observations (Panel B)	32958	36620	25690	25780	36620	36620
Treated (Panel A)	265	267	261	261	267	267
Treated (Panel B)	3580	3580	2515	2524	3580	3580
Never treated (Panel A)	377	380	356	356	380	380
Never treated (Panel B)	82	82	54	54	82	82
Average dep var (Panel A)	1.646	34.115	0.209	0.223	3.201	0.445
Average dep var (Panel B)	1.076	39.315	0.175	0.194	3.078	1.213
P-value diff.	0.207	0.292	0.001	0.000	0.237	0.000

**Notes:** This table presents the overall ATT following Callaway and Sant'Anna (2020) for the treatments of humanitarian demining and military demining during peace. *Post demining* is the weighted average of all group-time average treatment effects with weights proportional to group size. The set of controls include not yet treated and never treated. The outcomes were computed using a radius of 5km around the demining. The sample is restricted to the optimal selection rule from Crump et al. (2009) over the propensity score, probability of being a humanitarian demining. The covariates used to predict the probability were selected following Belloni et al. (2014) machine learning algorithm, which selects the best covariates predicting humanitarian demining. The selected covariates were a poverty index, the logarithm of population, the distance to the closest department's capital, a coca suitability index, the distance to the country's capital, and the distance to the closest national park. From the pool of potential covariates that also included a rurality index, elevation, precipitation, and temperature. We also present the p-value that test the difference between the coefficients of Panel A and B. Bootstrap standard errors clustered at the event level. \* is significant at the 10% level, \*\* is significant at the 5% level, \*\*\* is significant at the 1% level.

TABLE A8. Bias from hypothesized linear pre-trend

	(1)	(2)	(3)	(4)
	Estimate	Slope	Unconditional bias	Conditional bias
<b>Panel A: Humanitarian demining during peace</b>				
Nighttime lights	0.117	0.034	0.086	0.082
Population density	0.938	0.444	1.109	1.081
Test scores: Math	0.067	0.019	0.048	0.046
Test scores: Reading	0.081	0.020	0.050	0.046
Forest loss	-0.031	0.058	0.145	0.136
Coca	-0.110	0.046	0.114	0.115
<b>Panel B: Military demining during peace</b>				
Nighttime lights	0.009	0.003	0.010	0.009
Population density	-0.983	0.080	0.200	0.207
Test scores: Math	-0.001	0.004	0.010	0.010
Test scores: Reading	-0.016	0.004	0.011	0.011
Forest loss	0.258	0.012	0.030	0.029
Coca	-0.097	0.011	0.028	0.029
<b>Panel C: Military demining during conflict</b>				
Nighttime lights	-0.013	0.003	0.006	0.006
Population density	-0.478	0.039	0.098	0.109
Forest loss	0.098	0.012	0.031	0.032
Coca	-0.009	0.012	0.030	0.029

**Notes:** This table presents the estimated parameter from our baseline specification in Panel A from Tables 1 to 3 and the main estimates based on Roth (2021). In column 2, we present the pre-trend that has a 50% power of being detected given the precision of the estimates in the pre-period. In column 3, we present the average bias suggested by this trend, while in column 4, the bias from the adjusted pre-trend that takes into account the pre-testing bias that arises from the fact that the analysis shown is conditional on passing a pre-test.

TABLE A9. Demining and conflict

Dep. variable:	(1)	(2)	(3)	(4)	(5)
	Total	Forced displacement	Army	FARC	Paramilitaries
Conflict $\times$ Military demining	0.20*** (0.02)	0.10*** (0.02)	0.04*** (0.01)	0.15*** (0.01)	0.05*** (0.01)
Military demining	0.11*** (0.02)	0.08*** (0.02)	0.01 (0.01)	-0.03*** (0.01)	-0.02** (0.01)
Humanitarian demining	-0.18*** (0.05)	-0.30*** (0.05)	-0.02 (0.02)	-0.09*** (0.03)	-0.01 (0.01)
Observations	17,472	12,110	17,472	17,472	17,472
R-squared	0.912	0.883	0.309	0.458	0.473
Municipality fixed effect	Yes	Yes	Yes	Yes	Yes
Dept-year fixed effect	Yes	Yes	Yes	Yes	Yes
Baseline controls	Yes	Yes	Yes	Yes	Yes
Municipalities	1092	1073	1092	1092	1092
Average dep var	3.466	4.775	0.0473	0.0750	0.0850

**Notes:** This table presents the relationship between military and humanitarian demining and conflict related victims and attacks (see equation 6.1). All the dependent variables are measured using the hyperbolic sine transformation. Military (Humanitarian) demining is the total number of landmines demined by the army (humanitarian organizations) transformed using the hyperbolic sine transformation.  $Conflict_t$  is a dummy that takes the value one before 2013. All the regressions include the set of covariates: the total population, log distance to the capital, a rurality index, and a poverty index all measured in 2003 and interacted with year fixed effects. Robust standard errors are clustered at the municipality level. \* is significant at the 10% level, \*\* is significant at the 5% level, \*\*\* is significant at the 1% level.

TABLE A10. The local effects of demining on fires

	(1)	(2)	(3)
	<b>Dep. variable: <i>Fires</i></b>		
During:	Peace		Conflict
	Humanitarian	Military	Military
Post demining	-0.048 (0.049)	0.033** (0.014)	0.049*** (0.010)
Observations	7460	100560	213000
Treated	294	9630	15150
Never treated	452	426	2600
Average dep var	0.592	0.879	0.865

**Notes:** This table presents the overall ATT following Callaway and Sant'Anna (2020) for fires for the three demining treatments. *Post demining* is the weighted average of all group-time average treatment effects with weights proportional to group size. The set of controls include not yet treated and never treated. The outcomes were computed using a radius of 5km around the demining. Bootstrap standard errors clustered at the event level. \* is significant at the 10% level, \*\* is significant at the 5% level, \*\*\* is significant at the 1% level.

TABLE A11. The local effects of demining during peace on illegal gold mining

	(1)	(2)	(3)	(4)
Demining:	Humanitarian		Military	
Dep. variable:	Area illegal gold mining	Illegal gold mining	Area illegal gold mining	Illegal gold mining
Post demining	-0.103 (0.146)	-0.048 (0.046)	0.059*** (0.020)	0.019*** (0.006)
Observations	2020	2020	11604	11604
Treated	53	53	2475	2475
Never treated	452	452	426	426
Average dep var	0.216	0.065	0.415	0.084

**Notes:** This table presents the overall ATT following Callaway and Sant'Anna (2020) for illegal alluvial gold meaning using the humanitarian and military demining treatments during peace. *Post demining* is the weighted average of all group-time average treatment effects with weights proportional to group size. The set of controls include not yet treated and never treated. The outcomes were computed using a radius of 5km around the demining. \* significant at the 10% level, \*\* is significant at the 5% level, \*\*\* is significant at the 1% level



TABLE A12. Heterogeneous effects for coca cultivation by substitution program

	(1)	(2)	(3)	(4)	(5)	(6)
	<b>Dependent variable: <i>Coca cultivation</i></b>					
	During peace:					
Demining:	Humanitarian			Military		
PNIS:	No	Yes	p-value diff.	No	Yes	p-value diff.
Post demining	0.131* (0.068)	-0.511** (0.257)	0.016	-0.047 (0.046)	-0.149*** (0.037)	0.084
Observations	4650	580		7860	11060	
Treated	60	14		639	828	
Never treated	405	44		147	278	
Average dep var	0.766	1.213		2.212	5.089	

**Notes:** This table presents the overall ATT following Callaway and Sant'Anna (2020) for coca cultivation using for both demining treatments during peace. *Post demining* is the weighted average of all group-time average treatment effects with weights proportional to group size. The set of controls include not yet treated and never treated. The outcomes were computed using a radius of 5km around the demining. For each type of treatment, we divide the treated and never treated into events in municipalities without and within the PNIS program. \* is significant at the 10% level, \*\* is significant at the 5% level, \*\*\* is significant at the 1% level.

TABLE A13. Summary statistics

	(1)	(2)	(3)	(4)	(5)
	Average	Standard deviation	90th percentile	50th percentile	10th percentile
<b>Panel A: Humanitarian demining during peace</b>					
Nighttime lights	1.675	0.917	2.632	1.932	0.255
Population density	34.407	30.752	66.871	26.092	7.839
Test scores: Math	0.191	0.35	0.861	0	0
Test scores: Reading	0.205	0.374	0.927	0	0
Forest loss	3.29	1.305	4.928	3.38	1.473
Coca hts	0.636	1.488	2.809	0	0
N Schools	10.138	7.809	19	9	1
<b>Panel B: Military demining during peace</b>					
Nighttime lights	0.973	1.156	2.611	0.196	0
Population density	33.381	56.452	76.459	14.362	2.332
Test scores: Math	0.103	0.266	0.618	0	0
Test scores: Reading	0.116	0.292	0.75	0	0
Forest loss	3.721	1.521	5.48	3.963	1.443
Coca hts	2.392	2.636	6.369	1.346	0
N Schools	8.084	12.688	18	5	1
<b>Panel C: Military demining during conflict</b>					
Nighttime lights	0.435	0.849	1.759	0	0
Population density	26.349	50.676	56.319	10.939	2.269
Forest loss	3.403	1.39	5.063	3.566	1.408
Coca hts	1.728	2.137	5.022	0	0

**Notes:** This table presents summary statistics for our main variables of interest. The outcomes in panels A, B, and C were computed using a radius of 5km around the demining event.

TABLE A14. Robustness to other estimation methods: Humanitarian demining

	(1)	(2)	(3)	(4)	(5)	(6)
			Test scores:			
Dep. variable:	Nighttime Lights	Population density	Math	Reading	Forest loss	Coca
<b>Panel A: Borusyak et al. (2021)</b>						
Post demining	0.081** (0.033)	0.203 (0.456)	0.042*** (0.015)	0.041*** (0.016)	0.108** (0.048)	-0.119*** (0.037)
<b>Panel B: De Chaisemartin and d'Haultfoeuille (2020)</b>						
Post demining	0.080*** (0.028)	0.528** (0.214)	0.061*** (0.023)	0.071*** (0.020)	0.002 (0.042)	-0.084*** (0.024)
<b>Panel C: Wooldridge (2021)</b>						
Post demining	0.074** (0.033)	0.202 (0.457)	0.042*** (0.015)	0.041*** (0.016)	0.108** (0.049)	-0.119*** (0.037)
Observations	5983	7460	6960	6960	7460	7460
Treated	291	294	283	283	294	294
Never treated	449	452	413	413	452	452
Average dep var	1.675	34.407	0.207	0.222	3.290	0.636

**Notes:** This table presents the overall ATT using two different models for the treatment of humanitarian demining during peace. The outcomes were computed using a radius of 5km around the demining. In Panel A, we present the imputation method suggested by [Borusyak et al. \(2021\)](#). We estimate the model for the window of three year around the event. In Panel B, we present the model suggested by [De Chaisemartin and d'Haultfoeuille \(2020\)](#) computing the ATT for the three years after the the treatment. In Panel C, we present the ATT suggested by [Wooldridge \(2021\)](#) where the estimated coefficient is a weighted average of group-year dummies after treatment. Standard errors are clustered at the event level. \* is significant at the 10% level, \*\* is significant at the 5% level, \*\*\* is significant at the 1% level.

TABLE A15. Robustness to other estimation methods: Military demining during peace

	(1)	(2)	(3)	(4)	(5)	(6)
			Test scores:			
Dep. variable:	Nighttime Lights	Population density	Math	Reading	Forest loss	Coca
<b>Panel A: Borusyak et al. (2021)</b>						
Post demining	0.039*** (0.015)	-1.524*** (0.253)	0.010* (0.006)	-0.004 (0.006)	0.097*** (0.018)	-0.153*** (0.028)
<b>Panel B: De Chaisemartin and d'Haultfoeuille (2020)</b>						
Post demining	0.030*** (0.011)	-0.719*** (0.155)	-0.002 (0.004)	-0.019*** (0.005)	0.145*** (0.016)	-0.090*** (0.022)
<b>Panel C: Wooldridge (2021)</b>						
Post demining	0.039** (0.015)	-1.524*** (0.253)	0.010* (0.006)	-0.004 (0.006)	0.096*** (0.018)	-0.153*** (0.028)
Observations	90504	100560	69340	69340	100560	100560
Treated	9630	9630	6641	6641	9630	9630
Never treated	426	426	293	296	426	426
Average dep var	0.973	33.381	0.149	0.168	3.721	2.392

**Notes:** This table presents the overall ATT using two different models for the treatment of military demining during peace. The outcomes were computed using a radius of 5km around the demining. In Panel A, we present the imputation method suggested by [Borusyak et al. \(2021\)](#). We estimate the model for the window of three year around the event. In Panel B, we present the model suggested by [De Chaisemartin and d'Haultfoeuille \(2020\)](#) computing the ATT for the three years after the the treatment. In Panel C, we present the ATT suggested by [Wooldridge \(2021\)](#) where the estimated coefficient is a weighted average of group-year dummies after treatment. Standard errors are clustered at the event level. Standard errors are clustered at the event level. \* is significant at the 10% level, \*\* is significant at the 5% level, \*\*\* is significant at the 1% level.

TABLE A16. Robustness to other estimation methods: Military demining during conflict

	(1)	(2)	(3)	(4)
Dep. variable:	Nighttime Lights	Population density	Forest Loss	Coca
<b>Panel A: Borusyak et al. (2021)</b>				
Post demining	-0.020*** (0.003)	-0.716*** (0.103)	0.017* (0.009)	0.018 (0.014)
<b>Panel B: De Chaisemartin and d'Haultfoeuille (2020)</b>				
Post demining	-0.014*** (0.003)	-0.424*** (0.064)	0.043*** (0.010)	-0.020* (0.012)
<b>Panel C: Wooldridge (2021)</b>				
Post demining	-0.020*** (0.003)	-0.716*** (0.103)	0.017** (0.009)	0.018 (0.014)
Observations	213000	213000	213000	213000
Treated	15150	15150	15150	15150
Never treated	2600	2600	2600	2600
Average dep var	0.435	26.349	3.403	1.728

**Notes:** This table presents the overall ATT using two different models for the treatment of military demining during conflict. The outcomes were computed using a radius of 5km around the demining. In Panel A, we present the imputation method suggested by [Borusyak et al. \(2021\)](#). We estimate the model for the window of three year around the event. In Panel B, we present the model suggested by [De Chaisemartin and d'Haultfoeuille \(2020\)](#) computing the ATT for the three years after the the treatment. In Panel C, we present the ATT suggested by [Wooldridge \(2021\)](#) where the estimated coefficient is a weighted average of group-year dummies after treatment. Standard errors are clustered at the event level. Standard errors are clustered at the event level. \* is significant at the 10% level, \*\* is significant at the 5% level, \*\*\* is significant at the 1% level.

TABLE A17. Spillover effects of humanitarian demining during peace: Covariates

	(1)	(2)	(3)	(4)	(5)	(6)
			Test scores:			
Dep. variable:	Nighttime Lights	Population density	Math	Reading	Forest loss	Coca
<b>Panel A: Buffer of 3 km</b>						
Post demining	0.093** (0.047)	1.053*** (0.375)	0.034 (0.026)	0.062** (0.027)	-0.094 (0.066)	-0.061*** (0.024)
<b>Panel B: Buffer of 5 km</b>						
Post demining	0.114*** (0.044)	1.019** (0.425)	0.047* (0.027)	0.072*** (0.026)	-0.061 (0.068)	-0.069*** (0.026)
<b>Panel C: Buffer of 7 km</b>						
Post demining	0.115** (0.047)	0.938** (0.411)	0.058** (0.028)	0.080*** (0.027)	-0.007 (0.059)	-0.076*** (0.027)
Observations	5983	7460	6960	6960	7460	7460
Treated	291	294	283	283	294	294
Never treated	449	452	413	413	452	452
Average dep var	1.675	34.407	0.207	0.222	3.290	0.636

**Notes:** This table presents the overall ATT following Callaway and Sant'Anna (2020) for the treatment of humanitarian demining during peace. *Post demining* is the weighted average of all group-time average treatment effects with weights proportional to group size. The outcomes were computed using a radius of 5km around the demining. Panel A, B, and C present the results controlling with an indicator that takes the value one if there was a demining event within a buffer of 3, 5, and 7 Km the year before the demining. In the case of the never treated the dummy takes the value if there was a demining event the year before within a buffer of 3km/5km/7km. We include the covariate following Sant'Anna and Zhao (2020) doubly robust method. We include the covariate following Sant'Anna and Zhao (2020) doubly robust method. Bootstrap standard errors clustered at the event level. \* is significant at the 10% level, \*\* is significant at the 5% level, \*\*\* is significant at the 1% level

TABLE A18. Spillover effects of military demining during peace: Covariates

	(1)	(2)	(3)	(4)	(5)	(6)
			Test scores:			
Dep. variable:	Nighttime Lights	Population density	Math	Reading	Forest loss	Coca
<b>Panel A: Buffer of 3 km</b>						
Post demining	0.015 (0.013)	-0.614*** (0.200)	0.006 (0.006)	-0.010 (0.007)	0.207*** (0.025)	-0.118*** (0.027)
<b>Panel B: Buffer of 5 km</b>						
Post demining	0.028* (0.015)	-0.325* (0.183)	0.002 (0.006)	-0.015** (0.007)	0.155*** (0.024)	-0.111*** (0.029)
<b>Panel C: Buffer of 7 km</b>						
Post demining	0.038** (0.015)	-0.224 (0.194)	0.003 (0.007)	-0.012 (0.008)	0.127*** (0.026)	-0.085*** (0.029)
Observations	90504	100560	69340	69370	100560	100560
Treated	9630	9630	6641	6641	9630	9630
Never treated	426	426	293	296	426	426
Average dep var	0.973	33.381	0.149	0.168	3.721	2.392

**Notes:** This table presents the overall ATT following Callaway and Sant'Anna (2020) for the treatment of military demining during peace. *Post demining* is the weighted average of all group-time average treatment effects with weights proportional to group size. The outcomes were computed using a radius of 5km around the demining. Panel A, B, and C present the results controlling with an indicator that takes the value one if there was a demining event within a buffer of 3, 5, and 7 Km the year before the demining. In the case of the never treated the dummy takes the value if there was a demining event the year before within a buffer of 3km/5km/7km. We include the covariate following Sant'Anna and Zhao (2020) doubly robust method. Bootstrap standard errors clustered at the municipality level. \* is significant at the 10% level, \*\* is significant at the 5% level, \*\*\* is significant at the 1% level

TABLE A19. Spillover effects of military demining during conflict: Covariates

Dep. variable:	(1) Nighttime Lights	(2) Population	(3) Forest Loss	(4) Coca
<b>Panel A: Buffer of 3 km</b>				
Post demining	-0.009*** (0.003)	-0.422*** (0.093)	0.115*** (0.010)	-0.003 (0.014)
<b>Panel B: Buffer of 5 km</b>				
Post demining	-0.010*** (0.003)	-0.373*** (0.088)	0.119*** (0.011)	0.003 (0.014)
<b>Panel C: Buffer of 7 km</b>				
Post demining	-0.011*** (0.004)	-0.354*** (0.095)	0.119*** (0.011)	0.003 (0.014)
Average dep var	0.435	26.349	3.403	1.728
Observations	213000	213000	213000	213000
Treated	15150	15150	15150	15150
Never treated	2600	2600	2600	2600

**Notes:** This table presents the overall ATT following Callaway and Sant'Anna (2020) for the treatment of military demining during conflict. *Post demining* is the weighted average of all group-time average treatment effects with weights proportional to group size. The outcomes were computed using a radius of 5km around the demining. Panel A, B, and C present the results controlling with an indicator that takes the value one if there was a demining event within a buffer of 3, 5, and 7 Km the year before the demining. In the case of the never treated the dummy takes the value if there was a demining event the year before within a buffer of 3km/5km/7km. We include the covariate following Sant'Anna and Zhao (2020) doubly robust method. Bootstrap standard errors clustered at the municipality level. \* is significant at the 10% level, \*\* is significant at the 5% level, \*\*\* is significant at the 1% level



TABLE A20. Spillover effects of humanitarian demining during peace: Excluding treated buffers

	(1)	(2)	(3)	(4)	(5)	(6)
			Test scores:			
Dep. variable:	Nighttime Lights	Population density	Math	Reading	Forest loss	Coca
<b>Panel A: Buffer of 3 km</b>						
Post demining	0.098** (0.044)	0.699 (0.454)	0.054** (0.025)	0.058** (0.024)	0.062 (0.055)	-0.129*** (0.047)
<b>Panel B: Buffer of 5 km</b>						
Post demining	0.127*** (0.041)	0.829* (0.449)	0.063** (0.026)	0.072*** (0.026)	0.083 (0.058)	-0.134** (0.055)
<b>Panel C: Buffer of 7 km</b>						
Post demining	0.133*** (0.042)	0.877* (0.521)	0.079*** (0.028)	0.087*** (0.026)	0.068 (0.065)	-0.139** (0.056)
Observations (Panel A)	5208	6600	6110	6110	6600	6600
Observations (Panel B)	5035	6390	5900	5900	6390	6390
Observations (Panel C)	4902	6210	5720	5720	6210	6210
Treated (Panel A)	205	208	198	198	208	208
Treated (Panel B)	184	187	177	177	187	187
Treated (Panel C)	166	169	159	159	169	169
Never treated (Panel A)	449	452	413	413	452	452
Never treated (Panel B)	449	452	413	413	452	452
Never treated (Panel C)	449	452	413	413	452	452
Average dep var (Panel A)	1.662	33.197	0.205	0.221	3.297	0.703
Average dep var (Panel B)	1.659	32.964	0.204	0.220	3.291	0.716
Average dep var (Panel C)	1.663	32.868	0.203	0.220	3.288	0.728

**Notes:** This table presents the overall ATT following Callaway and Sant'Anna (2020) for the treatment of humanitarian demining during peace. *Post demining* is the weighted average of all group-time average treatment effects with weights proportional to group size. The outcomes were computed using a radius of 5km around the demining. Panel A, B, and C presents the results excluding from the sample treated buffers with at least one demining in the previous year of the event around 3 km, 5km, and 7km to the demining, respectively. Bootstrap standard errors clustered at the municipality level. \* is significant at the 10% level, \*\* is significant at the 5% level, \*\*\* is significant at the 1% level

TABLE A21. Spillover effects of military demining during peace: Excluding treated buffers

	(1)	(2)	(3)	(4)	(5)	(6)
			Test scores:			
Dep. variable:	Nighttime Lights	Population density	Math	Reading	Forest loss	Coca
<b>Panel A: Buffer of 3 km</b>						
Post demining	0.007 (0.015)	-0.548** (0.272)	0.004 (0.009)	-0.002 (0.009)	0.188*** (0.028)	-0.076** (0.032)
<b>Panel B: Buffer of 5 km</b>						
Post demining	-0.001 (0.019)	-0.384 (0.282)	0.006 (0.010)	0.005 (0.011)	0.175*** (0.032)	-0.050 (0.032)
<b>Panel C: Buffer of 7 km</b>						
Post demining	-0.007 (0.020)	-0.328 (0.319)	0.009 (0.011)	0.011 (0.012)	0.197*** (0.034)	-0.023 (0.037)
Observations (Panel A)	53082	58980	39170	39170	58980	58980
Observations (Panel B)	43542	48380	32110	32090	48380	48380
Observations (Panel C)	38007	42230	27890	27860	42230	42230
Treated (Panel A)	5472	5472	3624	3621	5472	5472
Treated (Panel B)	4412	4412	2918	2913	4412	4412
Treated (Panel C)	3797	3797	2496	2490	3797	3797
Never treated (Panel A)	426	426	293	296	426	426
Never treated (Panel B)	426	426	293	296	426	426
Never treated (Panel C)	426	426	293	296	426	426
Average dep var (Panel A)	0.971	34.491	0.158	0.178	3.550	2.244
Average dep var (Panel B)	0.977	34.402	0.159	0.180	3.497	2.222
Average dep var (Panel C)	0.974	34.205	0.160	0.181	3.468	2.204

**Notes:** This table presents the overall ATT following Callaway and Sant'Anna (2020) for the treatment of military demining during peace. *Post demining* is the weighted average of all group-time average treatment effects with weights proportional to group size. The outcomes were computed using a radius of 5km around the demining. Panel A, B, and C presents the results excluding from the sample treated buffers with at least one demining in the previous year of the event around 3 km, 5km, and 7km to the demining, respectively. Bootstrap standard errors clustered at the municipality level. \* is significant at the 10% level, \*\* is significant at the 5% level, \*\*\* is significant at the 1% level

TABLE A22. Spillover effects of military demining during conflict: Excluding treated buffers

Dep. variable:	(1) Nighttime Lights	(2) Population	(3) Forest Loss	(4) Coca
<b>Panel A: Buffer of 3 km</b>				
Post demining	-0.013*** (0.004)	-0.428*** (0.099)	0.056*** (0.013)	-0.019 (0.016)
<b>Panel B: Buffer of 5 km</b>				
Post demining	-0.018*** (0.004)	-0.606*** (0.119)	0.008 (0.016)	0.002 (0.019)
<b>Panel C: Buffer of 7 km</b>				
Post demining	-0.019*** (0.005)	-0.795*** (0.119)	-0.028 (0.017)	0.010 (0.021)
Average dep var (Panel A)	0.395	24.879	3.349	1.706
Average dep var (Panel B)	0.401	25.390	3.331	1.696
Average dep var (Panel C)	0.414	26.142	3.331	1.690
Treated (Panel A)	9110	9110	9110	9110
Treated (Panel B)	6554	6554	6554	6554
Treated (Panel C)	4836	4836	4836	4836
Never treated (Panel A)	2600	2600	2600	2600
Never treated (Panel B)	2600	2600	2600	2600
Never treated (Panel C)	2600	2600	2600	2600
Observations (Panel A)	140520	140520	140520	140520
Observations (Panel B)	109848	109848	109848	109848
Observations (Panel C)	89232	89232	89232	89232

**Notes:** This table presents the overall ATT following Callaway and Sant'Anna (2020) for the treatment of military demining during conflict. *Post demining* is the weighted average of all group-time average treatment effects with weights proportional to group size. The outcomes were computed using a radius of 5km around the demining. Panel A, B, and C presents the results excluding from the sample treated buffers with at least one demining in the previous year of the event around 3 km, 5km, and 7km to the demining, respectively. Bootstrap standard errors clustered at the municipality level. \* is significant at the 10% level, \*\* is significant at the 5% level, \*\*\* is significant at the 1% level

TABLE A23. Robustness to only using never treated as controls

	(1)	(2)	(3)	(4)	(5)	(6)
			Test scores:			
Dep. variable:	Nighttime Lights	Population density	Math	Reading	Forest loss	Coca
<b>Panel A: Humanitarian demining during peace</b>						
Post demining	0.112*** (0.035)	0.924** (0.367)	0.064*** (0.020)	0.079*** (0.021)	-0.043 (0.054)	-0.105*** (0.036)
<b>Panel B: Military demining during peace</b>						
Post demining	0.004 (0.018)	-0.888*** (0.340)	0.003 (0.009)	-0.004 (0.010)	0.259*** (0.031)	-0.129*** (0.038)
<b>Panel C: Military demining during conflict</b>						
Post demining	-0.018*** (0.004)	-0.606*** (0.121)	—	—	0.081*** (0.011)	0.014 (0.016)
Observations (Panel A)	5983	7460	6960	6960	7460	7460
Observations (Panel B)	90504	100560	69340	69370	100560	100560
Observations (Panel C)	213000	213000	—	—	213000	213000
Treated (Panel A)	291	294	283	283	294	294
Treated (Panel B)	9630	9630	6641	6641	9630	9630
Treated (Panel C)	15150	15150	—	—	15150	15150
Never treated (Panel A)	449	452	413	413	452	452
Never treated (Panel B)	426	426	293	296	426	426
Never treated (Panel C)	2600	2600	—	—	2600	2600
Average dep var (Panel A)	1.675	34.407	0.207	0.222	3.290	0.636
Average dep var (Panel B)	0.973	33.381	0.149	0.168	3.721	2.392
Average dep var (Panel C)	0.435	26.349	—	—	3.403	1.728

**Notes:** This table presents the overall ATT following Callaway and Sant'Anna (2020). Panel A presents the results for humanitarian demining during peace, panel B for military demining during peace, and panel C for military demining during conflict. *Post demining* is the weighted average of all group-time average treatment effects with weights proportional to group size. The set of controls include only the never treated. The outcomes were computed using a radius of 5km around the demining. Bootstrap standard errors clustered at the municipality level. \* is significant at the 10% level, \*\* is significant at the 5% level, \*\*\* is significant at the 1% level.

TABLE A24. Robustness to excluding never treated as controls

	(1)	(2)	(3)	(4)	(5)	(6)
			Test scores:			
Dep. variable:	Nighttime Lights	Population density	Math	Reading	Forest loss	Coca
<b>Panel A: Humanitarian demining during peace</b>						
Post demining	0.144** (0.070)	1.026 (0.831)	0.176*** (0.043)	0.211*** (0.040)	0.271** (0.137)	-0.267*** (0.091)
<b>Panel B: Military demining during peace</b>						
Post demining	0.014 (0.015)	-0.871*** (0.226)	0.012* (0.006)	-0.008 (0.007)	0.174*** (0.023)	-0.039 (0.025)
<b>Panel C: Military demining during conflict</b>						
Post demining	-0.006* (0.003)	-0.308*** (0.096)	—	—	0.137*** (0.010)	-0.033** (0.014)
Observations (Panel A)	2424	2940	2830	2830	2940	2940
Observations (Panel B)	86670	96300	66410	66410	96300	96300
Observations (Panel C)	181800	181800	—	—	181800	181800
Treated (Panel A)	291	294	283	283	294	294
Treated (Panel B)	9630	9630	6641	6641	9630	9630
Treated (Panel C)	15150	15150	—	—	15150	15150
Average dep var (Panel A)	1.733	38.005	0.212	0.223	3.312	0.338
Average dep var (Panel B)	0.981	33.375	0.150	0.169	3.708	2.299
Average dep var (Panel C)	0.425	25.971	—	—	3.386	1.710

**Notes:** This table presents the overall ATT following Callaway and Sant'Anna (2020). Panel A presents the results for humanitarian demining during peace, panel B for military demining during peace, and panel C for military demining during conflict. *Post demining* is the weighted average of all group-time average treatment effects with weights proportional to group size. The set of controls exclude the never treated. The outcomes were computed using a radius of 5km around the demining. Bootstrap standard errors clustered at the municipality level. \* is significant at the 10% level, \*\* is significant at the 5% level, \*\*\* is significant at the 1% level.

TABLE A25. Robustness to allow for one-year anticipation

	(1)	(2)	(3)	(4)	(5)	(6)
			Test scores:			
Dep. variable:	Nighttime Lights	Population density	Math	Reading	Forest loss	Coca
<b>Panel A: Humanitarian demining during peace</b>						
Post demining	0.078* (0.042)	0.953 (0.649)	0.035** (0.018)	0.032* (0.020)	0.124** (0.056)	-0.190*** (0.055)
<b>Panel B: Military demining during peace</b>						
Post demining	-0.004 (0.015)	-1.877*** (0.272)	0.008 (0.005)	-0.010 (0.006)	0.123*** (0.023)	-0.119*** (0.027)
<b>Panel C: Military demining during conflict</b>						
Post demining	-0.022*** (0.004)	-0.718*** (0.111)	—	—	0.083*** (0.010)	-0.001 (0.016)
Observations (Panel A)	5983	7460	6960	6960	7460	7460
Observations (Panel B)	90504	100560	69340	69370	100560	100560
Observations (Panel C)	213000	213000	—	—	213000	213000
Treated (Panel A)	291	294	283	283	294	294
Treated (Panel B)	9630	9630	6641	6641	9630	9630
Treated (Panel C)	15150	15150	—	—	15150	15150
Never treated (Panel A)	449	452	413	413	452	452
Never treated (Panel B)	426	426	293	296	426	426
Never treated (Panel C)	2600	2600	—	—	2600	2600
Average dep var (Panel A)	1.675	34.407	0.207	0.222	3.290	0.636
Average dep var (Panel B)	0.973	33.381	0.149	0.168	3.721	2.392
Average dep var (Panel C)	0.435	26.349	—	—	3.403	1.728

**Notes:** This table presents the overall ATT following Callaway and Sant'Anna (2020). We allow for an anticipation of the treatment of one-year. Panel A presents the results for humanitarian demining during peace, panel B for military demining during peace, and panel C for military demining during conflict. *Post demining* is the weighted average of all group-time average treatment effects with weights proportional to group size. The set of controls include not yet treated and never treated. The outcomes were computed using a radius of 5km around the demining. Bootstrap standard errors clustered at the municipality level. \* is significant at the 10% level, \*\* is significant at the 5% level, \*\*\* is significant at the 1% level.

TABLE A26. Robustness to different radii: Humanitarian demining during peace

	(1)	(2)	(3)	(4)	(5)	(6)
			Test scores:			
Dep. variable:	Nighttime Lights	Population density	Math	Reading	Forest loss	Coca
<b>Panel A: 3 km radius</b>						
Post demining	0.143*** (0.053)	0.907** (0.442)	0.039 (0.026)	0.058** (0.023)	0.013 (0.057)	-0.117*** (0.027)
<b>Panel B: 4 km radius</b>						
Post demining	0.181*** (0.046)	0.910** (0.410)	0.081*** (0.021)	0.093*** (0.021)	-0.021 (0.054)	-0.118*** (0.032)
<b>Panel C: 6 km radius</b>						
Post demining	0.068** (0.034)	0.941*** (0.347)	0.101*** (0.018)	0.110*** (0.023)	-0.023 (0.051)	-0.122*** (0.039)
<b>Panel D: 7 km radius</b>						
Post demining	0.064** (0.027)	1.120*** (0.352)	0.076*** (0.020)	0.091*** (0.022)	-0.035 (0.044)	-0.135*** (0.041)
Observations	5983	7460	6960	6960	7460	7460
Treated	291	294	283	283	294	294
Never treated	449	452	413	413	452	452
Average dep var (Panel A)	1.829	32.578	0.137	0.147	2.249	0.415
Average dep var (Panel B)	1.729	33.354	0.166	0.179	2.830	0.533
Average dep var (Panel C)	1.630	35.842	0.211	0.229	3.684	0.727
Average dep var (Panel D)	1.600	37.372	0.223	0.244	4.018	0.807

**Notes:** This table presents the overall ATT following Callaway and Sant'Anna (2020) for the treatment of humanitarian demining during peace. *Post demining* is the weighted average of all group-time average treatment effects with weights proportional to group size. The set of controls include not yet treated and never treated. Panels A, B, C, and D present the results where dependent variable was computed using a radius of 3, 4, 6, and 7km around the event, respectively. Bootstrap standard errors clustered at the municipality level. \* is significant at the 10% level, \*\* is significant at the 5% level, \*\*\* is significant at the 1% level.

TABLE A27. Robustness to different radii: Military demining during peace

Dep. variable:	(1)	(2)	Test scores:		(5)	(6)
	Nighttime Lights	Population density	Math	Reading	Forest loss	Coca
<b>Panel A: 3 km radius</b>						
Post demining	0.019 (0.015)	-1.193*** (0.335)	-0.002 (0.007)	-0.019** (0.007)	0.300*** (0.025)	-0.134*** (0.026)
<b>Panel B: 4 km radius</b>						
Post demining	0.014 (0.013)	-1.021*** (0.257)	-0.001 (0.006)	-0.014** (0.007)	0.290*** (0.025)	-0.108*** (0.025)
<b>Panel C: 6 km radius</b>						
Post demining	0.003 (0.012)	-0.945*** (0.196)	-0.002 (0.006)	-0.018*** (0.006)	0.246*** (0.023)	-0.087*** (0.026)
<b>Panel D: 7 km radius</b>						
Post demining	-0.005 (0.012)	-0.947*** (0.193)	0.006 (0.005)	-0.001 (0.006)	0.246*** (0.023)	-0.075*** (0.025)
Observations	90504	100560	69340	69370	100560	100560
Treated	9630	9630	6641	6641	9630	9630
Never treated	426	426	293	296	426	426
Average dep var (Panel A)	0.958	37.058	0.069	0.078	2.670	1.828
Average dep var (Panel B)	0.968	34.944	0.088	0.099	3.255	2.130
Average dep var (Panel C)	0.974	32.171	0.116	0.131	4.103	2.625
Average dep var (Panel D)	0.975	31.400	0.129	0.145	4.427	2.834

**Notes:** This table presents the overall ATT following Callaway and Sant'Anna (2020) for the treatment of military demining during peace. *Post demining* is the weighted average of all group-time average treatment effects with weights proportional to group size. The set of controls include not yet treated and never treated. Panels A, B, C, and D present the results where dependent variable was computed using a radius of 3, 4, 6, and 7km around the event, respectively. Bootstrap standard errors clustered at the municipality level. \* is significant at the 10% level, \*\* is significant at the 5% level, \*\*\* is significant at the 1% level.



TABLE A28. Robustness to different radii: Military demining during conflict

	(1)	(2)	(3)	(4)
Dep. variable:	Nighttime Lights	Population	Forest Loss	Coca
<b>Panel A: 3 km radius</b>				
Post demining	-0.016*** (0.004)	-0.552*** (0.130)	0.074*** (0.011)	-0.036*** (0.012)
<b>Panel B: 4 km radius</b>				
Post demining	-0.015*** (0.003)	-0.533*** (0.102)	0.094*** (0.011)	-0.020 (0.013)
<b>Panel C: 6 km radius</b>				
Post demining	-0.011*** (0.003)	-0.424*** (0.077)	0.099*** (0.010)	-0.001 (0.014)
<b>Panel D: 7 km radius</b>				
Post demining	-0.011*** (0.003)	-0.414*** (0.072)	0.092*** (0.010)	0.008 (0.014)
Observations	213000	213000	213000	213000
Treated	15150	15150	15150	15150
Never treated	2600	2600	2600	2600
Average dep var (Panel A)	0.424	29.553	2.339	1.251
Average dep var (Panel B)	0.430	27.580	2.928	1.507
Average dep var (Panel C)	0.440	25.550	3.797	1.923
Average dep var (Panel D)	0.446	25.221	4.134	2.101

**Notes:** This table presents the overall ATT following Callaway and Sant'Anna (2020) for the treatment of military demining during conflict. *Post demining* is the weighted average of all group-time average treatment effects with weights proportional to group size. The set of controls include not yet treated and never treated. Panels A, B, C, and D present the results where dependent variable was computed using a radius of 3, 4, 6, and 7km around the event, respectively. Bootstrap standard errors clustered at the municipality level. \* is significant at the 10% level, \*\* is significant at the 5% level, \*\*\* is significant at the 1% level.

TABLE A29. Humanitarian demining during peace and local activity excluding ETCR zones

	(1)	(2)	(3)	(4)	(5)	(6)
			Test scores:			
Dep. variable:	Nighttime Lights	Population density	Math	Reading	Forest loss	Coca
<b>Panel A: Humanitarian demining - 1km around ETCR zone</b>						
Post demining	0.117*** (0.037)	0.963*** (0.357)	0.067*** (0.022)	0.081*** (0.021)	-0.038 (0.053)	-0.111*** (0.035)
Observations	5970	7440	6950	6950	7440	7440
Treated	289	292	282	282	292	292
Never treated	449	452	413	413	452	452
Average dep var	1.672	34.273	0.207	0.222	3.290	0.633

**Notes:** This table presents the overall ATT following Callaway and Sant'Anna (2020) for the treatments of humanitarian demining. *Post demining* is the weighted average of all group-time average treatment effects with weights proportional to group size. The set of controls include not yet treated and never treated. The outcomes were computed using a radius of 5km around the demining. Those buffers that intersect a buffer of 1km around the ETCR zones and demined after 2016, were removed from the sample. Bootstrap standard errors clustered at the event level. \* is significant at the 10% level, \*\* is significant at the 5% level, \*\*\* is significant at the 1% level.

TABLE A30. Robustness to outliers

Dep. variable:	(1)	(2)	(3)	(4)	(5)	(6)	(7)	(8)	(9)	(10)	(11)	(12)
	Test scores:											
Nighttime Lights	Population density		Math		Reading		Forest loss		Coca			
Winsorizing at:	1%	3%	1%	3%	1%	3%	1%	3%	1%	3%	1%	3%
<b>Panel A: Humanitarian demining during peace</b>												
Post demining	0.112*** (0.036)	0.119*** (0.039)	1.025*** (0.364)	1.011*** (0.336)	0.067*** (0.023)	0.067*** (0.022)	0.081*** (0.021)	0.081*** (0.020)	-0.030 (0.056)	-0.026 (0.049)	-0.101*** (0.032)	-0.072*** (0.035)
<b>Panel B: Military demining during peace</b>												
Post demining	0.008 (0.013)	0.005 (0.013)	-0.983*** (0.214)	-1.343*** (0.186)	-0.001 (0.006)	-0.001 (0.006)	-0.016*** (0.006)	-0.016*** (0.006)	0.257*** (0.026)	0.253*** (0.024)	-0.083*** (0.025)	-0.023 (0.025)
<b>Panel C: Military demining during conflict</b>												
Post demining	-0.013*** (0.003)	-0.012*** (0.003)	-0.478*** (0.092)	-0.295*** (0.067)	-	-	-	-	0.100*** (0.011)	0.098*** (0.010)	-0.004 (0.014)	0.003 (0.013)
Observations (Panel A)	7460	7460	7460	7460	6960	6960	6960	6960	7460	7460	7460	7460
Observations (Panel B)	90504	90504	100560	100560	69340	69340	69370	69370	100560	100560	100560	100560
Observations (Panel C)	213000	213000	213000	213000	-	-	-	-	213000	213000	213000	213000
Treated (Panel A)	294	294	294	294	283	283	283	283	294	294	294	294
Treated (Panel B)	9630	9630	9630	9630	6641	6641	6641	6641	9630	9630	9630	9630
Treated (Panel C)	15150	15150	15150	15150	-	-	-	-	15150	15150	15150	15150
Never treated (Panel A)	452	452	452	452	413	413	413	413	452	452	452	452
Never treated (Panel B)	426	426	426	426	293	293	296	296	426	426	426	426
Never treated (Panel C)	2600	2600	2600	2600	-	-	-	-	2600	2600	2600	2600
Average dep var (Panel A)	1.341	1.332	33.598	33.103	0.207	0.207	0.222	0.222	3.288	3.280	0.632	0.610
Average dep var (Panel B)	0.969	0.961	33.381	29.951	0.149	0.149	0.168	0.168	3.718	3.711	2.390	2.377
Average dep var (Panel C)	0.429	0.417	26.349	22.547	-	-	-	-	3.399	3.392	1.723	1.711

**Notes:** This table presents the overall ATT following Callaway and Sant'Anna (2020). Panel A presents the results for humanitarian demining during peace, panel B for military demining during peace, and panel C for military demining during conflict. *Post demining* is the weighted average of all group-time average treatment effects with weights proportional to group size. The set of controls include not yet treated and never treated. The outcomes were computed using a radius of 5km around the demining. Odd columns present the results for the dependent variable winsorized at 1% and even columns at 3%. Bootstrap standard errors clustered at the municipality level. \* is significant at the 10% level, \*\* is significant at the 5% level, \*\*\* is significant at the 1% level.

